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Evaluation Of The Use Of DUMPStat To Detect The Impact Of Landfills On Groundwater Quality

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LIST OF ABBREVIATIONS

ACF Autocorrelation Function

ACL Alternate Concentration Limit

BOD Biochemical Oxygen Demand

CC Combined Shewhart-CUSUM Control Chart

COD Chemical Oxygen Demand

CUSUM Cumulative Sum

DATCP Wisconsin Department of Agriculture, Trade and Consumer

Protection

DILHR Wisconsin Department of Industry, Labor, and Human

Relations

DMZ Design Management Zone

DOT Wisconsin Department of Transportation

DUMPStat Downgradient Upgradient Monitoring Program Statistics

EPA United States Environmental Protection Agency

ES Enforcement Standard

IQR Interquartile Range

KW Kruskal-Wallis Test

NR Natural Resources

PAL Preventative Action Limit

PL Prediction Limit

TDS Total Dissolved Solids

TOC Total Organic Carbon

TOX Total Organic Halogen

VOC Volatile Organic Compounds

WDNR or DNR Wisconsin Department of Natural Resources

CHAPTER 1

INTRODUCTION

1.1 Objectives

The primary objective of this research is to evaluate the use of the computer program DUMPStat as an alternative or supplement to the use of indicator preventive action limits (PAL) for determining whether a landfill is in compliance with groundwater quality standards.

1.2 Background

Careful monitoring of the impact of landfills on groundwater quality is a critical component of Wisconsin's groundwater protection program. Currently, the Wisconsin Department of Natural Resources (WDNR) administrates the monitoring programs in accordance with Chapter NR 140 of the Wisconsin Administrative Code. This regulation sets two types of groundwater quality standards: preventive action limits (PAL's) and enforcement standards (ES). The lower PAL serves two purposes. First, the PAL is used in design codes so that contamination is prevented through use of stringent designs. Second, the PAL is used as a "trigger" for remedial actions. PAL's are used to prevent contamination but the ES defines when a violation has occurred. If an ES is

violated, the landfill is subject to immediate enforcement action. If substances detected in groundwater exceed either the ES or PAL for a health or welfare parameter, the WDNR may set an alternative concentration limit (ACL). The WDNR also sets PAL's for indicator parameters which occur naturally and may indicate contamination. These indicator parameters provide an early warning for possible contamination.

The focus of this research is on PAL's for indicator parameters. NR 140 specifies the PAL for indicator parameters as the mean plus three standard deviations or a minimum increase, based on at least eight background samples. NR 140 does not specify how ACL's are set. However, the WDNR Waste Management Program policy is to use the sample mean plus two sample standard deviations. Consequently, both the indicator PAL and ACL are determined statistically. Also, the indicator PAL and the ACL are both determined on a well specific basis (intrawell) when possible. The health and welfare PAL and ES are uniform across the landfill as well as across the state.

The PAL procedure is set by law and rule while the ACL procedure is set by rule only. The rule can be changed by changing policy and operating procedures but the law can be changed only by an act of the legislature. Since changing the PAL procedure would require a substantial legislative change, this was our focus. Also, the nature of the current statistical procedures is intrawell, so the focus of this research is intrawell.

Gibbons(1994) has developed statistical procedures for analyzing groundwater quality data at landfills. These procedures are included in the

statistical package DUMPStat (Downgradient, Upgradient Monitoring Program Statistics), developed by Gibbons and Discerning Systems, Inc. DUMPStat has two intrawell methods which could potentially replace the indicator PAL, prediction limits (PL's) and the combined Shewhart-CUSUM control charts (CC's). The objective of this research focuses on statistical issues, although operational issues are also considered. This report is intended to provide information which the WDNR can use to determine whether or not DUMPStat should be adopted in whole or in part as a tool for assessing groundwater impacts at landfills. Again, we focused on the well specific indicator PAL, since the use of DUMPStat would require a legislative change.

1.3 Project Summary and Organization

In order to evaluate the intrawell DUMPStat algorithms, we proceeded in three steps. First, we applied DUMPStat to indicator parameter data from the Portage County Landfill. From this analysis, we found that the DUMPStat algorithms are far more conservative than the PAL. For indicator parameters, the use of DUMPStat would result in more violations than would the use of PAL's. This would be true for both upgradient and downgradient wells.

Our second step was to test the underlying assumptions of the DUMPStat calculations: independence, stationarity, and normality. These assumptions are not critical to the use of PAL's, as they are not based on statistical significance.

In order to evaluate these assumptions, we tested three indicator parameters for 26 upgradient wells at ten landfills. Normality was rejected at 12 of 26 wells and the assumptions of independence was rejected at 15 of 26 wells. In addition to these assumptions, we tested each well and parameter for trend. For this test, at least 50 percent of the wells and parameters exhibited an upward or downward trend.

The third step in this study was to evaluate how violations of assumptions and presence of trend affects DUMPStat and the use of the indicator PAL. DUMPStat was applied to indicator parameters for the 26 wells at the ten landfills. Next, we determined whether or not each violation was in a data set that rejected normality, rejected independence or exhibited trend. For all parameters combined, we found that over 60 percent of the violations for DUMPStat and PAL were in these data sets.

This report is intended to provide an in depth analysis of the results reported above as well as additional background information. The report is organized in the following manner:

- Chapter one is a simple introduction.
- Chapter two reviews Wisconsin's groundwater monitoring regulations.
- Chapter three reviews all the options available in DUMPStat. Our focus is on intrawell prediction limits and combined Shewhart-CUSUM control charts, however, we discuss in brief the other DUMPStat options.

- Chapter four is the Portage County Landfill case study. In this case study we
 evaluate all wells using PL's, CC's and PAL's using both eight and 25
 background samples.
- Chapter five evaluates the assumptions of parametric tests for 26 upgradient
 wells at ten landfills. Then we evaluated these wells using PL's, CC's and
 PAL's. Lastly, we considered how violations of the assumptions affect each
 statistical test.
- Chapter six is the conclusion and recommendation section.

CHAPTER 2

MONITORING AT WISCONSIN LANDFILLS

2.1 Wisconsin Regulations

As stated in Chapter 1, the primary legislation regulating Wisconsin landfills is NR 140. The purpose of the legislation is:

"...to establish groundwater quality standards for substances detected in or having a reasonable probability of entering the groundwater resources of the state; to specify scientifically valid procedures for determining if a numerical standard has been attained or exceeded; to specify procedures for establishing points of standards application, and for evaluating groundwater monitoring data; to establish ranges of responses the department may require if a groundwater standard is attained or exceeded; and to provide for exemptions for facilities, practices and activities regulated by the department."

The purpose of Chapter 2 of this report is to outline the regulations of the state, to outline the monitoring requirements, to review the calculations required for PAL, and to show how these requirements relate to intrawell and interwell analysis.

Subchapter III of NR 140 applies only to the DNR while subchapters I and II apply to the DNR as well as the following Wisconsin government agencies:

- Department of Agriculture, Trade and Consumer Protection (DATCP),
- Department of Industry, Labor and Human Relations (DILHR),
- Department of Natural Resources (DNR), and
- Department of Transportation (DOT).

Also, NR 140 applies to facilities and practices which may affect groundwater quality but are regulated by other agencies. However, the regulation does not apply to mining or prospecting activities, which are covered by other regulations.

2.2 Preventive Action Limits and Enforcement Standards

Wisconsin has two water quality standards for each parameter. The preventive action limit (PAL) is the lower of the two; exceedance of this value raises a red flag. The enforcement standard (ES) is the maximum allowable level. NR 140.10 specifies PAL's and ES's for 101 public health related standards. The PAL is ten percent of the ES for parameters with carcinogenic effects. For example, the ES for benzene is 5 micrograms per liter while the PAL for is 0.5 micrograms per liter. For other parameters with mutagenic, teratogenic or interactive effects the PAL is twenty percent of the ES. For example, for cyanide the ES is 200 micrograms per liter and the PAL is 40 micrograms per liter. Similarly, NR 140.12 specifies PAL and ES for parameters having aesthetic or other public welfare concerns. The PAL for these are 50 percent of the ES. For example the ES and PAL for chloride are 250 mg/l and 125 mg/l respectively.

PAL's are calculated for indicator parameters on a well specific basis.

These are the parameters that are measured regularly but do not directly affect public health or welfare. For these parameters the PAL is calculated as:

$$PAL = \overline{x} + max[3s, M]$$
 (2.1)

where \overline{x} is the sample mean, s is the sample standard deviation and M is the minimum increase. s and \overline{x} are estimated from a minimum of eight background samples for the particular constituent. Table 2-1 lists the minimum increase, M, for indicator parameters.

Parameter	Minimum Increase, mg/l
Alkalinity	100
Biochemical Oxygen Demand (BOD)	25
Boron	2
Calcium	25
Chemical Oxygen Demand (COD)	25
Magnesium	25
Ammonia Nitrogen	2
Organic Nitrogen	2
Total Nitrogen	5
Potassium	5
Sodium	10
Specific Conductance	200 μmhos/cm
Total Dissolved Solids (TDS)	200
Hardness	100
Total Organic Carbon (TOC)	1
Total Organic Halogen (TOX)	0.25

Table 2-1 Minimum Increase for Indicator Parameters

2.3 Monitoring Requirements

NR 140 covers many of the important factors necessary for environmental protection. However it does not provide specific information regarding the requirements for environmental monitoring for landfills. NR 507 defines the monitoring requirements. The purpose of this rule is:

"... to help insure that efficient, nuisance-free and environmentally acceptable solid waste management procedures are practiced in this state, to outline environmental monitoring requirements at solid waste facilities and to implement groundwater standards according to NR 140 and ch. 160 stats."

This rule governs all environmental monitoring for solid waste disposal facilities except hazardous waste facilities and mining operations.

Prior to accepting waste, a landfill must have a WDNR approved sampling plan. This plan must include the number and location of all monitoring wells. The number of required monitoring wells is based on the facility size, waste type, design and hydrogeologic or geologic properties. However, each new municipal solid waste landfill (MSW) must have a minimum of four Subtitle D wells.

The baseline monitoring requirements vary depending on the type of waste accepted. For example, the groundwater monitoring requirements for a municipal solid waste landfill are different from a landfill that accepts foundry waste. Table 2-2 lists the detection parameters and sampling frequencies currently required for routine sampling at municipal solid waste landfills. Table 2-3 lists the detection parameters and sampling frequency requirements for

other waste types. All proposed landfills must take four samples for all the detection parameters listed in Table 2-2 or 2-3 plus any public health and welfare parameters not included as detection monitoring parameters prior to submittal of the feasibility report. If exemptions to the groundwater standards for public health and welfare are granted in the feasibility determination, four more rounds of sampling for those parameters are required so that an ACL can be calculated and approved as a part of the plan of operation. For detection parameters, the next four samples must be taken before submittal of the plan of operation in order to accumulate a total of eight background samples. Each sample must have a minimum of 30 days between sampling rounds.

Detection Parameters	Frequency for All Wells	Frequency for Sub D. Wells
Alkalinity	Semi-annual	Semi-annual
Chloride	Semi-annual	Semi-annual
COD	Semi-annual	Semi-annual
Field Conductivity	Semi-annual	Semi-annual
Field pH	Semi-annual	Semi-annual
Field Temperature	Semi-annual	Semi-annual
Groundwater Elev.	Semi-annual	Semi-annual
Hardness	Semi-annual	Semi-annual
*VOC Scan	Annual	Semi-annual

Table 2-2 Monitoring Requirements for Municipal Solid Waste Landfills

^{*} A list of VOC parameters for detection monitoring is in Appendix III of NR 507

Waste Type	Detection Parameters	Frequency for All Wells
Municipal Solid Waste	Alkalinity	Semi-annual
Combustor Residue	Boron	Semi-annual
	Cadmium	Semi-annual
	Chloride	Semi-annual
	COD	Semi-annual
	Field Conductivity	Semi-annual
	Field pH	Semi-annual
	Field Temperature	Semi-annual
	Groundwater Elevation	Semi-annual
	Hardness	Semi-annual
	Lead	Semi-annual
	Selenium	Semi-annual
	Sulfate	Semi-annual
Paper Mill Sludge	Ammonia Nitrogen	Semi-annual
	Alkalinity	Semi-annual
	Chloride	Semi-annual
	COD	Semi-annual
	Field Conductivity	Semi-annual
	Field pH	Semi-annual
	Field Temperature	Semi-annual
	Groundwater Elevation	Semi-annual
	Hardness	Semi-annual
	Nitrate + Nitrate as N	Semi-annual
	Sulfate	Semi-annual
Fly or Bottom Ash	Alkalinity	Semi-annual
	Boron	Semi-annual
	COD	Semi-annual
	Field Conductivity	Semi-annual
	Field pH	Semi-annual
	Field Temperature	Semi-annual
	Groundwater Elevation	Semi-annual
	Hardness	Semi-annual
	Sulfate	Semi-annual
		Semi-annual Semi-annual
	A II - I' - it -	Semi-annual
Foundry Waste	Alkalinity	Semi-annual
	COD Field Conductivity	Semi-annual
	Field Conductivity	Semi-annual Semi-annual
	Field pH	Semi-annual Semi-annual
	Field Temperature Fluoride	Semi-annual Semi-annual
		I .
	Groundwater Elevation	Semi-annual
	Hardness	Semi-annual
Other Solid Waste	Sodium	Semi-annual string by the WDNR
Other Solid Waste	As Specified in Wi	Iding by the AADIAK

Table 2-3 Monitoring Requirements for Other Solid Waste Landfills

2-6

2.4 Responses When a Groundwater Standard is Attained or Exceeded

Wisconsin's rules differ slightly from the U.S. EPA Subtitle D. Subtitle D requires analysis of groundwater samples from designated wells for indicator parameters. These parameters include fifteen metals and VOC's. Wisconsin's approved plan allows for routine monitoring of detection parameters as listed in Table 2-2 and 2-3. Wisconsin's detection monitoring does not include routine metals monitoring. Also, the VOC monitoring is as specified in Table 2-2. If a PAL is exceeded for an indicator parameter, assessment monitoring in accordance with Wisconsin's approved plan may be required.

If the concentration of a substance, including indicator parameters, in groundwater attains or exceeds a PAL this does not automatically lead to assessment monitoring. Sections NR 140.24, NR 140.26 and Chapter NR 508 outline the appropriate responses for exceedances. First, the landfill owner must notify the WDNR in writing of the exceedance. Second, the owner must provide a preliminary analysis of the cause and significance of the concentration. In response the WDNR evaluates the notice and preliminary analysis. The WDNR may require the owner to further assess the cause and significance of the increased concentration and prepare a report evaluating the degree and extent of the problem.

The owner may demonstrate that a reported value represents a false exceedance of the groundwater standard. While reviewing this information the WDNR may look at a number of different items to determine the cause and significance of the exceedance. They may consider:

- the location of the monitoring well;
- the specific characteristics of the site;
- the nature of the substances involved and the likelihood of migration;
- background water quality;
- reliability of sampling data;
- public health, welfare and environmental effects;
- probability that a PAL or ES may be attained or exceeded outside the design management zone (DMZ) of the landfill;
- performance of the facility;
- other known or suspected sources of the substance in the area;
- hydrogeologic conditions;
- extent of groundwater contamination;
- alternate responses.

Within 30 days, if the WDNR does not agree that the exceedance is a false positive, the owner must begin assessment monitoring. If the exceedance is for one of the inorganic compounds listed in Table 2-2 then the WDNR may allow an alternative assessment monitoring plan. Also, an alternate monitoring plan may be approved to include only the parameters that are present in the leachate collection system. If the full assessment monitoring proceeds, all the parameters

in Appendix II of NR 507 must be included in the monitoring program. This includes but is not limited to the following:

Public Welfare Parameters	Public Health Parameters
Copper	Arsenic
Manganese	Barium
Sulfate	Cadmium
Zinc	Chromium
	Fluoride
	Lead
	Mercury
	Nitrate + Nitrite (as N)
	Selenium
	Silver
	Antimony
	Beryllium
	Nickel
	Thallium

Table 2-4 Assessment Monitoring Parameters

Assessment monitoring is only one of the options available to the WDNR if a well exceeds a PAL, ES or ACL. The range of responses which the WDNR may take are:

- no action (only for PAL's at NR 140 wells);
- sample wells or require sampling of wells;
- require a change in the monitoring program, including increased monitoring;
- require an investigation of the extent of groundwater contamination;

- require a revision of the operational procedures at the facility, practice or activity;
- require an alternate method of waste treatment of disposal;
- require prohibition or closure and abandonment of a facility, practice or activity;
- require remedial action to renovate or restore groundwater quality;
- revise rules or criteria on facility design, location or management practices.

CHAPTER 3

DUMPStat ALGORITHMS

3.1 Introduction

In the previous chapter we introduced the current groundwater monitoring regulations as enforced by the Wisconsin Department of Natural Resources (WDNR). In this chapter we consider the potential use of DUMPStat as an approach for evaluating groundwater monitoring data at landfill sites. DUMPStat (Downgradient, Upgradient Monitoring Program Statistics) is a statistical package developed by Gibbons. It has five analysis options:

- upgradient versus downgradient (interwell) prediction limits;
- intrawell prediction limits (PL's) and combined Shewhart-CUSUM;
 control charts (CC's)
- time series plots;
- VOC detection;
- assessment monitoring.

The two intrawell options could potentially replace the indicator PAL. The objective of this research is to evaluate DUMPStat as an alternative to the current methods for calculating the PAL for indicator parameters. The objective of this chapter is to discuss operational issues, to introduce all the options available in DUMPStat, and to take a closer look at the two intrawell options.

3.2 **DUMPStat Operation**

We found the DUMPStat program easy to use. The program runs well both in Windows 3.11 and Windows 95. We had no installation problems. In addition, the menus are easy to understand and use. The program, in its current form, would be difficult to misuse because all important statistical parameters cannot be changed. For example, the user cannot select the significance level for the prediction limit. Also, the statistical parameters for the combined Shewhart -CUSUM control chart are set and cannot be changed.

The WDNR data was easily merged into the DUMPStat database, with only a few problems. The problems we experienced were with nondetect heavy metals. DUMPStat requires that the nondetect data include detection limits. Some data did not include the detection limit and could not be merged. Most indicator parameter data are above the detection limit. So for our study, merging indicator parameters was not a problem. However, merging large data sets (over 10 Mb) can take over an hour. Merging smaller data sets takes only minutes. In order to shorten merging times, we used a database program to isolate the wells and parameters of interest.

After the data have been merged into the database a few more steps need to be taken to run the DUMPStat options. First, the time window for the background data needs to be defined. Second, the minimum number of background samples needs to be selected (typically eight for indicator

parameters). Third, a rare event statistic needs to be selected, either Poisson prediction limits or nonparametric prediction limits. We did not examine parameters that would require use of rare event statistics. Fourth, we decided to identify historical trends, an optional step. Fifth, we isolated the wells and constituents of interest. If these steps are followed most DUMPStat options can be performed.

3.3 **DUMPStat Options**

DUMPStat has five options for analysis. After merging the data and performing the five steps listed in the previous section these options can be used. The first option is the upgradient versus downgradient prediction limits. The second option includes the intrawell prediction limits (PL's) and combined Shewhart-CUSUM control charts (CC's). This option is the main focus of this study and is discussed below. The third option is the time series plots. This option can be used to plot the concentration versus time for the selected parameters. The fourth option, VOC detection, lists all the selected volatile organic compounds (VOC) above the detection limit. The fifth option is assessment monitoring. For each parameter for which a health or welfare standard is set (by the user), DUMPStat makes a time series plot, determines if any trends exist, and performs a t-test for the last four independent samples. Section NR 140.14 allows the t-test or another valid statistical analysis for the data being considered for health and welfare standards.

3.4 Intrawell Methods

As mentioned previously, the intrawell methods in DUMPStat could potentially replace the use of PAL's for detecting changes in values for indicator parameters. DUMPStat has two methods for intrawell testing: prediction limits (PL's) and the combined Shewhart-CUSUM control charts (CC's).

3.4.1 Prediction Limits

When the prediction limit is used, DUMPStat first removes outliers from the background data using Dixon's test. Next, DUMPStat tests for increasing trend in the background data using Sen's test. Then the prediction limit is calculated as:

$$PL = \overline{X} + s \cdot t_{\left[\frac{1-\alpha,n-1}{n}\right]} \cdot \sqrt{1 + \frac{1}{n}}$$
 (3.1)

where \bar{x} is the sample mean, s is the sample standard deviation, n is the number of background samples, and $t_{\left[\frac{1-\alpha,n-1}{2}\right]}$ is the t-statistic based on the 1- α confidence level (or α significance level) and n-1 degrees of freedom.

The significance level, α , is the minimum of: 0.01 or

$$\alpha = \sqrt{1 - 0.95^{\frac{1}{k}}} \,. \tag{3.2}$$

where k is the number of future comparisons. From equation 3.2 we can determine that when k is 100, α is the minimum of 0.01 and 0.0226. k will always be less than 100; consequently α will always be 0.01 for intrawell analysis. Further, equation 3.2 is based on one verification resample. We were not able to verify historical results for our analysis. Consequently, DUMPStat only indicated a violation after two consecutive exceedances.

The PL is a logical alternative to the PAL because a limit is set based on the background mean and standard deviation. Similar to the indicator PAL, the PL is compared to future data points. Also, prediction limits and indicator PAL's are better suited for detecting sudden increases in concentration. Unlike the indicator PAL, DUMPStat does not indicate a violation unless the resample or the next sampling round (depending on the verification resampling plan) also violates the PL. The PL procedure outlined here as well as the tests for historical trend and outliers are illustrated in detail with examples in appendix A.

A comparison can be made between the indicator PAL and PL without calculating either. Most of the time, the PL as implemented by DUMPStat will be more conservative and will produce more violations than the indicator PAL. This is true for two reasons. First, the PL does not allow for a minimum increase like the indicator PAL. This minimum increase allows an additional benefit to the landfill owner for indicator parameters with a low variance. Second, the number of standard deviations used to compute the PAL will generally be greater than that used to compute the PL. Consider the second point. The PAL is always

three standard deviations above the mean. However, for the PL the number of standard deviations is $t_{[1-a,n-1]} \cdot \sqrt{1+\frac{1}{n}}$. The dependence on n, the number of background samples, is illustrated in Figure 3-1 for a 99 percent confidence level. From this figure, we can see with eight background samples the PL is 3.18 standard deviations above the mean which is almost the same as the indicator PAL. Further, with ten background samples the PL is 2.96 standard deviations above the mean which is essentially the same as the indicator PAL. Looking at Figure 3-1 we can conclude that the PL is more conservative than the PAL for ten or more background samples. From eight to ten background samples the PL is essentially the same as the indicator PAL. Since eight is the minimum number of background samples for indicator parameters the PL will at worst be 0.18 standard deviations above the PAL.

All in all, the PL, as implemented by DUMPStat, is more conservative than the indicator PAL. The PL will cause more violations than the indicator PAL. This could result in additional monitoring or possibly remedial action.

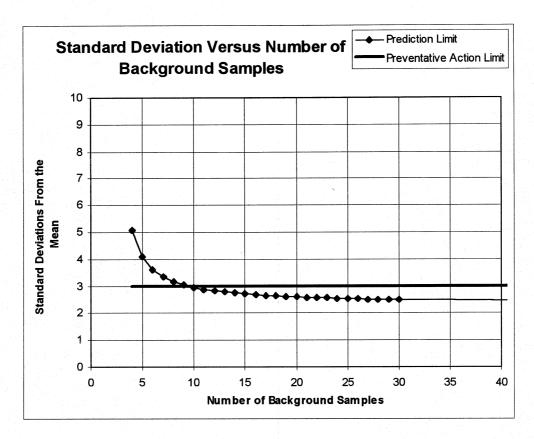


Figure 3-1 Comparison of PAL and PL Based on Number of Background Samples

3.4.2 Combined Shewhart-CUSUM Control Charts

The second intrawell method that could replace the indicator PAL is the combined Shewhart-CUSUM control chart (CC). As its name suggests, the combined Shewhart-CUSUM control chart combines the Shewhart control chart with the CUSUM control chart. The CC can pick up sudden changes as well as long term upward trends in water quality.

The Shewhart control chart is often used in industrial practice. Consider a bolt manufactured with a specified diameter. The diameter of the bolt is assumed to be a random variable. A mean and standard deviation for the bolt diameter can be calculated. The original Shewhart control chart is "out of control" if the bolt diameter is three standard deviations above or below the mean. If the control chart is "out of control" then the problem is considered deterministic. Theoretically, a deterministic problem can be fixed by changing the machinery or the operator. This is very similar to the current indicator PAL. In fact, the WDNR is currently using a one sided Shewhart control chart since the indicator PAL is set three standard deviations above the mean.

The Shewhart test used in DUMPStat is a variation of the original Shewhart control chart. It is a one sided test and the control chart is "out of control" if the sample exceeds the mean by 4.5 standard deviations for eleven or fewer background samples or by 4.0 standard deviations for twelve or more background samples.

The cumulative sum (CUSUM) portion of the control chart is much different. While the Shewhart portion of the control chart detects sudden increases in concentration the CUSUM portion detects steady or slow increases in concentration. The CUSUM, S_i is calculated as

$$S_{i} = \max[0, (z_{i} - k) + S_{i-1}]$$
 (3.3)

where z_i is defined as

$$z_{i} = \frac{x_{i} - \overline{x}}{s}, \qquad (3.4)$$

 \overline{x} is the sample mean, s is the sample standard deviation, and x_i is the concentration of interest. Also, k is a parameter related to the displacement which should be detected quickly (one for eleven or fewer background samples and 0.75 for twelve or more background samples) and S_0 =0. The CUSUM portion of the control chart is considered "out of control" if S_i is greater than 4.5 for eleven or fewer samples and 4.0 for twelve or more samples.

Like the prediction limits, the DUMPStat CC does not indicate a violation unless it is confirmed by a resample or the next round of sampling. If a resample is taken it should replace the original value in order to provide an unbiased confirmation of the exceedance for the CC. A complete explanation of the CC is included in Appendix A with a sample calculation.

3.6 Chapter Summary

All in all we found the DUMPStat program easy to install and use. The program has several simple and useful options including: assessment monitoring, time series plotting and VOC detection. DUMPStat also has interwell prediction limits which do not fit well into the WDNR system since

intrawell analysis is the preferred method. DUMPStat has two intrawell options: prediction limits and the combined Shewhart-CUSUM control charts. Of these two, the prediction limits are more conservative than the indicator PAL for ten or more background samples. Further, they are essentially the same as the PAL for eight or nine background samples. The combined Shewhart-CUSUM control chart on the other hand is more difficult to evaluate since a limit can only be set for the next sampling round while the prediction limit sets a limit for the next k sampling rounds. In the next chapter we apply these intrawell methods to data from a Wisconsin landfill.

CHAPTER 4

PORTAGE COUNTY CASE STUDY

4.1 Introduction

As previously mentioned, the indicator PAL could potentially be replaced by either of the two DUMPStat intrawell options: prediction limits (PL's) or the combined Shewhart-CUSUM control charts (CC's). In order to examine and compare the effectiveness of the PAL's, PL's and CC's we applied all three methods to the Portage County landfill (license number 2966), a landfill which is not apparently affecting groundwater quality.

We examined both upgradient and downgradient wells using the indicator parameters: alkalinity, hardness and specific conductance. Using box plots, time series plots and a site map we classified each well into one of four categories: up/sidegradient unimpacted; downgradient impact possible; downgradient impact suspected; and downgradient unimpacted. Next for each parameter and each well, we applied the indicator PAL, PL and CC methods. The effectiveness of each method was evaluated using the downgradient impacted wells, while the false positive rate was estimated based on the unimpacted upgradient and sidegradient wells.

Based on this case study we found that the CC's were more conservative than the PL's, and both methods were more conservative than the indicator PAL. That is, the CC and PL would put the landfill in violation much more often than the indicator PAL. This result confirms the arguments made in the third chapter.

In addition, we found an unexpected number of violations at upgradient wells. We used these violations to estimate the false positive rate and we found it to be very high for some unimpacted wells. The high number of violations and high false positive rates led us to question the assumptions of these statistical tests. This issue is considered in Chapter 5.

4.2 Portage County Landfill Characteristics

The Portage County landfill is located approximately in the center of the state of Wisconsin. As stated previously, the landfill has a number of upgradient and sidegradient wells which are not likely to have been impacted by the landfill. The landfill also has a number of downgradient wells. The location of all wells and the direction of groundwater flow are shown on Figure 4-1.

Fifteen of the wells have continuous data from 1983 to the present. Each well has approximately 51 data points for each indicator parameter. Of these fifteen wells, three can be clearly classified as upgradient (W-10, W-11 and W-12). Two of the wells are sidegradient but appear to be beyond the influence of the landfill (W-9 and W-9p). The site has ten clearly downgradient wells (W-17, W-18, W-20, W-20p, W-21, W-21p, W-22, W-22p, W-23 and W-23p). The "p" designates a deep well (piezometer), screened below the water table. These wells are screened in sandy glacial till (in a sand fraction that is coarser than above). The wells without the "p" designation are screened at (or near) the water table which is in fine to very fine sandy glacial till.

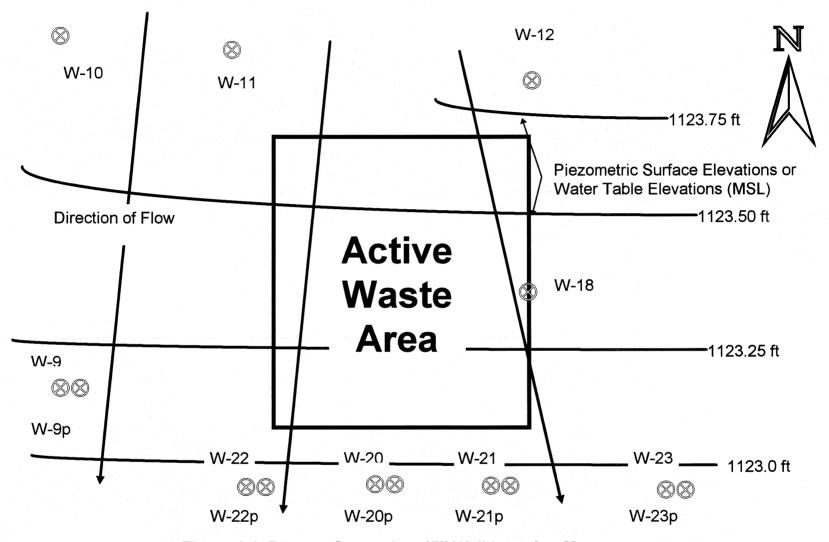


Figure 4-1 Portage County Landfill Well Location Map

4.3 Box Plots and Time Series Plots

As a preliminary step in the evaluation of the landfill, we prepared box plots and time series plots for all wells. These plots were discussed and recommended in the two previous WDNR studies (Goodman, 1987 and Fisher, 1989). These studies suggested that a high median or interquartile range (IQR) may indicate contamination. Similarly, multiple outliers or extreme values may indicate contamination. Box plots for all the wells for a single parameter should be plotted simultaneously. Aside from the box plots, time series plots can be visually inspected for increasing trends. Multiple time series can be plotted on one graph; however, this should be limited to five or fewer wells and one parameter per graph. In addition, at least one upgradient or sidegradient well should be plotted on each graph for comparison.

Figure 4-2, 4-3 and 4-4 show the box plots from the fifteen wells for alkalinity, hardness and specific conductance respectively. The box plots were created by Statistica (1997) and the box plot construction is indicated in Figure 4-5. Time series plots for alkalinity (Figures 4-6, 4-7 and 4-8), hardness (Figures 4-9, 4-10 and 4-11) and specific conductance (Figures 4-12, 4-13 and 4-14) also are given. These box plots and time series plots are based on quarterly samples for each constituent at each well (approximately 51 samples for each parameter). By visually inspecting the box plots, time series plots and well location map, we made some preliminary classifications. Based on these

classifications we will compare the effectiveness of the indicator PAL to the PL and CC. The preliminary classifications are:

- 1. Up/sidegradient unimpacted: W-10, W-11, W-12, W-9 and W-9p
- 2. Downgradient impact possible: W-17, W-20p, W-21p, W-22 and W-22p
- 3. Downgradient impact suspected: W-20, W-21 and W-23p
- 4. Downgradient impacted: W-18 and W-23

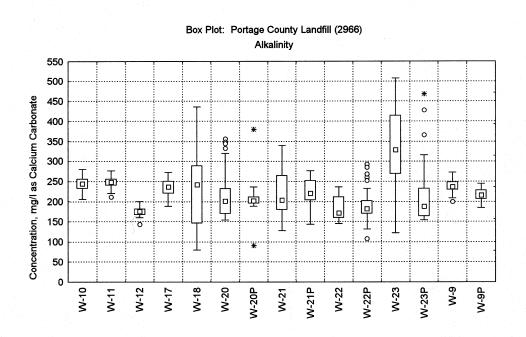


Figure 4-2 Box Plots for Alkalinity

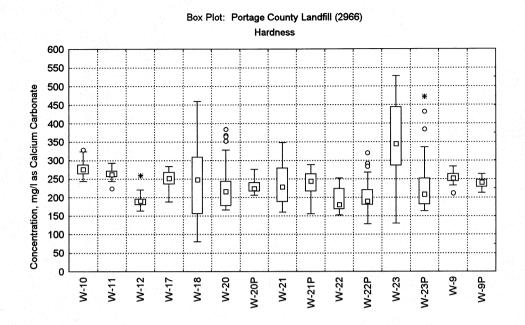


Figure 4-3 Box Plots for Hardness

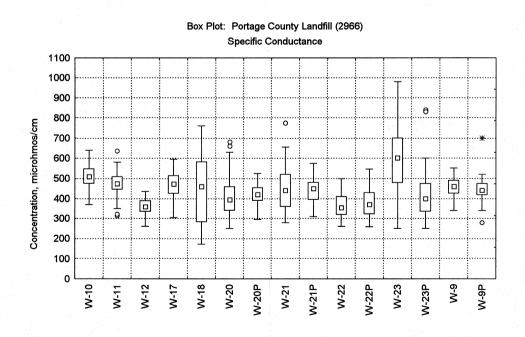


Figure 4-4 Box Plots for Specific Conductance

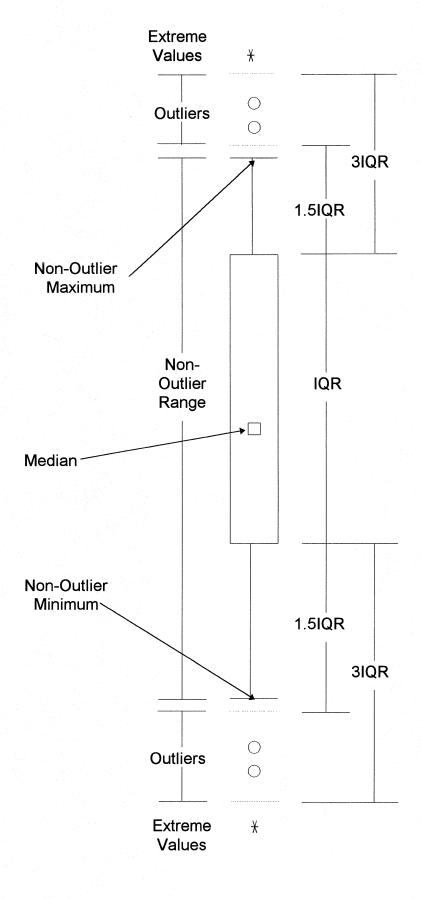


Figure 4-5 Box Plot Construction by Statistica

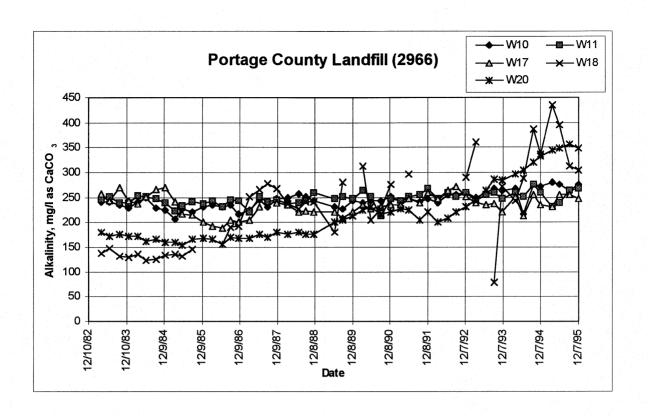


Figure 4-6 Time Series Plots Alkalinity

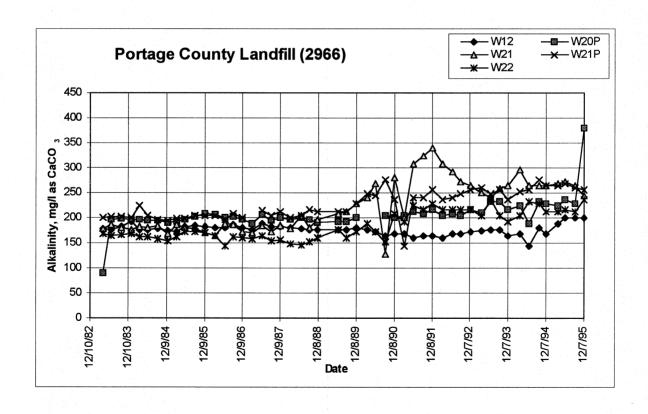
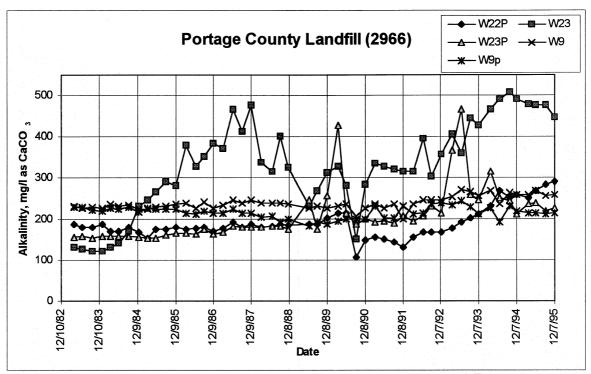


Figure 4-7 Time Series Plots Alkalinity



Note that y-axis is different from Figures 4-6 and 4-7

Figure 4-8 Time Series Plots Alkalinity

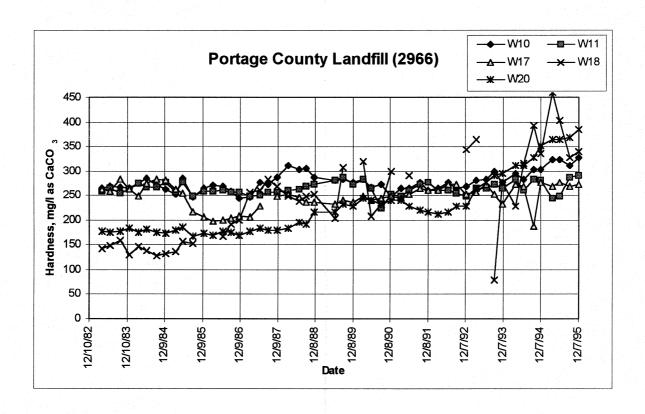


Figure 4-9 Time Series Plots Hardness

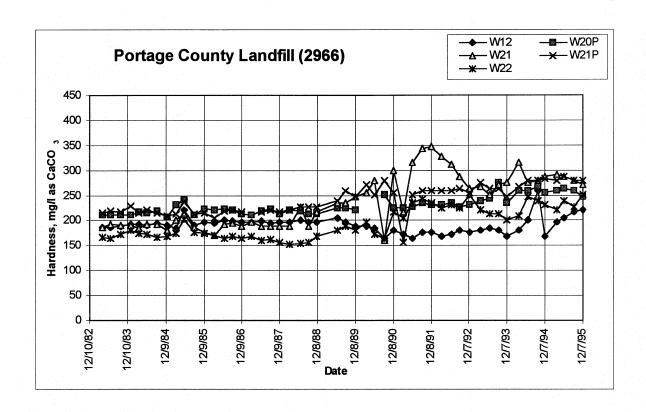
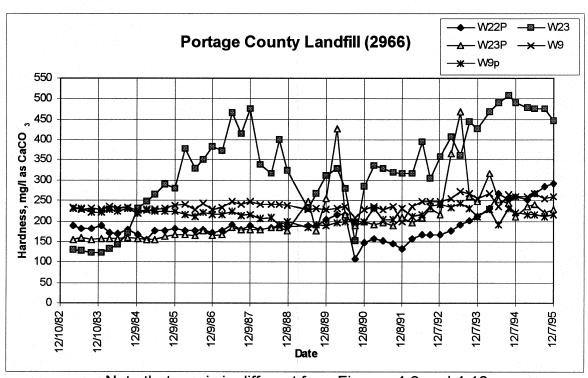


Figure 4-10 Time Series Plots Hardness



Note that y-axis is different from Figures 4-9 and 4-10

Figure 4-11 Time Series Plots Hardness

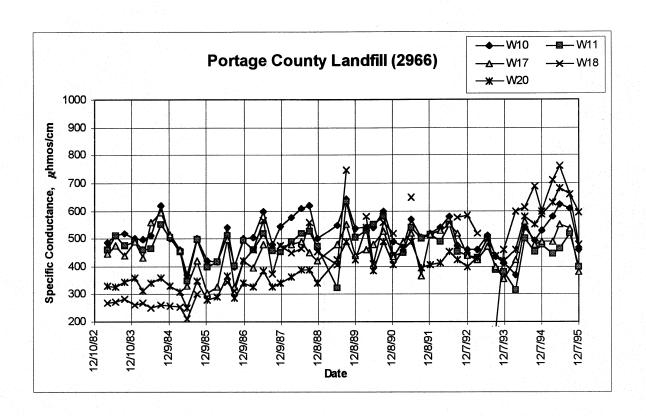


Figure 4-12 Time Series Plots Specific Conductance

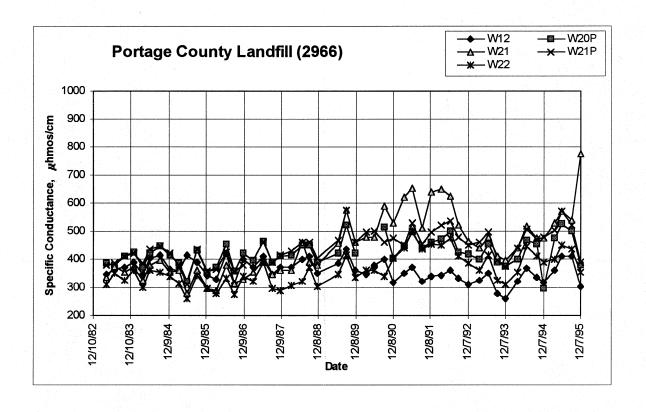


Figure 4-13 Time Series Plots Specific Conductance

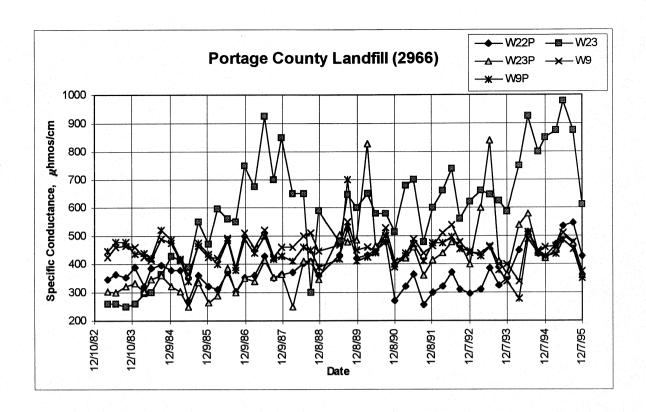


Figure 4-14 Time Series Plots Specific Conductance

4.4 Intrawell Analysis All Wells

Box plots and time series plots, as discussed in the previous section, can be an important tool for groundwater monitoring. In fact, the use of these visual aids may be better than the use of any statistic. However, these visual aids must be used with appropriate statistics governed by law. Currently, the appropriate statistic for alkalinity, hardness and specific conductance is the indicator PAL.

The DUMPStat alternatives to the indicator PAL are the PL and CC.

First let us compare the PL to the indicator PAL. Figures 4-15, 4-16 and 4-17 show the indicator PAL and PL for alkalinity, hardness and specific conductance, respectively, based on eight background samples. From these figures we can see the PL is lower for all but a few of wells for each parameter. The PL is higher than the indicator PAL for W-23 for all parameters and for W-20p for alkalinity, because the standard deviation for each is greater than M/3 where M is the minimum increase. Recall (figure 2-1) that the minimum increase for alkalinity, hardness and specific conductance is 100 mg/l, 100 mg/l and 200 µhmos/cm, respectively. So if the variance is above 33.3 mg/l, 33.3 mg/l or 66.7 µhmos/cm for alkalinity, hardness and specific conductance, respectively, the PL will be 3.18 standard deviations above the mean while the indicator PAL will be 3.00 standard

deviations above the mean for eight background samples. Otherwise the minimum increase above the mean is set as the indicator PAL which is well above the PL.

We also calculated the PL and indicator PAL based on 25 background samples. As expected, the PL is always less than the indicator PAL. The indicator PAL and PL for 25 background samples are given in Figures 4-18, 4-19 and 4-20 for alkalinity, hardness and specific conductance. For 25 background samples, the PL for parameters with a high variance (over M/3) will be close to the indicator PAL while those with a low variance tend to be much lower than the indicator PAL. Using 25 background samples also produced lower PL than using eight background samples. This result is expected and explained in the last chapter.

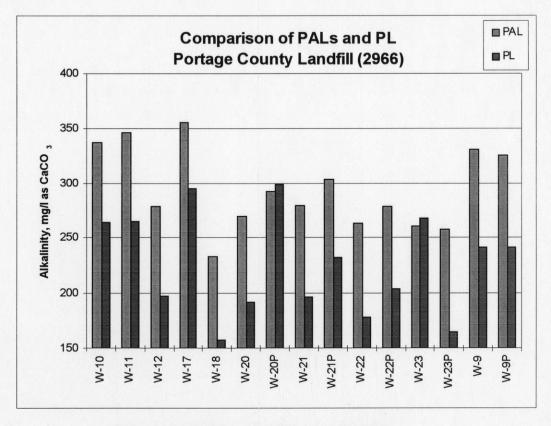


Figure 4-15 PAL and PL for Alkalinity (8 Background Samples)

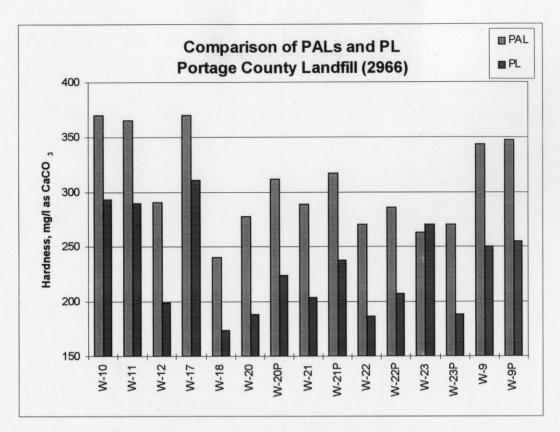


Figure 4-16 PAL and PL for Hardness

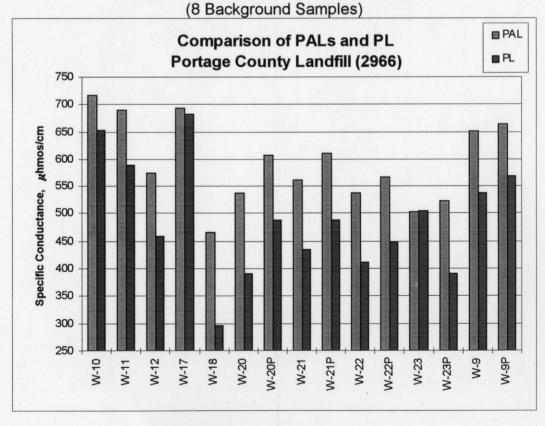


Figure 4-17 PAL and PL for Specific Conductance (8 Background Samples)

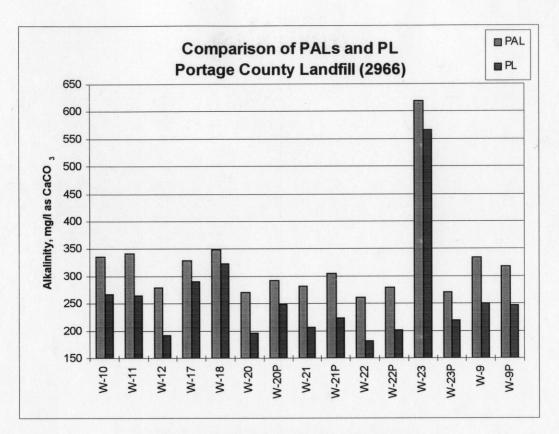


Figure 4-18 PAL and PL for Alkalinity

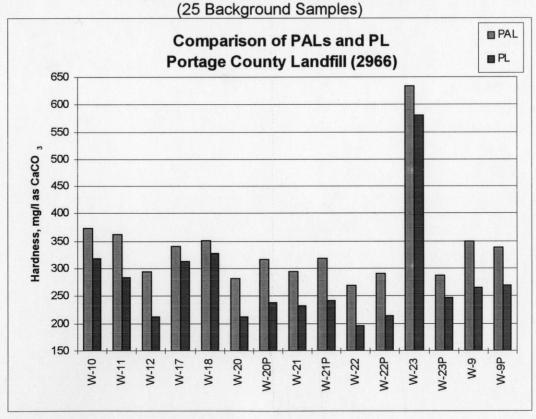


Figure 4-19 PAL and PL for Hardness (25 Background Samples)

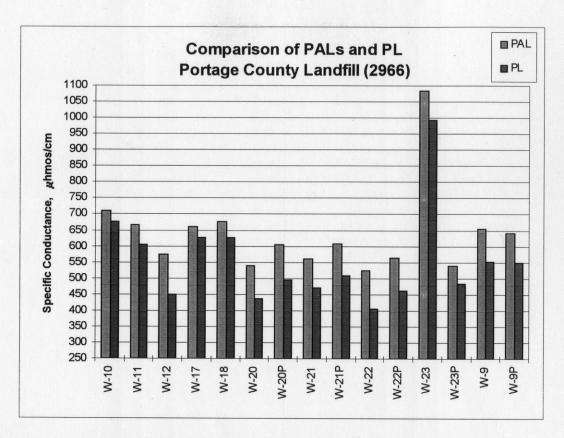


Figure 4-20 PAL and PL for Specific Conductance (25 Background Samples)

To compare the CC to the indicator PAL, we calculated the number of data points in violation for alkalinity, hardness and specific conductance for each well. This procedure was conducted using both eight and 25 background samples. This analysis was also performed for the PL for comparison. The results from this analysis are listed in Table 4-1 and Table 4-2 for eight and 25 background samples. Figures 4-21 and 4-22 present the total number of violations for each of the parameters for all the wells combined.

From this analysis, we found that the greatest number of violations sitewide (for alkalinity, hardness and specific conductance) was 89 for the PAL

with eight background samples and 27 for 25 background samples. The greatest number of violations sitewide for the PL was 308 for eight background samples and 172 violations for 25 background samples. The CC had the highest number violations sitewide with 390 violations for eight background samples and 198 violations for 25 background samples. In other words on a sitewide basis, the PL produced over three times as many violations as the indicator PAL for eight background samples and over six times as many violations based on 25 background samples. Similarly, the CC produced four and seven times as many violations as the indicator PAL for eight and 25 background samples respectively.

	Alkalinity			Hardness			Specific Conductance		
	PAL	PL	CC	PAL	PL	CC	PAL	PL	CC
W-10	0	9	12	0	11	25	0	0	0
W-11	0	3	12	0	1	0	0	1	2
W-12	0	3	2	0	9	28	0	0	0
W-17	0	0	0	0	0	0	0	0	0
W-18	22	28	37	21	28	37	19	31	38
W-20	10	27	26	10	30	30	7	26	28
W-20P	1	1	1	0	22	43	0	6	21
W-21	7	27	28	8	29	31	9	26	27
W-21P	0	23	24	0	27	27	0	12	26
W-22	0	22	23	0	24	27	0	11	20
W-22P	2	12	11	2	16	17	0	8	15
W-23	40	39	42	41	40	42	37	37	41
W-23P	5	38	39	7	36	35	5	27	30
W-9	0	17	28	0	28	41	0	2	0
W-9P	0	1	2	0	7	9	1	1	1

Table 4-1 Number of Data Points in Violation (Based on Eight Background Samples)

	Alkalinity			Hardness			Specific Conductance		
	PAL	PL	СС	PAL	PL	CC	PAL	PL	CC
W-10	0	7	11	0	3	4	0	0	0
W-11	0	3	15	0	3	4	0	1	1
W-12	0	4	3	0	4	3	0	0	0
W-17	0	0	0	0	0	0	0	0	0
W-18	4	5	27	4	7	26	4	6	26
W-20	10	27	26	10	27	27	7	16	26
W-20P	1	1	5	0	13	20	0	6	20
W-21	6	25	26	7	25	26	9	20	26
W-21P	0	25	25	0	25	26	0	7	26
W-22	0	22	22	0	23	21	0	14	21
W-22P	2	13	11	2	15	14	0	6	15
W-23	0	0	6	0	0	6	0	0	4
W-23P	4	15	26	5	15	27	5	9	26
W-9	0	11	12	0	12	12	0	0	0
W-9P	0	0	2	0	0	0	1	1	11

Table 4-2 Number of Data Points in Violation (Based on 25 Background Samples)

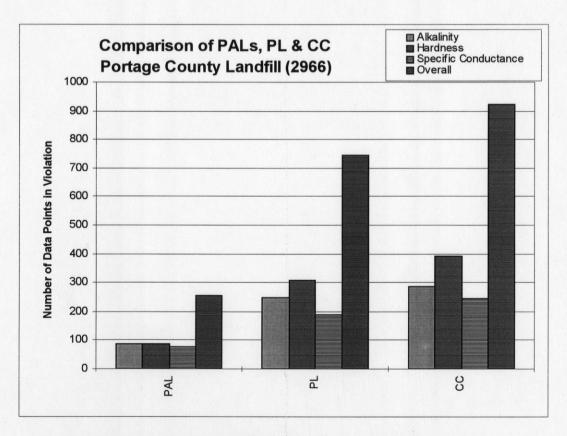


Figure 4-21 Number of Data Points in Violation Sitewide (Based on Eight Background Samples)

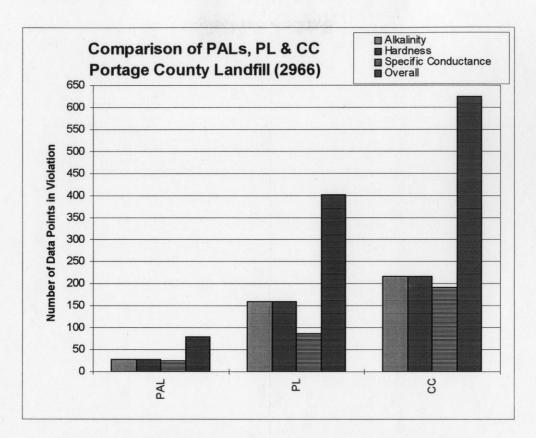


Figure 4-22 Number of Data Points in Violation Sitewide (Based on Approximately 25 Background Samples)

4.5 Upgradient Wells

As discussed in the previous section, the PL and CC are more conservative than the indicator PAL. Therefore, DUMPStat will produce more violations when compared to the indicator PAL. Because of the increased number of violations, DUMPStat will pick up an impact sooner than the indicator PAL. However, some of these violations are false. In order to evaluate this effect, five wells were considered. All were either upgradient or sidegradient and beyond the influence of the landfill. We estimated the false positive rate by dividing the number of violations by the total number of data points for each parameter for these wells. We performed this analysis based on both eight and 25 background samples.

The results of this analysis are given in Figures 4-23 and 4-24 for alkalinity, hardness and specific conductance. From these plots, one can see that hardness and alkalinity have a higher false positive rate than specific conductance for both eight and 25 background samples. Specifically for the two DUMPStat algorithms, specific conductance had less than a five percent false positive rate for both eight and 25 background samples, while alkalinity and hardness had over a five percent false rate. The worst case was a 40 percent false positive rate for hardness using eight background samples and CC.

Both of the DUMPStat algorithms had higher false positive rates than the indicator PAL. However, CC had the highest false positive rate. With eight

background samples, the overall false positive rate was 0.1, 12 and 21 percent for the PAL, PL and CC, respectively. Similarly with approximately 25 background samples, the overall false positive rates improved to 0.1, 6, and 9 percent. We would expect the false positive rates to decrease with more background samples since the mean and standard deviation can be better estimated for upgradient wells. This result is not true for downgradient wells because they may be impacted by the landfill. Hence, the background data for downgradient wells should be selected with care.

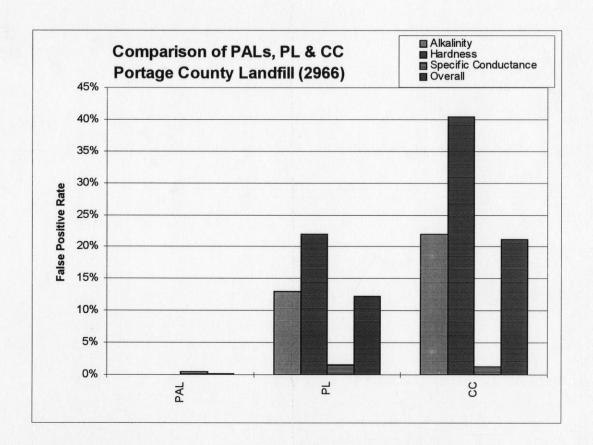


Figure 4-23 False Positive Rate Based on Five Upgradient Wells (Based on Eight Background Samples)

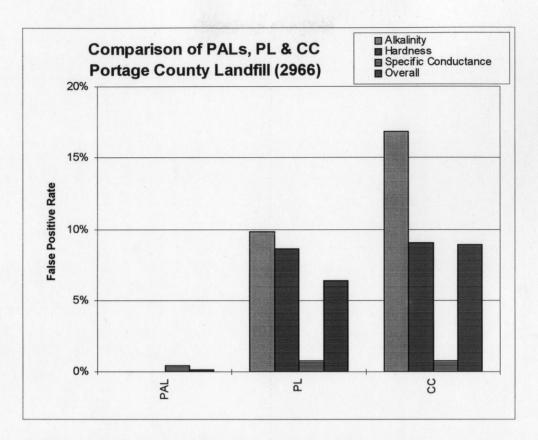


Figure 4-24 False Positive Rate Based on Five Upgradient Wells (Based on 25 Background Samples)

4.6 Downgradient Wells

The false positive rates for the upgradient wells appear to be high. In addition to determining the false positive rate, we are interested in evaluating the effectiveness of the test. We did this evaluation by examining two downgradient wells which were preliminarily determined to be impacted.

Consider W-18 and W-23 which we determined to be impacted based on box plots and time series plots for alkalinity, hardness and specific conductance.

As noted in Table 4-1, one can see, regardless of the method, that W-23 and W-18 have high numbers of violations based on eight background samples. W-23 has approximately 40 violations (see Table 4-1) for alkalinity, hardness and specific conductance regardless of the test, based on eight background samples. Each method is equally effective for detecting the impact on W-23 since each method indicates 4/5 of the data points are in violation. For W-18 (8 background samples) the PAL, PL and CC have approximately 21, 29, and 37 violations for each parameter. Although the number of violations are different, each test is effective for determining impact.

Based on eight background samples each test was effective for determining impact. However, the same cannot be said for using 25 background samples. Based on 25 background samples, the PL and indicator PAL do not indicate impact for these wells and parameters. Alternatively, the CC indicates a substantial impact for W-18 but not for W-23. This result indicates that background data should be selected with great care. The 25 background samples likely include impacted data which increases the mean and standard deviation and decreases the likelihood of impact detection for all tests.

4.7 Chapter Summary

We have shown that the intrawell DUMPStat algorithms are more conservative than the indicator PAL. In addition, it is clear that the CC will produce more violations than the PL. Each of these methods is effective for determining impact. Of the violations, we noticed some upgradient wells with a substantial number of violations. Consequently, we estimated the false positive rate for the each parameter and statistical test. The false positive rate at these wells for the PL and CC was very high for alkalinity and hardness when compared to the indicator PAL. The number of violations at upgradient wells leads us to question the underlying assumptions of these statistical tests, which will be examined in the next chapter. Regardless of the statistical test or parameter, the selection of background data appears to a very important factor. If too much data is included in the background, all the statistical tests will fail to detect an impact. In the next chapter, we will review the underlying assumptions of these statistical tests.

CHAPTER 5

ASSUMPTIONS FOR INTRAWELL ANALYSIS

5.1 Introduction

In the previous chapters we examined three statistical tests used for intrawell analysis of indicator parameters including: indicator preventive action limits (PAL), prediction limits (PL) and combined Shewhart-CUSUM control charts (CC). Currently, the PAL is the statistical test mandated by the Wisconsin Administrative Code and enforced by the Wisconsin Department of Natural Resources (WDNR) for indicator parameters. Prediction limits and combined Shewhart-CUSUM control charts are intrawell tests used in the computer program DUMPStat. These tests are potential replacements for the indicator PAL.

In the last chapter we applied the prediction limits and combined

Shewhart-CUSUM control charts on data from a Wisconsin landfill and found a
large number of false positives. For example, using 25 background samples we
found the false positive rate to be over fifteen percent for alkalinity at the

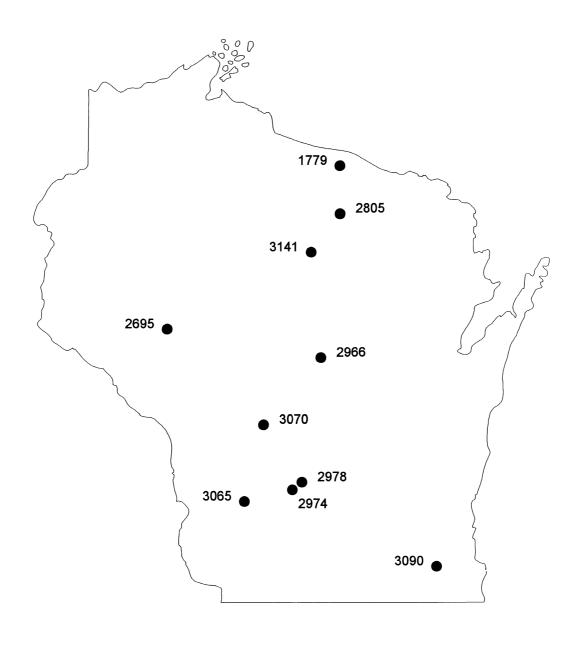
Portage County Landfill. In this chapter we investigate when this high incidence
of false positives is due to violations of the assumptions; in particular, we
explore distributional assumptions and the assumptions of stationarity and
independence. These assumptions are investigated using alkalinity, hardness

and specific conductance data from 26 upgradient wells at ten Wisconsin landfills.

5.2 Landfill Selection

The locations of the landfills we considered are indicated in Figure 5-1. These landfills cover a majority of the state and are located in a variety of geologic formations. Table 5-1 gives information on each landfill. Most of the landfills accept municipal solid waste (MSW), and three of the landfills accept industrial waste including papermill sludge and foundry waste. The size of these landfills varies from the small Lincoln Landfill (license number 1779), which has a capacity of 0.05 million cubic yards, to the large Troy area landfill (license number 3090), which has a capacity of 3.9 million cubic yards.

Based on site maps provided by the WDNR, we identified 26 wells that were unlikely to have been impacted by these ten landfills. Most of the wells are clearly upgradient. A few are sidegradient wells but are clearly beyond the influence of the landfill. Each of the wells had at least fifteen data points for each of the indicator parameters of interest. The selected wells are listed in Table 5-2.



License	Landfill	License	Landfill
Number	Name	Number	Name
1779	Lincoln	2978	Sauk County
2695	Pope & Talbot	3065	City of Richland Center
2805	Oneida County	3070	Juneau County
2966	Portage County	3090	Troy Area
2974	Grede Foundries	3141	Lincoln County

Figure 5-1 Name and Location of Landfills Studied

5-3

Table 5-1 Landfill Characteristics for Sites Studied

Landfill	Facility Design	Waste	Design Volume	Site Size	Site	Year	Number of
(License)		Received	(Millions of Cubic Yards)	(Acres)	Life	Sampling Began	Upgradient Wells Analyzed
Lincoln (1779)	Natural Attenuation	MSW	0.05	20	1984- 1990	1984	2
Pope & Talbot (2695)	Partial Clay Lined Leachate Collection	Papermill Sludge	1.2	19	1976- present	1976	3
Oneida County (2805)	Partial Clay Lined Leachate Collection	MSW Papermill Sludge	1.1	16.3	1979- present	1979	2
Portage County (2966)	Clay Lined Leachate Collection	MSW	1.37	18.6	1984- Present	1983	5
Grede Foundries (2974)	Partial Clay Lined Leachate Collection	Foundry Waste	0.966	19.5	1981- present	1981	2
Sauk County (2978)	Clay Lined Leachate Collection	MSW	1.28	20	1983- Present	1978	3
City of Richland Center (3065)	Clay Lined Leachate Collection	MSW	0.34	7.3	1986- 1993	1986	3
Juneau County (3070)	Clay Lined Leachate Collection	MSW	0.42	8	1987- present	1987	1
Troy Area (3090)	Clay Lined	MSW	3.9	26	1987- 1996	1987	3
Lincoln County (3141)	Clay Lined Leachate Collection	MSW	0.825	14.6	1987- present	1987	2

License #	Landfill	Common Well #	DNR Well#	Location
1779	Lincoln	MW-1	1	Up
1779	Lincoln	MW-4	104	Up
2695	Pope & Talbot	W-13	11	Up
2695	Pope & Talbot	W-18	18	Up
2695	Pope & Talbot	W-6	3	Up
2805	Oneida County	MW-1	11	Up
2805	Oneida County	MW-126	126	Up
2966	Portage County	W-10	14	Up
2966	Portage County	W-11	16	Up
2966	Portage County	W-12	17	Up
2966	Portage County	W-9	12	Side
2966	Portage County	W-9P	13	Side
2974	Grede Foundries	B-3	103	Up
2974	Grede Foundries	B-5	105	Up/Side
2978	Sauk County	W-30	101	Up
2978	Sauk County	W-30A	102	Up
2978	Sauk County	W-31	103	Up
3065	City of Richland Center	MW-6	106	Up
3065	City of Richland Center	MW-7	107	Side
3065	City of Richland Center	MW-7P	108	Side
3070	Juneau County	OW-5	1	Up
3090	Troy Area	B-1	201	Up
3090	Troy Area	B-1B	203	Up
3090	Troy Area	B-2	204	Up
3141	Lincoln County	M-4	706	Up
3141	Lincoln County	M -9	710	Up/Side

Table 5-2 Selected Wells

5.3 Stationarity, Independence and Trend

Most statistical tests, such as those based on the combined Shewhart-CUSUM control charts or predictions limits are predicated on the assumption that the data are taken from a random sample. Herein random means that the data are independent and identically distributed. In this section, we will determine whether the upgradient alkalinity, hardness and specific conductance data represent random samples. In particular, we will test for two kinds of nonrandomness: nonstationarity; and serial correlation.

For data collected in time, nonstationarity means that the distribution changes in time. Nonstationarity can be due to seasonality of the mean and/or variance or to long-term drift of the mean and/or variance. Figure 5-2 is an example of nonstationarity due to periodic variation in the mean. Figure 5-3 is an example of nonstationarity in variance. Figure 5-4 shows data for which both the mean and variance change in time.

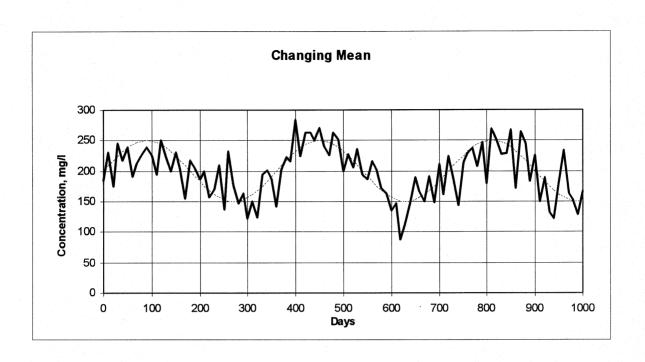


Figure 5-2 Nonstationary Process: Changing Mean

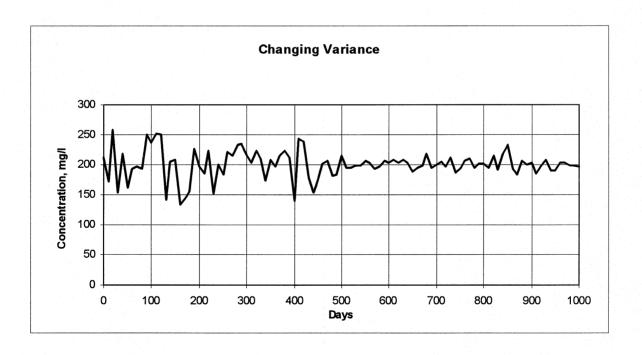


Figure 5-3 Nonstationary Process: Changing Variance

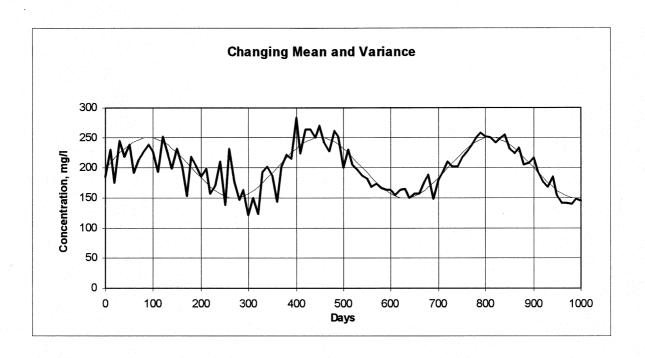


Figure 5-4 Nonstationary Process: Changing Mean and Variance

Figures 5-2 through 5-4 are hypothetical examples of nonstationarity;

Figure 5-5 shows an example of apparent nonstationarity of alkalinity data from two wells at the Portage County landfill. These data show apparent variation in the mean and variance. For example, for W-9 the variance for data prior to June 1990 is 38.5 mg/l and after is 374.6 mg/l. The variance of the data from W-9p also appears to increase, although not as significantly as W-9. These wells are clearly beyond the influence of the landfill; hence the apparent nonstationary behavior is not due to contamination by the landfill. Determining the cause of this apparent nonstationarity was beyond the scope of this study. One possibility is consistent measurement error, due to changes in sampling and

measurement protocols. Another possibility is that the apparent nonstationarity is due to variation in the flow field at the site.

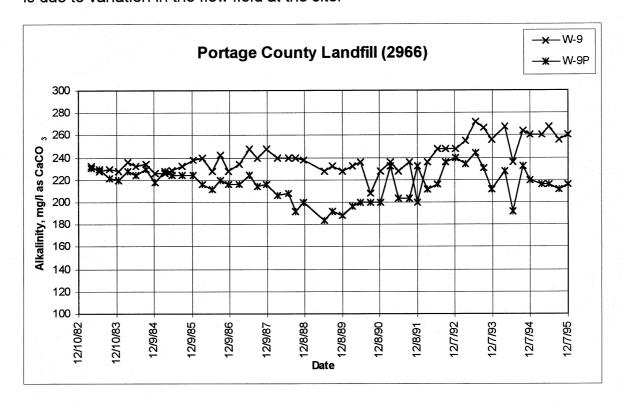


Figure 5-5 Nonstationary Process: Changing Variance

Inspection of all of the time series plots of the water quality data from 26 upgradient wells did not reveal any obvious seasonal behavior. Goodman (1987) and Montgomery et. al. (1987) also found little significant seasonality in groundwater quality data. Consequently, we did not conduct any formal statistical tests for seasonality. Instead, we conducted statistical tests for serial correlation and for linear trends, both of which are causes of nonrandomness. In short datasets, such as those analyzed here, correlation and trend can be confounding. Strictly speaking, serial correlation should only be estimated from

stationary data, as the presence of a trend can falsely indicate serial correlation. Similarly, serial correlation can indicate a trend which does not exist. We chose to use tests of both serial correlation and trend to detect nonrandomness. We also tested the data for normality and lognormality, since both prediction limits and combined Shewhart-CUSUM control charts are typically predicated on the assumption of normality or lognormality.

Serial correlation was evaluated using the lag-k autocorrelation coefficient, the correlation coefficient for data points separated by k sampling intervals. This coefficient varies from -1 to +1. For the lag-1 autocorrelation, a positive coefficient indicates high values tend to be followed by high values and low values tend to be followed by low values in the time series. For a negative lag-1 autocorrelation coefficient, the opposite is true.

For our study, we calculated the autocorrelation function coefficient for up to fifteen lags (k=15) depending on the length of the data sets. This analysis was performed for 26 upgradient wells for alkalinity, hardness and specific conductance at ten landfills. We used a five percent significance level for the two sided test. It should be noted that the test procedure offered in Statistica (1997) is slightly different from the procedure used by Goodman (1987). Appendix C outlines the procedure for calculating any lag-k autocorrelation coefficient and it includes a sample calculation.

Before reviewing our results, it is interesting to look at graphical representation of the data. Specifically, a plot for the concentration, C_{i-1}, versus the next concentration, C_i, provides a visual indication of the strength of the lag-1 Autocorrelation coefficient. For perfect positive correlation the plot would be linear with a positive slope. For strong correlation, the plot would show a clustering of the data points with a positive slope. Figures 5-6 and 5-7 show the results for specific conductance for two wells. These two wells show no clustering and no clear slope. The lag-1 autocorrelation coefficient for the data are 0.005 and -0.053 respectively; neither is significant at the 5% level. In contrast, strong lag-1 autocorrelations are shown in Figures 5-8 and 5-9. The positive slopes are clear as well as the clustering of both high values and low values. Lastly, figures 5-10 and 5-11 show strong lag-1 autocorrelations but the high values are more disperse than the low values.

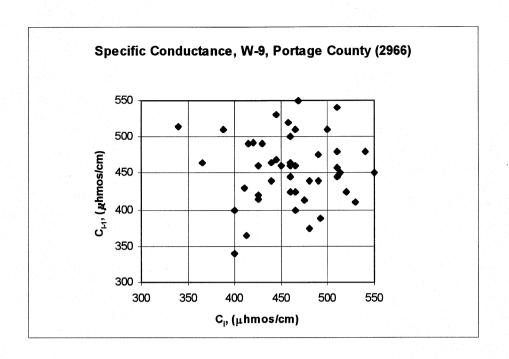


Figure 5-6 No Significant Serial Correlation $(r_1=0.005, p=0.97)$

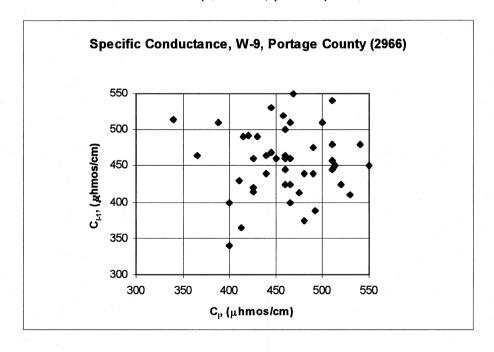


Figure 5-7 No Significant Serial Correlation $(r_1=-0.053, p=0.69)$

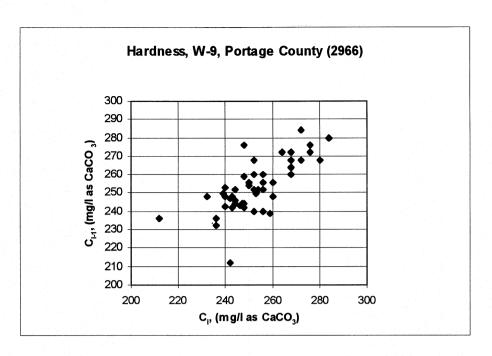


Figure 5-8 Significant Positive Serial Correlation $(r_1=0.716, p=0.000)$

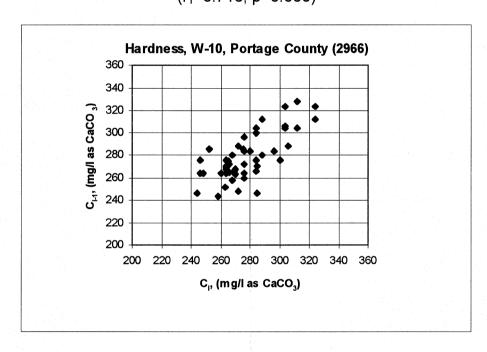


Figure 5-9 Significant Positive Serial Correlation $(r_1=0.676, p=0.000)$

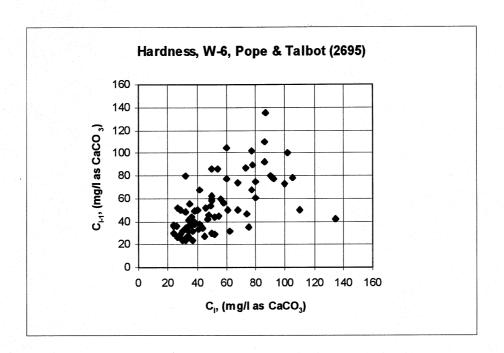


Figure 5-10 Clustering of Lower Values and Dispersion of High Values $(r_1=0.611, p=0.000)$

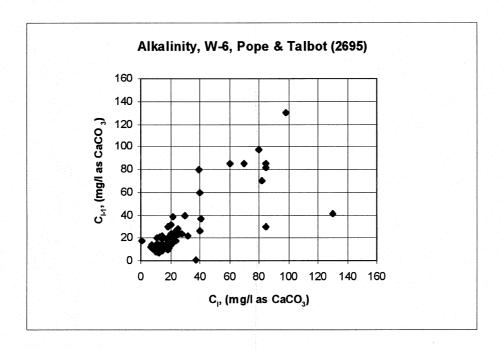


Figure 5-11 Clustering of Lower Values and Dispersion of High Values $(r_1=0.802, p=0.000)$

Autocorrelations were displayed in the form of correlograms, plots of the estimated lag-k autocorrelation coefficient as a function of k. Figures 5-12 and 5-13 show an example for which the autocorrelation was insignificant for all fifteen lags calculated. Figures 5-14 and 5-15 show correlograms for which significant correlation persists for a number of lags. It is not surprising that these two correlograms are almost the same. Figure 5-14 is hardness and Figure 5-15 is the alkalinity for the same well. Since this is a carbonate aquifer, we can expect hardness and alkalinity to be similar. Figure 5-16 and 5-17 are also correlograms for the alkalinity and hardness. Here, the correlations decay to zero after eight lags and then become negative (although not significantly). All the correlograms are included in appendix C.

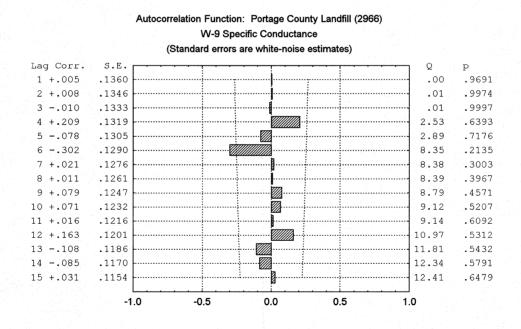


Figure 5-12 Corellogram with Insignificant Autocorrelation

Autocorrelation Function: Portage County Landfill (2966) W-9P Specific Conductance (Standard errors are white-noise estimates)

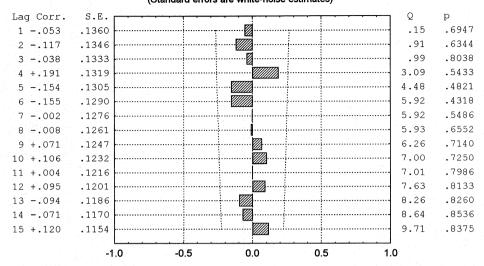


Figure 5-13 Corellogram with Insignificant Autocorrelation

Autocorrelation Function: Pope & Talbot Landfill (2695)
W-6 Hardness

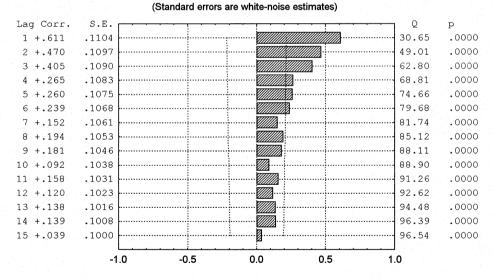


Figure 5-14 Corellogram with Autocorrelation Showing Longer Term Persistence

Autocorrelation Function: Pope & Talbot Landfill (2695) W-6 Alkalinity

(Standard errors are white-noise estimates)

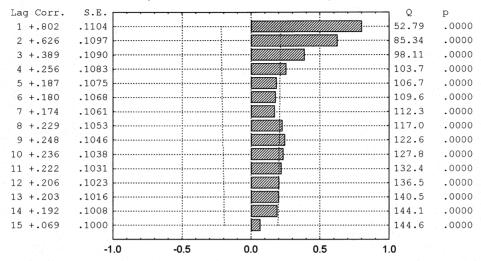


Figure 5-15 Corellogram with Autocorrelation Showing Longer Term Persistence

Autocorrelation Function: Portage County Landfill (2966)
W-9 Hardness
(Standard errors are white-noise estimates)

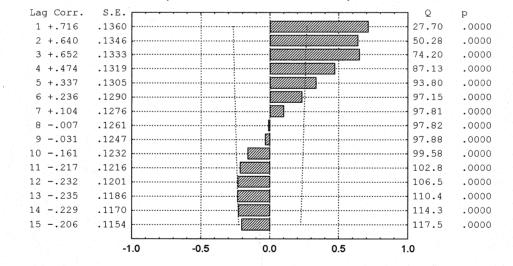


Figure 5-16 Corellogram with Autocorrelation Showing Shorter Term
Persistence

(Standard errors are white-noise estimates) Q Lag Corr. S.E. .0000 24.73 1 +.676 .1360 .0000 45.32 2 +.611 .1346 3 +.405 .1333 54.55 .0000 .0000 4 +.279 .1319 59.03 5 +.162 .1305 60.57 .0000 .0000 6 +.055 .1290 60.75 7 +.051 .1276 60.91 .0000 8 -.011 .1261 60.92 .0000 9 -.117 .1247 61.80 .0000 10 -.210 .1232 64.71 .0000 .0000 11 -.241 .1216 68.64 .1201 71.94 .0000 12 -.218

75.00

76.01

77.45

1.0

.0000

.0000

.0000

13 -.208

14 -.117

15 -.139

.1186

.1170

.1154

-1.0

-0.5

Autocorrelation Function: Portage County Landfill (2966)
W-10 Hardness

Figure 5-17 Corellogram with Autocorrelation Showing Shorter Term Persistence

0.0

0.5

Goodman (1987) and Montgomery et. al. (1987) also evaluated serial correlation coefficients of groundwater quality data. Goodman tested for serial correlation using two methods: lag-1 autocorrelation function; and the nonparametric runs test. The results were similar for both tests. For the parameters examined in this study, Goodman (1987) found a smaller proportion of significant lag-1 autocorrelation coefficients: 23, 26 and 24 percent compared to the 62, 65 and 58 percent found in this study for alkalinity, hardness and specific conductance, respectively. Much of the difference may be attributed to differences in record lengths. The average record length in Goodman (1987) was less than three years; in this study the average record length is about thirteen years. Montgomery et. al. (1987) estimated lag-1 autocorrelation coefficients for hardness and specific conductance at two and eight wells, respectively. Of these, only 50% of the hardness data and 25% of the specific conductance data showed significant autocorrelation. The average record length was only about seven years.

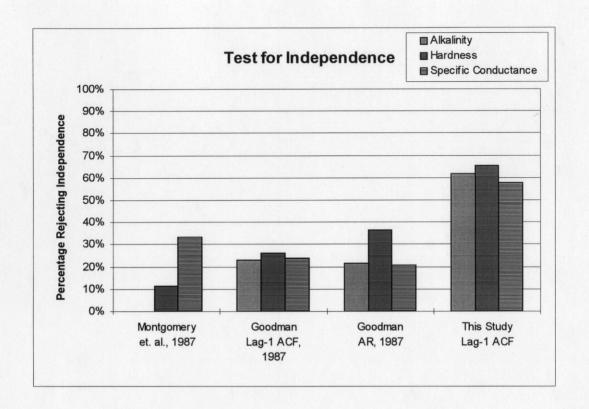


Figure 5-18 Results for Independence Tests

Our study indicates the data for upgradient wells are not random. One or more correlation coefficients for over 50 percent of the data sets are significantly different from zero. In the next section we will test the data for trend.

A number of parametric and nonparametric tests for trend are available. Stoline et. al. (1993) used parametric regression to detect both linear and quadratic downward trends for a superfund site. Frapporti et. al. (1994) used the Spearman rho test to evaluate changing geochemical processes. DUMPStat uses Sen's test to evaluate background data for upward historical trends or the entire data set for upward or downward trends for assessment monitoring. We

chose the gamma test, which is a slight variation of the Kendall tau test. The Gamma test and Sen's test procedures are outlined in appendix A.

The results of the gamma test are shown in Figure 5-19, for a significance level of five percent. Overall, at least 50 percent of the wells had a significant trend for alkalinity, hardness and specific conductance. Slightly fewer wells showed trend than autocorrelation for hardness and specific conductance.

Slightly more wells exhibited autocorrelation than trend for alkalinity. Further, the breakdown of the gamma trends were 53 percent positive and 47 negative. Figure 5-20 shows the totals for autocorrelation and trend on the same graph.

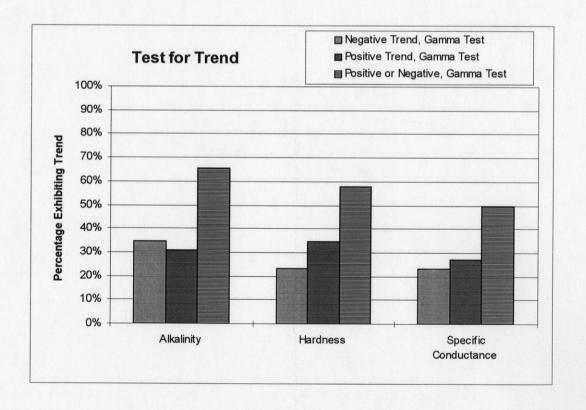


Figure 5-19 Overall Results for Trend Tests

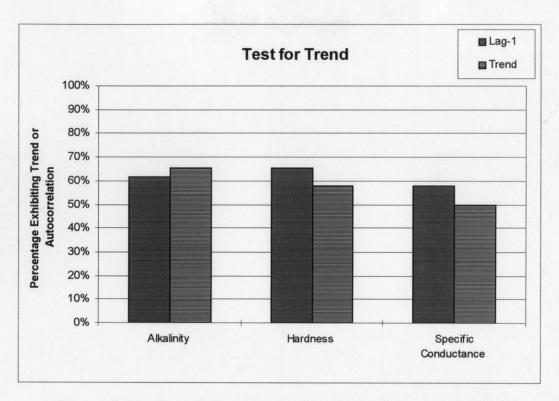


Figure 5-20 Overall Results for Trend and Autocorrelation

5.4 Normality

Prediction limits and combined Shewhart-CUSUM control charts, as used in DUMPStat, are generally predicated on the assumption of normality. A number of tests for normality are suggested in the literature. Gibbons (1994) lists several including:

- Normal Probability Plots
- Shapiro-Wilk Test
- Shapiro-Francia Test
- D'Agostino's Test

- Skewness
- Kolmogorov-Smirnov (mean and variance known) Goodness-of
 Fit Test
- Kolmogorov-Smirnov with Lilliefors Generalization (uses sample mean and sample standard deviation)

Montgomery et. al. (1987) and Gilbert (1987) suggest:

- Histograms
- Normal Probability Plots
- Chi-Squared Goodness-of-Fit Test
- Skewness

Before applying a test for normality, it is useful to examine histograms of the data. Histograms give qualitative information about data normality. For example, the alkalinity data shown in Figure 5-21 approximately exhibit normal behavior. Other data give a different picture. Figure 5-22 shows sampling data clearly skewed to the right while Figure 5-23 shows data clearly skewed to the left. Although the skewness coefficient would detect non-normal behavior for Figure 5-22 and 5-23 it would not detect the bimodality evident in Figure 5-24.

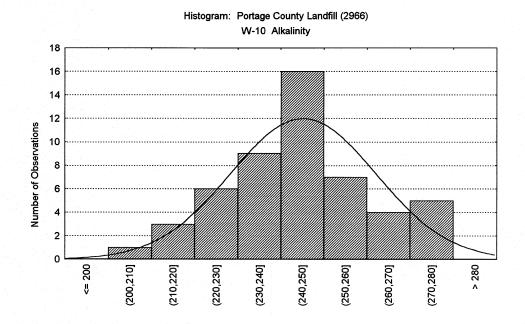


Figure 5-21 Histogram: Groundwater Data Approximately Normal (51 samples, 0.111 skewness coefficient)

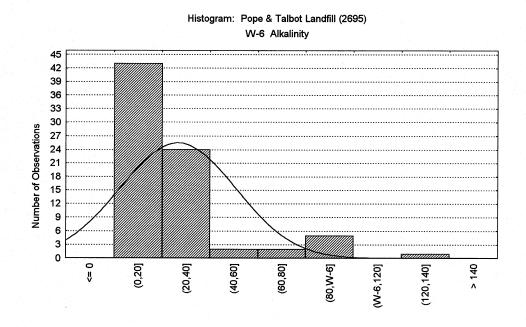


Figure 5-22 Histogram: Groundwater Data Skewed to the Right (77 samples, 2.294 skewness coefficient)

MW-6 Alkalinity 20 18 16 14 Number of Observations 10 (220,230] (230,240] (240,250] <= 220 (260,270] (270,280] (280,290] (300,310] > 310 (250, 260](290,300]

Histogram: City of Richland Center Landfill (3065)

Figure 5-23 Histogram: Groundwater Data Skewed to the Left (43 samples, -2.470 skewness coefficient)

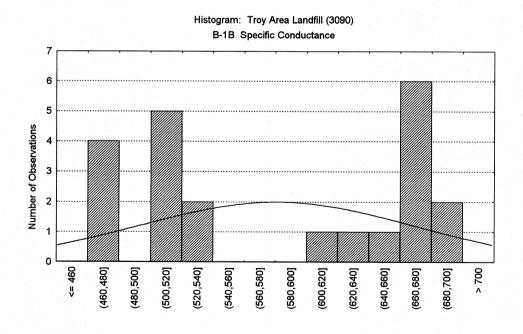


Figure 5-24 Histogram: Groundwater Data Bimodal Zero Skew (22 samples, -0.029 skewness coefficient)

For a quantitative look at normality we performed two numerical tests: the Shapiro-Wilk (SW) test and the Kolmogorov-Smirnov test (KS). The Lilliefors test procedure was used for the KS test since the mean and variance were not known a priori. For these tests, we analyzed 26 upgradient wells at ten landfills in Wisconsin for normality. For each well we used alkalinity, hardness and specific conductance.

Using a five percent significance level, we found that the KS and SW provided similar results for alkalinity and specific conductance and identical results for hardness. For the SW test, 54, 73 and 54 percent of the wells were rejected for normality for alkalinity, hardness and specific conductance respectively. Similarly for the KS test, 50, 73 and 46 percent of the wells were rejected for normality for alkalinity, hardness and specific conductance. Overall, normality was rejected 60 percent of the time for the SW test and 56 percent of the time for the KS test.

The results from our tests are summarized in Figure 5-25. In addition, this figure compares our results with the results from Montgomery et. al. (1987), Goodman (1987), and Fisher (1989). Note that Montgomery et. al. (1987) did not use alkalinity in any of the tests and only tested three wells for hardness. From Figure 5-25, one can see from all three studies that normality is rejected more often for hardness than for alkalinity and specific conductance. Basically, our study shows that the assumption of normality is frequently violated.

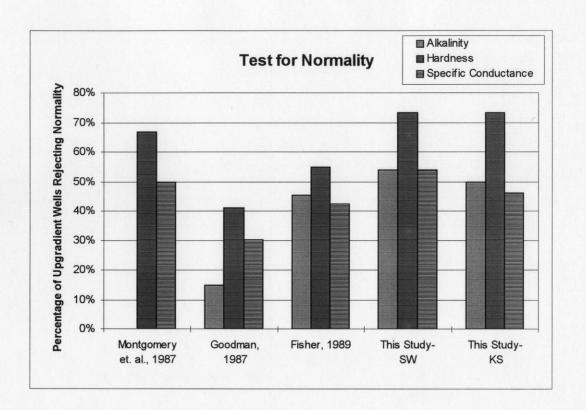


Figure 5-25 Results for Tests of Normality

We also applied the SW test to the log transform of the data. Table 5-3 gives the results of these tests. For the parameters alkalinity, hardness and specific conductance 58, 46 and 61 percent of the wells were rejected for lognormality. This is comparable to the rejection rate of 54, 73 and 54 percent for normality. Also, 46, 46 and 42 percent of the wells were rejected for both normality and lognormality. Alternately, 35, 27 and 27 percent of the wells were not rejected for both normality and lognormality. These results are not encouraging for the use of prediction limits or control charts based on normality

or lognormality. In about 50 percent of the cases considered neither assumption appears to be warranted.

	Rejecting	Rejecting	Rejecting	Not Rejecting
	Normality	Lognormality	Both	Both
Alkalinity	14/26	15/26	12/26	9/26
Hardness	19/26	12/26	12/26	7/26
Specific Conductance	14/26	16/26	11/26	7/26

Table 5-3 Results for Shapiro-Wilk Test for Normality and Lognormality

5.5 How do Violations of Assumptions Affect Parametric Tests?

In previous sections, we found that over 50 percent of the upgradient wells had a significant trend, non-normal distribution or serial correlation. In fact, 93, 100 and 88 percent of the wells were significant for at least one of the three for alkalinity, hardness and specific conductance, respectively. In the case study discussed in the previous chapter, we found that a number of upgradient wells had violations. Further, we found the false positive rate to be higher than expected. In this section, we will use DUMPStat and the indicator PAL to evaluate the same 26 upgradient wells at ten landfills. Again, the parameters alkalinity, hardness and specific conductance will be used. This analysis was

performed using eight background samples. After the analysis, we tabulated the number of data points out of compliance. Next, we estimated the false positive rate. Lastly, we determined if the data points that were out of compliance were in data sets exhibiting nonormality, serial correlation, and trend.

Figure 5-26 shows the number of violations for each parameter for all the upgradient wells. Figure 5-27 shows the estimates of the false positive rate based on the same wells. Clearly, the PL and CC indicate a large number of violations in the upgradient data. For all cases, the CC had more violations than the PL and the PL had more violations than the PAL. For instance, alkalinity had 2,151 and 202 violations for the PAL, PL and CC, respectively. Since all the data is upgradient, these violations are all false positive results. Overall the false positive rate estimate is 0.7, 10.9, and 21.4 percent for the PAL, PL and CC, respectively.

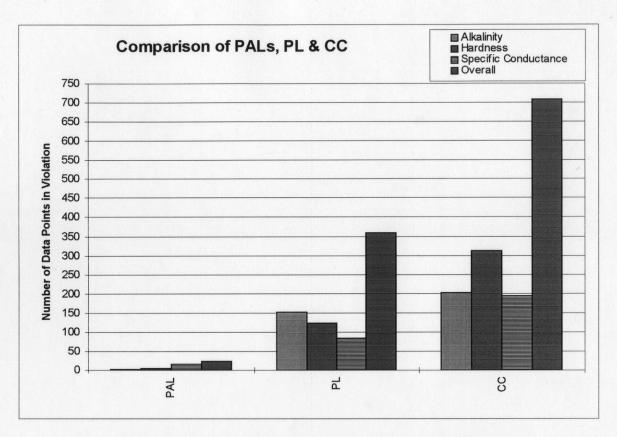


Figure 5-26 Number of Data Points in Violation (Based on Eight Background Samples)

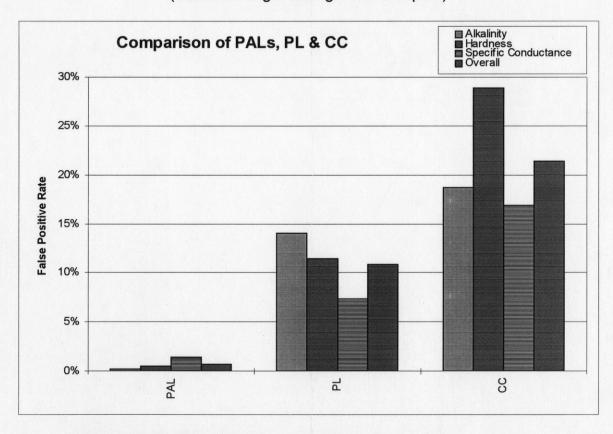


Figure 5-27 False Positive Rate Based on Upgradient Wells (Based on Eight Background Samples)

How do these false positive results relate to violation of the assumptions of normality and randomness in the indicator PAL, PL and CC. Figure 5-28 shows the percent of PAL exceedances which occur in datasets which exhibit serial correlation, nonormality, or trend. Figures 5-29 and 5-30 show the corresponding plots for PL and CC respectively. Of the PAL violations, 83, 78 and 74 percent of the violations were in data sets showing nonormality, serial correlation or trend respectively. (Note however, that they were close for PAL exceedances for alkalinity and hardness data.) For the PL, at least 65 percent of the violations are in data sets showing nonormality, serial correlation, or trend. For the CC, at least 57 percent of the violations are in data sets violating assumptions. From this, we can conclude data that violates the assumptions is more likely to be out of compliance than in compliance.

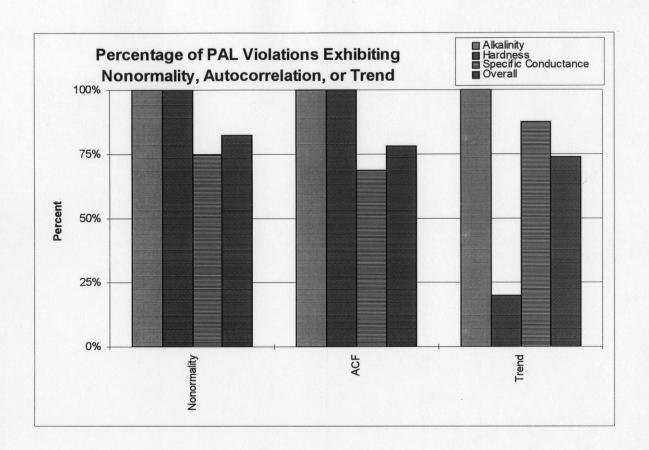


Figure 5-28 Percentages of PAL Violations in Data Sets Deviating from Parametric Assumptions

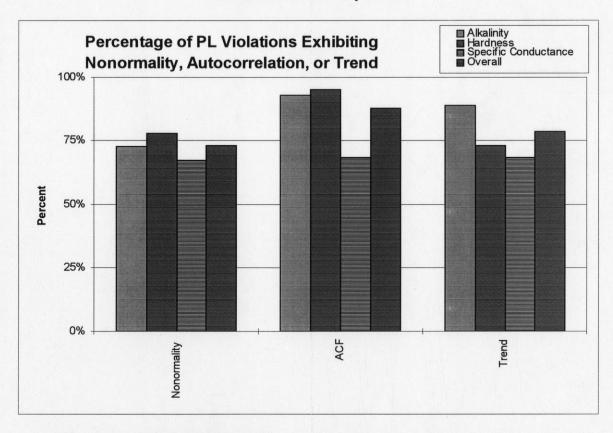


Figure 5-29 Percentages of PL Violations in Data Sets Deviating from Parametric Assumptions

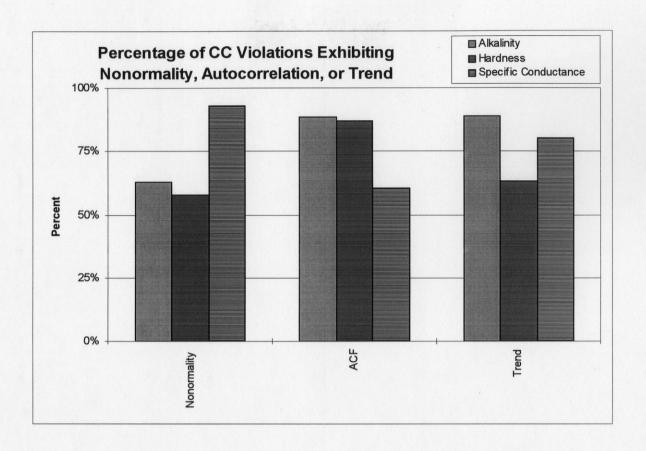


Figure 5-30 Percentages of CC Violations in Data Sets Deviating from Parametric Assumptions

5.6 Chapter Summary

The tests used in DUMPStat assume that the data are random and are normally distributed (or can be normally distributed with a log transformation).

Analysis of alkalinity, hardness and specific conductance data for 26 upgradient wells indicated that these assumptions are commonly violated. Over 50 percent of these wells, for each parameter, had a significant autocorrelation coefficient

or trend which indicates the data are not random. Also, over 50 percent of the upgradient wells, for each parameter, rejected normality and over 50 percent rejected lognormality.

CHAPTER 6

RECOMMENDATIONS AND CONCLUSIONS

We cannot recommend the use of DUMPStat as a supplement or alternative to the indicator PAL. Use of DUMPStat's methods for intrawell comparison, prediction limits and combined Shewhart-CUSUM control charts, will result in many false indications of contamination. This is true for two reasons. First, DUMPStat's intrawell tests are inherently more conservative than the indicator PAL, except at small sample sizes. Second, the assumptions on which these tests are based, that the data are random samples from normal distributions, are often violated by upgradient landfill monitoring data.

Nonrandomness in particular can trigger false indication of contamination.

CHAPTER 7

References

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APPENDIX A STATISTICAL TESTS

APPENDIX A STATISTICAL TESTS

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A1.0 Sen's Test

Reference:

Gibbons, Robert D. <u>Statistical Methods For Groundwater Monitoring</u>. New York: John Wiley & Sons, Inc., 1994.

Sen developed this nonparametric trend test. DUMPStat uses it to identify historical trends if the option is selected in the setup menu. The test estimates the slope of the concentration versus time regression line. According to Gibbons this method is robust to outliers, missing data and nondetects. The required data includes the sample values and the order of the samples in time for a well and constituent.

The one sided hypothesis test is:

Null hypothesis

H₀: The data has no upward trend

Alternate hypothesis

H₁: The data has an upward trend

A1.1 Test Procedure

First, compute N estimates of the slope Q where

$$N = \frac{n \cdot (n-1)}{2}$$
 (A1.1)

and

$$Q = \frac{X_j - X_i}{j - i} \quad (A1.2)$$

for all combinations of j greater than i where j=(i+1),....n and i=1,....(n-1).

n is the total number of sample points, x_i is the ith sample concentration and x_{j+1} is the jth sample concentration.

Second, rank the N values of Q from smallest to largest. Compute the median slope S, where

$$S = Q_{[(N+1)/2]} \text{ if N is odd}$$
 (A1.3)

$$S = Q_{[N/2]} + Q_{[(N+2)/2]}$$
 if N is even (A1.4)

Third, estimate the variance of S, V[S]. For n>10 the V[S] can be estimated by

$$V[S] = \frac{n \cdot (n-1) \cdot (2n+5) - \sum\limits_{p=1}^{q} t_{p} \cdot p \cdot (p-1) \cdot (2 \cdot p + 5)}{18}$$
 (A1.5)

where t_p is the number of ties with extent p. For example for t_4 =2 there are two ties of extent 4. There are two measurements repeated four times each.

Fourth, compute the approximate lower confidence limit which is the Mth smallest value of Q where

$$M = \frac{N - Z_{1-\alpha} \cdot V[S]^{1/2}}{2}$$
 (A1.6)

 $Z_{1-\alpha}$ is the lower one sided $(1-\alpha)100$ percent confidence Z statistic for the normal distribution, where α is the significance level of the test.

Lastly, locate the Mth smallest value of Q. Typically, M is not an integer so interpolate to approximate the Mth largest Q. If this interpolated value is greater than zero reject the null hypothesis.

A1.2 Sample Calculation

Compute the number of slope estimates, N

$$N=\frac{10\cdot 9}{2}=45$$

Compute slope estimates, Q

Time	Sample									
Period	Concentration									
1	200		_							
2	210	10.00								
3	225	12.50	15.00		_					
4	210	3.33	0.00	-15.00		_				
5	260	15.00	16.67	17.50	50.00		_			
6	260	12.00	12.50	11.67	25.00	0.00		_		
7	290	15.00	16.00	16.25	26.67	15.00	30.00		_	
8	250	7.14	6.67	5.00	10.00	-3.33	-5.00	-40.00		_
9	260	7.50	7.14	5.83	10.00	0.00	0.00	-15.00	10.00	
10	300	11.11	11.25	10.71	15.00	8.00	10.00	3.33	25.00	40.00

Table A1.1 Individual Slope Estimates

Now rank the Q values from smallest to largest.

Rank	Q	Rank	Q	Rank	Q	Rank	Q
1	-40	13	5.833	24	10.71	35	15
2	-15	14	6.667	25	11.11	36	16
3	-15	15	7.143	26	11.25	37	16.25
4	-5	16	7.143	27	11.67	38	16.67
5	-3.33	17	7.5	28	12	39	17.5
6	0	18	8	29	12.5	40	25
7	0	19	10	30	12.5	41	25
8	0	20	10	31	15	42	26.67
9	0	21	10	32	15	43	30
10	3.333	22	10	33	15	44	40
11	3	23	10	34	15	45	50
12	5						

Table A1.2 Ranked Slope Estimates

Compute The Median Slope

$$S = Q_{23} = 10.00$$

Estimate the variance of S.

$$V[S] = \frac{n \cdot (n-1) \cdot (2n+5) - \sum\limits_{p=1}^{q} t_p \cdot p \cdot (p-1) \cdot \left(2p+5\right)}{18} = \frac{10 \cdot 9 \cdot 25 - 1 \cdot 3 \cdot 2 \cdot 11 - 1 \cdot 2 \cdot 1 \cdot 9}{18} = 120.33$$

where n=10 is the number of sample points. Also, the data contains two ties. The first is a two way tie with a concentration of 210. Consequently t_2 =1. The second is a three way tie with a concentration of 260. Similarly t_3 =1.

For the 99 percent one sided test $Z_{0.99}$ =2.33.

Determine the Mth smallest slope estimate Q.

$$M = \frac{N - Z_{_{1-\alpha}} \cdot V[S]^{^{1/2}}}{2} = \frac{45 - 2.33 \cdot 120.33}{2} = 9.72$$

So select the ninth and tenth smallest values for the slope estimate, Q. Then interpolate to get the Mth smallest Q value.

$$Q_9=0$$
 $Q_{10}=3.33$
 $Q_{9.72}=2.40$

 $Q_{9.72}$ >0 so the reject null hypothesis that trend is zero (the data may have an upward trend).

A2.0 Dixon's Test

Reference:

Gibbons, Robert D. <u>Statistical Methods For Groundwater Monitoring.</u> New York: John Wiley & Sons, Inc., 1994.

Dixon developed this test for possible outliers. According to Gibbons (1994), Dixon's Test can be used where a small number of outliers are suspected. DUMPStat searches for outliers in the background data. If outliers are included in the background data, the control limit could be set too high or too low. This influences the false positive and false negative rate.

The two sided hypothesis test is:

Null hypothesis

H₀: The data set contains no outliers

Alternate hypothesis

H₁: The data set contains outliers

A2.1 Test Procedure

To perform the test, first rank the data from smallest to largest. Second, calculate the Dixon statistic for the highest and lowest values in the data set. Table A.2.1 shows the Dixon statistics based on the number of sample values, n.

n	Highest Value	Lowest Value		
3-7	$X_n - X_{n-1}$	$X_2 - X_1$		
	$X_n - X_1$	$X_n - X_1$		
8-10	$X_n - X_{n-1}$	$X_2 - X_1$		
	$X_n - X_2$	$X_{n-1} - X_1$		
11-13	$X_n - X_{n-2}$	$X_3 - X_1$		
	$X_n - X_2$	$X_{n-1} - X_1$		
14-25	$X_n - X_{n-2}$	$X_3 - X_1$		
	$X_n - X_3$	$X_{n-2} - X_1$		

Table A2.1 Dixon's Test Statistic Equations

Lastly compare these statistics with the critical Dixon's statistics provided in Table A2.2. The test may be repeated if additional outliers are suspected. However, this may increase the probability that a data point is removed unnecessarily. The probabilities in Table A2.2 are for a single test for outliers.

n	5%	1%	n	5%	1%
3	0.941	0.988	14	0.546	0.641
4	0.765	0.889	15	0.525	0.616
5	0.642	0.780	16	0.507	0.595
6	0.560	0.698	17	0.490	0.577
7	0.507	0.637	18	0.475	0.561
8	0.554	0.683	19	0.462	0.547
9	0.512	0.635	20	0.450	0.535
10	0.477	0.597	21	0.440	0.524
11	0.576	0.679	23	0.421	0.505
12	0.546	0.642	24	0.413	0.497
13	0.521	0.615	25	0.406	0.498

Table A2.2 Critical Values for Dixon's Statistic

A2.2 Sample Calculation

Rank	Sample
	Concentration
1	2.00
2	210
3	210
4	225
5	250
6	260
7	260
8	260
9	290
10	3000

Table A2.3 Sample Data for Dixon's Test

n=10 therefore use $\frac{X_n - X_{n-1}}{X_n - X_2}$ and $\frac{X_2 - X_1}{X_{n-1} - X_1}$ for the high value and low value Dixon statistic. The corresponding results are

$$x_{n-1}=2.90$$

$$x_n = 3000$$

$$x_2 = 210$$

$$x_1 = 2.00$$

The Dixon statistic for the upper value is

$$\frac{x_n^2 - x_{n-1}}{x_n - x_n^2} = \frac{3000 - 290}{3000 - 210} = 0.971$$

The Dixon statistic for the lower value is

$$\frac{x_2 - x_1}{x_{2,1} - x_1} = \frac{210 - 2.00}{290 - 2.00} = 0.722$$

The critical Dixon Statistic with a one percent false positive rate is 0.597. Since both Dixon statistics calculated above the critical Dixon statistic, reject the null hypothesis. Consequently, the highest and lowest value in the data set should be removed prior to calculating the mean, standard deviation, PAL, prediction limit or control chart.

A3.0 Intrawell Shewart-CUSUM Control Charts

Reference:

Gibbons, Robert D. <u>Statistical Methods For Groundwater Monitoring.</u> New York: John Wiley & Sons, Inc., 1994.

The Shewart-CUSUM control chart is one intrawell option available in DUMPStat. This method combines the Shewart control chart which can detect immediate releases, and the CUSUM control chart which can detect gradual releases. The data in order of measurement are required for this analysis.

DUMPStat notifies the user of historical trends using Sen's test and removes outliers using Dixon's test. Procedures and examples of these tests are included in A1.0 and A2.0.

The one sided hypothesis test is:

Null hypothesis

H₀: The well is not in violation

Alternate hypothesis

H₁: The well is in violation

A3.1 Test Procedure

DUMPStat selects three Shewart-CUSUM control chart parameters.

n	k	SCL	h
n < 12	1.0	4.5	4.5
n ≥ 12	0.75	4.0	4.0

Table A3.1 Shewart-CUSUM Control Chart Parameters

n is the number of background samples. The user may set the minimum number of n to 4, however, Gibbons and the WDNR recommend a minimum of 8 background samples. k is a parameter related to the displacement that should be quickly detected. h is value which the CUSUM will be compared. SCL is the number of standard deviations the Shewart will be compared. These values are slightly more conservative than the EPA recommendations.

First, compute the mean, \bar{x} , based on the background data where

$$\overline{X} = \frac{\sum_{i=1}^{n} X_{i}}{n}$$
 (A3.1)

Second, calculate the standard deviation, s, based on the background data where

$$s = \sqrt{\frac{\sum\limits_{i=1}^{n} x_i - \overline{x}}{n-1}}$$
 (A3.2)

Third, calculate the z score for all i=1,2....N where

$$z_{i} = \frac{x_{i} - \overline{x}}{s} \tag{A3.3}$$

and N is the total number of samples both background and beyond.

Fourth, calculate the CUSUM, S, for all i=1,2....N where

$$S_i = max[0, (z_i - k) + S_{i-1}]$$
 (A3.4)

and

$$S_0 = 0$$

The Shewart-CUSUM control chart is out of control if during a sampling period $S_i \ge h$ or $z_i \ge SCL$. This must be verified by the next round of sampling before action is required.

A3.2 Sample Calculation

For the data listed below use the first 8 samples for the background. Compute the Shewart-CUSUM control chart based on the entire sample.

Time Period	Sample Concentration	Time Period	Sample Concentration
1	200	7	240
2	210	8	250
3	225	9	260
4	210	10	310
- 5	260	11	320
6	260	12	260

Table A3.2 Sample Data for Shewart-CUSUM Control Chart

Calculate the mean and standard deviation based on the eight background samples. The results are listed in Table A3.3. Also, the three Shewart-CUSUM control chart parameters are listed.

Mean, \overline{x}	231.875
Standard Deviation, s	23.895
k	. 1
h=SCL	4.5
Number of Background, n	8
Total Samples, N	12

Table A3.3 Mean, Standard Deviation and Shewart-CUSUM Control Chart Parameters

Calculate z_i and S_i based on equations A3.3 and A3.4. Displays of the control chart in graphical and tabular form are included in table A3.4 and figure A3.1.

Time	Sample	Shewart	CUSUM	Control Chart
Period, i	Concentration, x _i	Z _i	Si	Status
1	200	-1.33	0.00	In Control
2	210	-0.92	0.00	In Control
3	225	-0.29	0.00	In Control
4	210	-0.92	0.00	In Control
5	260	1.18	0.18	In Control
6	260	1.18	0.35	In Control
7	240	0.34	0.00	In Control
8	250	0.76	0.00	In Control
9	260	1.18	0.18	In Control
10	310	3.27	2.45	In Control
11	320	3.69	5.13	Out of Control (Hit)
12	260	1.18	5.31	Out of Control (Verified Hit)

Table A3.4 Tabular Shewart-CUSUM Control Chart

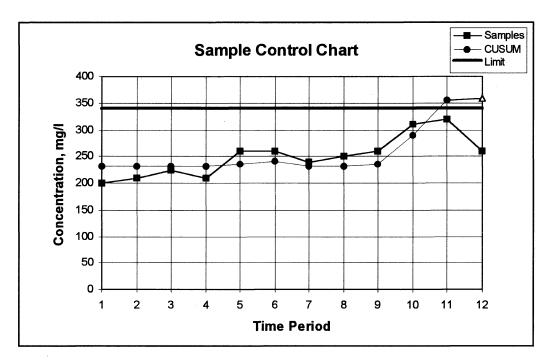


Figure A3.1 Graphical Shewart-CUSUM Control Chart

Note the eleventh concentration is out of control for the CUSUM portion of the control chart. This is verified on the twelfth sample. Therefor action should be taken.

A4.0 Intrawell Prediction Limits

Reference:

Gibbons, Robert D. Statistical Methods For Groundwater Monitoring. New York: John Wiley & Sons, Inc., 1994.

The prediction limit is one intrawell option available in DUMPStat. This method is detects immediate releases faster than the Shewart-CUSUM control charts but cannot detect slow or gradual releases. The data in order of measurement are required for this analysis.

DUMPStat notifies the user of historical trends using Sen's test and removes outliers using Dixon's test. Procedures and examples of these tests are included in A1.0 and A2.0.

The one sided hypothesis test is:

Null hypothesis

H₀: The well not in violation

Alternate hypothesis H₁: The well is in violation

A4.1 Test Procedure

First, select α as the minimum of 0.01 or $\alpha = \sqrt{1 - 0.95^{\frac{1}{k}}}$ for one verification resample. For this study pass 1 of 1 resamples was selected from the setup menu for the verification resampling plan. Other verification resampling plans can be selected in the DUMPStat setup menu. k is the number of future comparisons to be made.

Second, compute the intrawell prediction limit, PL, as:

$$PL = \overline{X} + s \cdot t_{\begin{bmatrix} 1-\alpha,n-1 \end{bmatrix}} \cdot \sqrt{1 + \frac{1}{n}}$$
 (A4.1)

where \overline{x} is the mean and s is the standard deviation of the background data (see equations A3.1 and A3.2). n is the number of background data points. $t_{t_{1-\alpha,n-1}}$ is the student's upper one sided t statistic based on n-1 degrees of freedom and a 1- α confidence level (see table A4.1 on the next page).

n-1	t _{0.99}	n-1	t _{0.99}
1	31.82054	17	2.566936
2	6.964523	18	2.552382
3	4.54071	19	2.539488
4	3.746956	20	2.527977
5	3.364931	21	2.517649
6	3.142664	22	2.508326
7	2.997957	23	2.499865
8	2.896457	24	2.492159
9	2.821438	25	2.485111
10	2.76377	26	2.478627
11	2.718078	27	2.472664
12	2.680996	28	2.467141
13	2.650311	29	2.46202
14	2.624493	30	2.457266
15	2.602481	50	2.403274
16	2.583488	31999	2.326463

Table A4.1 Student's t Statistics Based on n-1 Degrees of Freedom and a 99 Percent Confidence Level

Third compare the prediction limit to the data of interest. If the data exceeds the prediction limit two consecutive times the well is in violation.

A4.2 Sample Calculation

Compute the prediction limit based on the first eight background samples (n=8). See if samples nine through twelve exceed the prediction limit.

Time Period	Sample Concentration	Time Period	Sample Concentration
1	200	7	240
2	210	8	250
3	225	9	260
4	210	10	300
5	260	11	320
6	260	12	322

Table 4.2 Sample Data for Prediction Limit

There are four future samples therefor k=4. Consequently, α is the minimum of 0.01 and $\alpha = \sqrt{1 - 0.95^{\frac{1}{k}}} = \sqrt{1 - 0.95^{\frac{1}{4}}} = 0.113$.

Select $t_{[0.99,7]}$ =2.997957 from Table 4.2.

For the background data the mean, \bar{x} , is 231.875 and the standard deviation, s, is 23.895.

$$PL = \overline{x} + s \cdot t_{\left[\frac{1-\alpha,n-1}{4}\right]} \cdot \sqrt{1 + \frac{1}{n}} = 231.9 + 23.9 \cdot 2.998 \cdot \sqrt{1 + \frac{1}{8}} = 307.9$$

The eleventh and twelfth samples exceed the prediction limit. The well may not be in compliance and action should be taken.

A5.0 Gamma Test

References:

Kendall, M.G, and Jean Dickenson Gibbons. <u>Rank Order Correlation.</u> New York: Oxford University Press, 1990.

Statsoft, Inc. <u>Statistica for Windows</u> [Computer Program and Manual]. Tulsa, Oklahoma: Statsoft, Inc., 1997.

The Gamma test is a nonparametric trend test. It is equivalent to the Kendall tau, except that ties are explicitly taken into account. The Gamma test yields exactly the same results as the Kendall tau type b, except the calculation of the gamma statistic removes the ties from the denominator of the statistic. Like Sen's test the Gamma test is robust to outliers missing data and nondetects. The required data includes the sample values and the order of the samples in time for a well and constituent.

Null hypothesis

H₀: The data has no upward trend

Alternate hypothesis

H₁: The data has an upward or downward trend

A5.1 Test Procedure

First, compute N values of the $sgn(x_i-x_i)$ where

$$N = \frac{n \cdot (n-1)}{2} \tag{A5.1}$$

and

$$sgn(x_{i} - x_{i}) = \begin{cases} +1 & \text{if} \quad x_{i} - x_{i} > 0\\ 0 & \text{if} \quad x_{i} - x_{i} = 0\\ -1 & \text{if} \quad x_{i} - x_{i} < 0 \end{cases}$$
 (A5.2)

for all combinations of j greater than i where j=(i+1),...n and i=1,....(n-1). n is the total number of sample points, x_i is the ith sample concentration and x_{j+1} is the jth sample concentration.

Second, calculate the Mann-Kendall statistic, S as

$$S = \sum_{i=1}^{n-1} \sum_{j=1}^{n} sgn(x_{j} - x_{j})$$
 (A5.3)

Basically, this is the sum of the +1 and -1 values calculated in equation A5.2. If S is zero, no trends exist.

Third, calculate the gamma statistic as,

$$\gamma = \frac{S}{N - \sum_{p=1}^{q} t_p \cdot \frac{p}{2} \cdot (p-1)}$$
 (A5.4)

where t_p is the number of ties with extent p. For example for t_4 =2 there are two ties of extent four. There are two measurements repeated four times.

Fourth, estimate the variance of S, V[S]. For n>10 the V[S] can be estimated by

$$V[S] = \frac{n \cdot (n-1) \cdot (2n+5) - \sum_{p=1}^{q} t_{p} \cdot p \cdot (p-1) \cdot (2 \cdot p+5)}{18}$$
 (A5.3)

Fifth, for n>10 the z statistic can be estimated for the two sided normal test by

$$z = \begin{cases} \frac{S-1}{V[S]^{\frac{1}{2}}} & \text{if } S > 0\\ 0 & \text{if } S > 0\\ \frac{S+1}{V[S]^{\frac{1}{2}}} & \text{if } S < 0 \end{cases}$$
 (A5.4)

Sixth, compare z value with two sided z value with a 95 percent confidence level. Therefor if |z| > 1.96 then reject the null hypothesis. The data may have a trend.

A5.2 Sample Calculation

Compute the number of $sgn(x_j-x_i)$ values, N

$$N = \frac{10 \cdot 9}{2} = 45$$

Compute $sgn(x_i-x_i)$

Time	Sample									
Period	Concentration									
1	200		_							
2	210	1		-				•		
3	225	1	1		_					
4	210	1	0	-1		_				
5	260	1	1	1	1		_			
6	260	- 1	1	1	1	0		_		
7	290	1	1	1	1	1	1		_	
8	250	1	1	1	1	-1	-1	-1		_
9	260	1	1	1	1	0	0	-1	1	
10	300	1	1	1	1	1	1	1	1	1

Table A5.1 Individual sgn(x_i-x_i) for Gamma Test

Calculate the Mann-Kendall statistic, S by adding all the 1's, 0's and -1's in Table A1.1.

$$S = \sum\limits_{i=1}^{n-1} \sum\limits_{i+1}^{n} sgn(x_i - x_i) = 36 \cdot 1 + 4 \cdot 0 + 5 \cdot -1 = 31$$

Calculate the gamma statistic, γ

$$\gamma = \frac{S}{N - \left[\sum\limits_{p=1}^{q} t_{_{p}} \cdot \frac{p}{2} \cdot (p-1)\right]} = \frac{31}{45 - \left[1 \cdot \frac{2}{2} \left(2-1\right) + 1 \cdot \frac{3}{2} \left(3-1\right)\right]} = 0.75$$

Estimate the variance of S,

$$V\!\!\left[S\right] = \frac{n \cdot (n-1) \cdot (2n+5) - \sum\limits_{p=1}^{q} t_p \cdot p \cdot (p-1) \cdot \left(2p+5\right)}{18} = \frac{10 \cdot 9 \cdot 25 - 1 \cdot 3 \cdot 2 \cdot 11 - 1 \cdot 2 \cdot 1 \cdot 9}{18} = 120.33$$

\$>0 so calculate z, based on equation 5.4.

$$z = \frac{S - 1}{V[S]^{\frac{1}{2}}} = \frac{31 - 1}{120.33^{\frac{1}{2}}} = 2.735 > 1.96$$

Reject the null hypothesis. The data may have an upward trend.

A6.0 Autocorrelation Function

References:

Box, E. P. and G. M. Jenkins. <u>Time Series Analyis Forecasting and Control.</u> San Francisco: Holden Day, 1976.

Kendall, M.G. and Ord, J. K. <u>Time Series.</u> New York: Oxford University Press, 1990.

Statsoft, Inc. <u>Statistica for Windows</u> [Computer Program and Manual]. Tulsa, Oklahoma: Statsoft, Inc., 1997.

The autocorrelation or serial correlation coefficient relates a series of data with itself shifted by a lag of k observations. This is a test of independence. The required data includes the sample values and the order of the samples in time for a well and constituent.

Null hypothesis

H₀: The data does not have autocorrelation

Alternate hypothesis

H₁: the data has autocorrelation

A6.1 Test Procedure

The autocorrelation function, rk, may be estimated as

$$r_{k} = \frac{c_{k}}{c_{0}} \tag{A6.1}$$

where

$$c_{k} = \frac{1}{N} \sum_{i=1}^{N-k} (x_{i} - \overline{x}) (x_{i+k} - \overline{x})$$
 (A6.2)

is the estimate of the autocovariance and \bar{x} is the mean of the data. k is the number of the lag. In this study a k is one was used. Also, i=1,2....(N-k).

Assuming the series is a white noise process (all the autocorrelations are zero), the standard error of the autocorrelation function, $SE[r_k]$ can be estimated by:

$$SE[r_k] = \sqrt{\frac{1}{N} \cdot \frac{N - k}{N + 2}}$$
 (A6.3)

Lastly, calculate the normal z statistic and compare it to the two sided tabulated value at the 95 percent confidence level (1.96). If z>1.96 then reject the null hypothesis.

$$z = \frac{r_{k}}{SE[r_{k}]} \tag{A6.4}$$

A6.2 Sample Calculation

Calculate the autocorrelation function for the following data.

Time	Sample
Period	Concentration
1	222
2	210
3	200
4	220
5	240
6	260
7	280
8	300
9	320
10	340

Table A6.1 Sample Data for Autocorrelation Function

Calculate the autocovariance functions, co and c1.

$$c_{_{_{1}}} = \frac{1}{N} \sum_{_{_{i=1}}}^{9} \left(x_{_{_{i}}} - \overline{x}\right) \! \left(x_{_{_{i+1}}} - \overline{x}\right) = \frac{1}{10} \! \left[\! \left(222 - 255.2\right) \! \left(210 - 255.2\right) \! + \left(210 - 255.2\right) \! \left(200 - 255.2\right) \! + \left(210 - 255.2\right) \! \left(200 - 255.2\right) \! + \left(210 - 255.2\right) \! \left(200 - 255.2\right) \! + \left(210 - 255.2\right) \! + \left(21$$

Similarly,

$$c_0 = 1651.4$$

From these calculate the lag-1 autocorrelation function, r₁.

$$r_1 = \frac{c_1}{c_0} = \frac{1343.7}{1651.4} = 0.814$$

Now, estimate the standard error of the autocorrelation function

SE[r_k] =
$$\sqrt{\frac{1}{N} \cdot \frac{N-k}{N+2}} = \sqrt{\frac{1}{10} \cdot \frac{10-1}{10+2}} = 0.075$$

Next, compute the standard normal variate, z

$$z = \frac{r_k}{SE[r_k]} = \frac{0.814}{0.075} = 10.85 > 1.96$$

This z value is greater than 1.96 so reject the null hypothesis. The data may have serial correlation and the assumption of independence is in question.

APPENDIX B GAMMA TEST RESULTS

APPENDIX B GAMMA TEST RESULTS

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B1.1	Lincoln Landfill (1779)	B-1
B1.2	Pope & Talbot Landfill (2695)	B-2
B1.3	Oneida County Landfill (2805)	B-3
B1.4	Portage County Landfill (2966)	B-4
B1.5	Grede Foundries Landfill (2974)	B-5
B1.6	Sauk County Landfill (2978)	B-6
B1.7	City of Richland Center Landfill (3065)	B-7
B1.8	Juneau County Landfill (3070)	B-8
B1.9	Troy Area Landfill (3090)	B-9
B1.10	Lincoln County Landfill (3141)	B-10

B1.1 Lincoln Landfill (1779)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
MW-1	19	-0.31	-1.85	0.065
MW-4	19	-0.77	-4.48	7E-06

Hardness

	Valid			
	N	Gamma	Z	p-level
MW-1	19	-0.37	-2.14	0.032
MW-4	20	-0.7	-4.22	2E-05

	Valid			
	N	Gamma	Z	p-level
MW-1	46	0.131	1.248	0.212
MW-4	22	-0.75	-4.79	2E-06

B1.2 Pope & Talbot Landfill (2695)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
W-13	66	0.638	7.526	5E-14
W-18	52	-0.08	-0.79	0.43
W-6	77	0.206	2.609	0.009

Hardness

	Valid			
	N	Gamma	Z	p-level
W-13	66	0.261	2.99	0.003
W-18	52	0.449	4.634	4E-06
W-6	77	-0.15	-1.96	0.05

	Valid			
	N	Gamma	Z	p-level
W-13	60	0.364	4.098	4E-05
W-18	46	0.365	3.564	4E-04
W-6	71	0.235	2.896	0.004

B1.3 Oneida County Landfill (2805)

Alkalinity

	Valid			
·	N	Gamma	Z	p-level
MW-1	15	0.096	0.497	0.619
MW-126	16	-0.08	-0.38	0.702

Hardness

	Valid			
	N	Gamma	Z	p-level
MW-1	15	-0.06	-0.3	0.761
MW-126	16	-0.19	-1.02	0.31

	Valid			
	N	Gamma	Z	p-level
MW-1	69	0.414	4.937	8E-07
MW-126	17	-0.46	-2.53	0.011

B1.4 Portage County Landfill (2966)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
W-10	51	0.555	5.659	2E-08
W-11	51	0.378	3.797	1E-04
W-12	51	-0.24	-2.36	0.018
W-9	51	0.452	4.542	6E-06
W-9P	51	-0.11	-1.16	0.247

Hardness

	Valid			
	N	Gamma	· Z	p-level
W-10	51	0.357	3.629	3E-04
W-11	51	0.132	1.335	0.182
W-12	51	-0.08	-0.84	0.399
W-9	51	0.435	4.426	1E-05
W-9P	51	-0.08	-0.85	0.395

	Valid			
	N	Gamma	Z	p-level
W-10	51	0.059	0.608	0.543
W-11	51	-0.07	- 0.75	0.456
W-12	51	-0.27	-2.8	0.005
W-9	51	0.059	0.595	0.552
W-9P	51	-0.08	-0.81	0.421

B1.5 Grede Foundries Landfill (2974)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
B-3	56	-0.3	-3.18	0.001
B-5	51	-0.6	-6.13	9E-10

Hardness

	Valid			
	N	Gamma	Z	p-level
B-3	56	0.221	2.389	0.017
B-5	52	0.237	2.441	0.015

	Valid			
	N	Gamma	Z	p-level
B-3	56	0.081	0.882	0.378
B-5	51	0.058	0.599	0.549

B1.6 Sauk County Landfill (2978)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
W-30	48	-0.29	-2.89	0.004
W-30A	48	-0.45	-4.45	9E-06
W-31	47	-0.3	-2.97	0.003

Hardness

	Valid			
	N	Gamma	Z	p-level
W-30	48	-0.2	-1.99	0.046
W-30A	48	-0.3	-2.98	0.003
W-31	48	-0.29	-2.87	0.004

	Valid			
	N	Gamma	Z	p-level
W-30	48	-0.22	-2.17	0.03
W-30A	48	-0.21	-2.09	0.036
W-31	48	-0.18	-1.8	0.072

B1.7 City of Richland Center Landfill (3065)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
MW-6	43	0.339	3.152	0.002
MW-7	43	0.02	0.189	0.85
MW-7P	42	0.278	2.558	0.011

Hardness

	Valid			
	N	Gamma	Z	p-level
MW-6	43	-0.01	-0.11	0.915
MW-7	43	0.289	2.717	0.007
MW-7P	43	0.059	0.55	0.582

	Valid			
	N	Gamma	Z	p-level
MW-6	43	0.117	1.069	0.285
MW-7	43	0.249	2.312	0.021
MW-7P	42	0.249	2.29	0.022

B1.8 Juneau County Landfill (3070)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
OW-5	35	0.435	2.777	0.009

Hardness

	Valid			
	N	Gamma	Z	p-level
OW-5	35	0.681	5.34	7E-06

	Valid			
	N	Gamma	Z	p-level
OW-5	35	0.454	2.931	0.006

B1.9 Troy Area Landfill (3090)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
B-1	32	-0.37	-2.79	0.005
B-1B	21	0.077	0.485	0.627
B-2	30	0.197	1.524	0.128

Hardness

	Valid			
	N	Gamma	Z	p-level
B-1	32	-0.22	-1.75	0.081
B-1B	21	0.043	0.272	0.785
B-2	30	0.244	1.882	0.06

	Valid			
	N	Gamma	Z	p-level
B-1	33	-0.55	-4.49	7E-06
B-1B	22	0.214	1.388	0.165
B-2	31	0.137	1.078	0.281

B1.10 Lincoln County Landfill (3141)

Alkalinity

	Valid			
	N	Gamma	Z	p-level
M-4	32	0.218	1.737	0.082
M-9	30	-0.5	-3.83	1E-04

Hardness

	Valid			
	N	Gamma	Z	p-level
M-4	32	0.309	2.457	0.014
M -9	30	0	0	1

	Valid			
	N	Gamma	Z	p-level
M-4	32	0.191	1.513	0.13
M-9	30	-0.07	-0.52	0.605

APPENDIX C AUTOCORRELATION TEST RESULTS

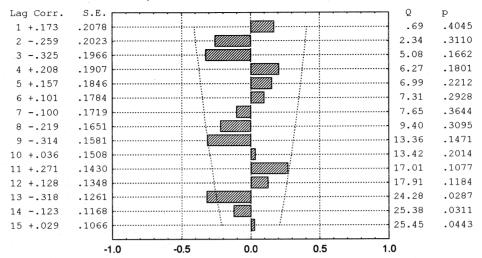
APPENDIX C AUTOCORRELATION TEST RESULTS

Table of Contents

C1.1	Lincoln Landfill (1779)	C-1
C1.2	Pope & Talbot Landfill (2695)	C-4
C1.3	Oneida County Landfill (2805)	C-7
C1.4	Portage County Landfill (2966)	C-10
C1.5	Grede Foundries Landfill (2974)	C-16
C1.6	Sauk County Landfill (2978)	C-19
C1.7	City of Richland Center Landfill (3065)	C-22
C1.8	Juneau County Landfill (3070)	C-25
C1.9	Troy Area Landfill (3090)	C-26
C1.10	Lincoln County Landfill (3141)	C-29

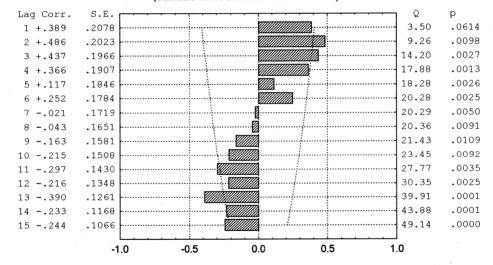
Autocorrelation Function: Lincoln Landfill (1779) MW-1 Alkalinity

(Standard errors are white-noise estimates)



Autocorrelation Function: Lincoln Landfill (1779)

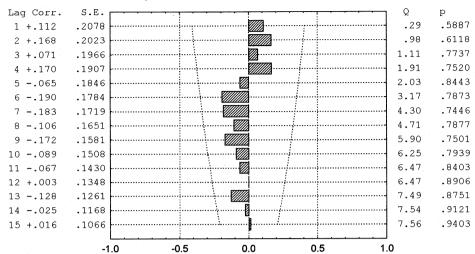
MW-4 Alkalinity



Autocorrelation Function: Lincoln Landfill (1779)

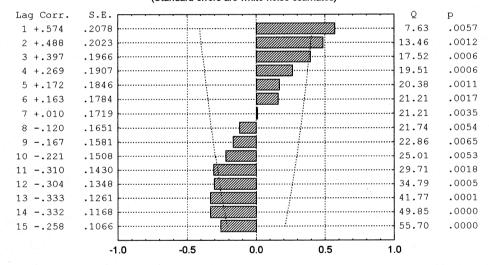
MW-1 Hardness

(Standard errors are white-noise estimates)



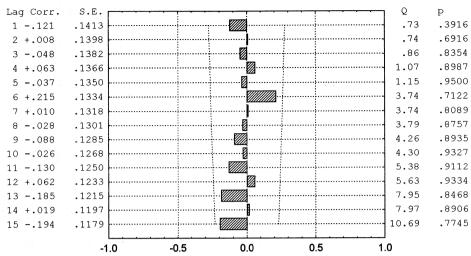
Autocorrelation Function: Lincoln Landfill (1779)

MW-4 Hardness

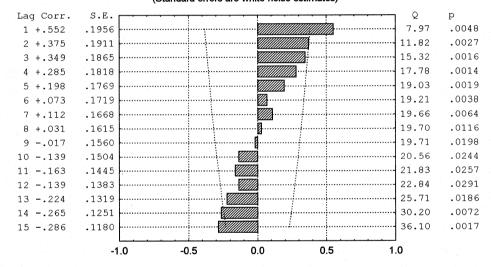


Autocorrelation Function: Lincoln Landfill (1779) MW-1 Specific Conductance

(Standard errors are white-noise estimates)

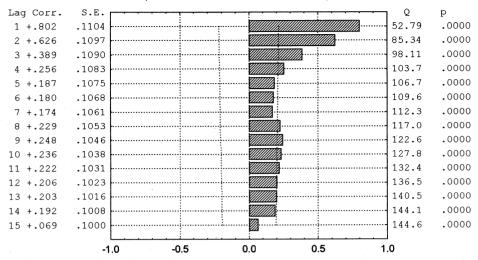


Autocorrelation Function: Lincoln Landfill (1779) MW-4 Specific Conductance



Autocorrelation Function: Pope & Talbot Landfill (2695) W-6 Alkalinity

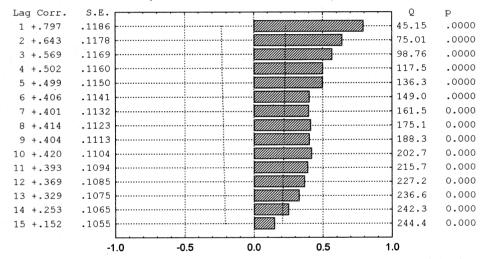
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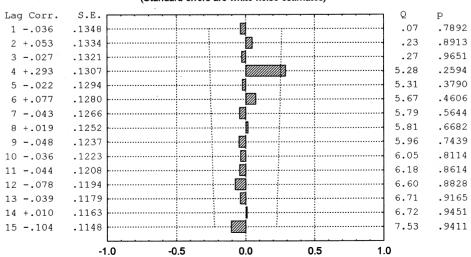
Autocorrelation Function: Pope & Talbot Landfill (2695)

W-13 Alkalinity

(Standard errors are white-noise estimates)

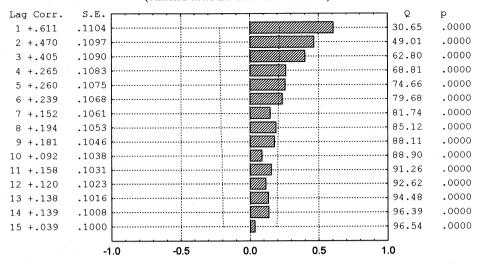


Autocorrelation Function: Pope & Talbot Landfill (2695) W-18 Alkalinity



Autocorrelation Function: Pope & Talbot Landfill (2695) W-6 Hardness

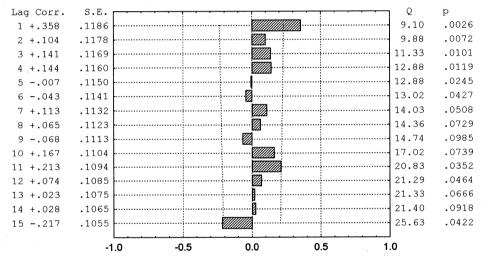
(Standard errors are white-noise estimates)



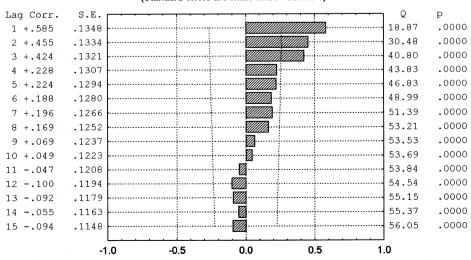
Autocorrelation Function: Pope & Talbot Landfill (2695)

W-13 Hardness

(Standard errors are white-noise estimates)



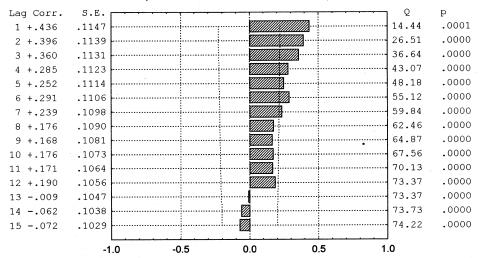
Autocorrelation Function: Pope & Talbot Landfill (2695) W-18 Hardness



Autocorrelation Function: Pope & Talbot Landfill (2695)

W-6 Specific Conductance

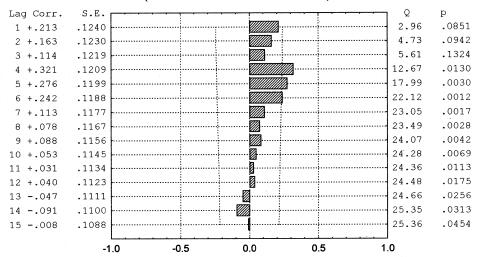
(Standard errors are white-noise estimates)



Autocorrelation Function: Pope & Talbot Landfill (2695)

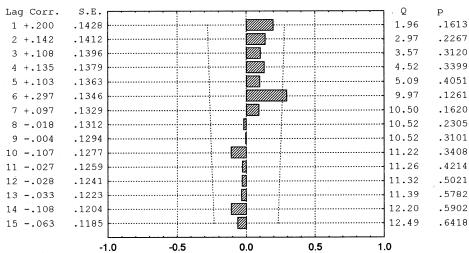
W-13 Specific Conductance

(Standard errors are white-noise estimates)



Autocorrelation Function: Pope & Talbot Landfill (2695)

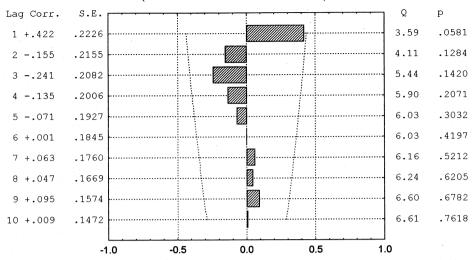
W-18 Specific Conductance



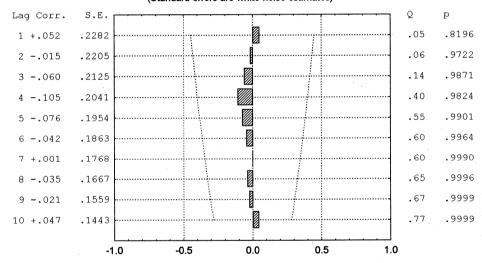
Autocorrelation Function: Oneida County Landfill (2805)

MW-1 Alkalinity

(Standard errors are white-noise estimates)

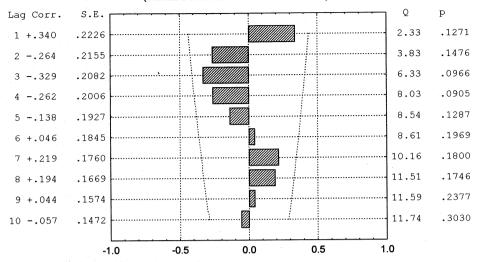


Autocorrelation Function: Oneida County Landfill (2805) MW-126 Alkalinity

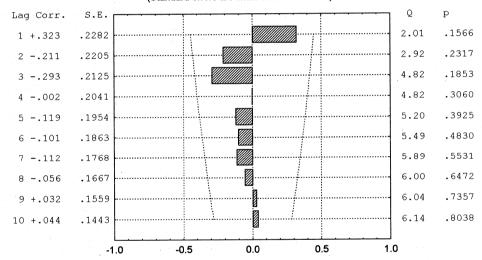


Autocorrelation Function: Oneida County Landfill (2805) MW-1 Hardness

(Standard errors are white-noise estimates)

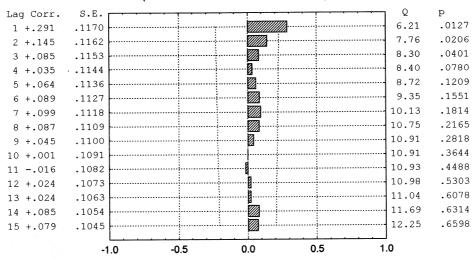


Autocorrelation Function: Oneida County Landfill (2805) MW-126 Hardness

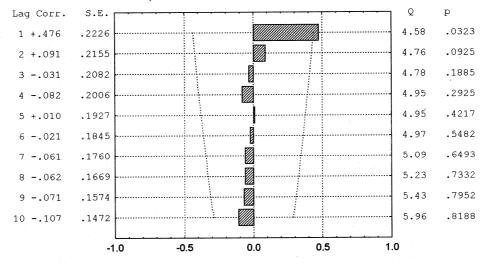


Autocorrelation Function: Oneida County Landfill (2805) MW-1 Specific Conductance

(Standard errors are white-noise estimates)

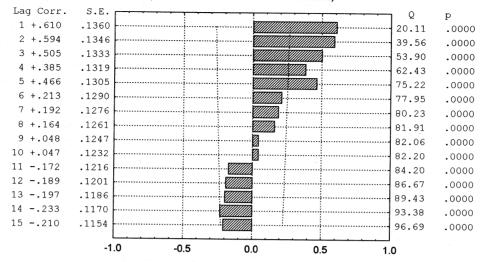


Autocorrelation Function: Oneida County Landfill (2805) MW-126 Specific Conductance



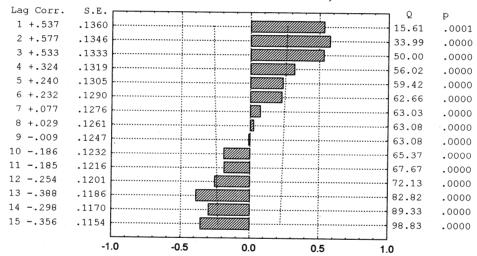
Autocorrelation Function: Portage County Landfill (2966) W-9 Alkalinity

(Standard errors are white-noise estimates)

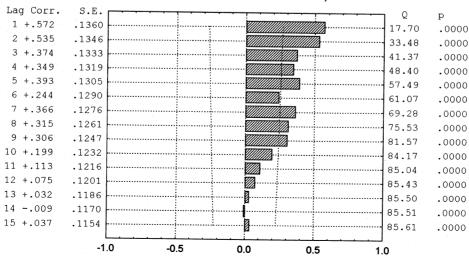


Autocorrelation Function: Portage County Landfill (2966) W-9P Alkalinity

(Standard errors are white-noise estimates)

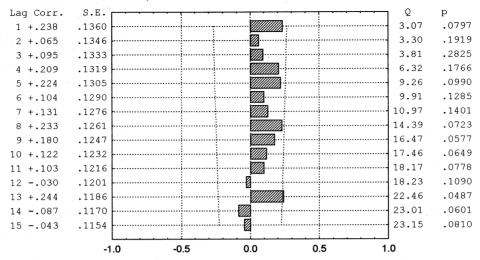


Autocorrelation Function: Portage County Landfill (2966) W-10 Alkalinity

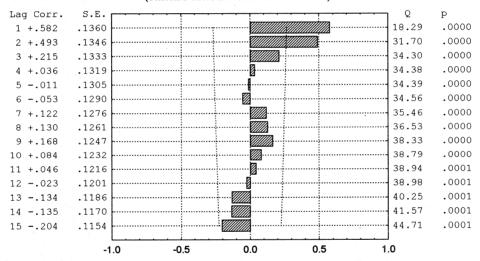


Autocorrelation Function: Portage County Landfill (2966) W-11 Alkalinity

(Standard errors are white-noise estimates)

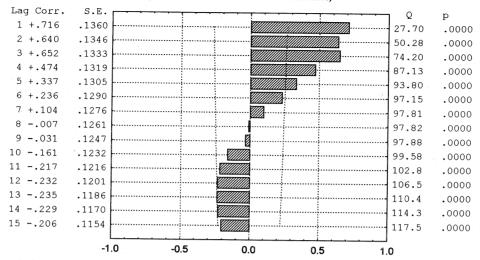


Autocorrelation Function: Portage County Landfill (2966) W-12 Alkalinity



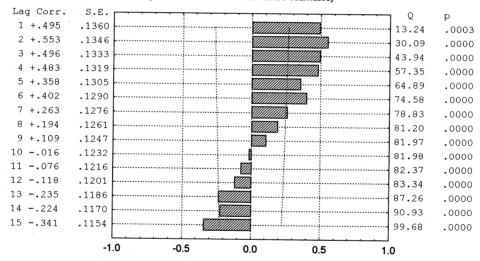
Autocorrelation Function: Portage County Landfill (2966) W-9 Hardness

(Standard errors are white-noise estimates)

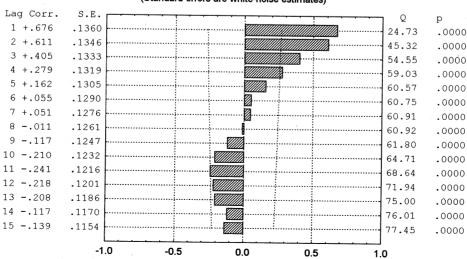


Autocorrelation Function: Portage County Landfill (2966) W-9P Hardness

(Standard errors are white-noise estimates)

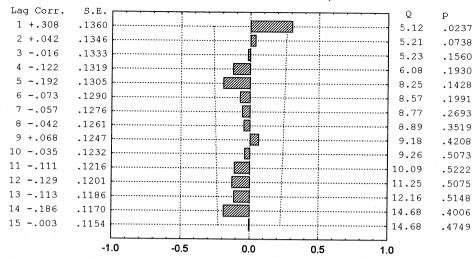


Autocorrelation Function: Portage County Landfill (2966) W-10 Hardness

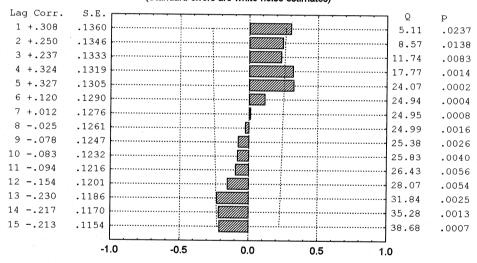


Autocorrelation Function: Portage County Landfill (2966) W-11 Hardness

(Standard errors are white-noise estimates)

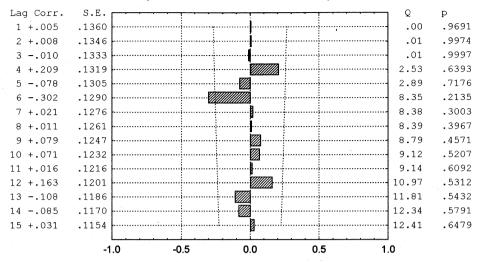


Autocorrelation Function: Portage County Landfill (2966) W-12 Hardness



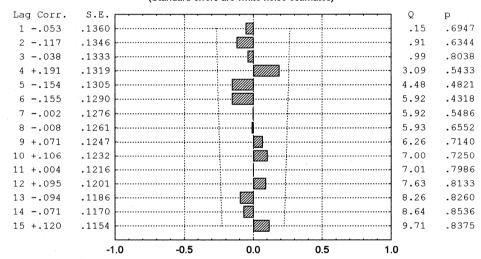
Autocorrelation Function: Portage County Landfill (2966) W-9 Specific Conductance

(Standard errors are white-noise estimates)

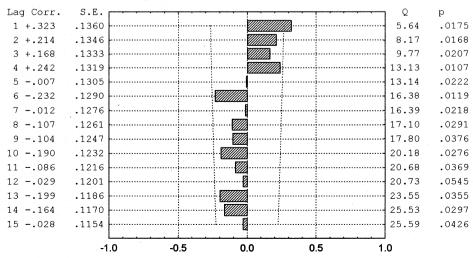


Autocorrelation Function: Portage County Landfill (2966) W-9P Specific Conductance

(Standard errors are white-noise estimates)



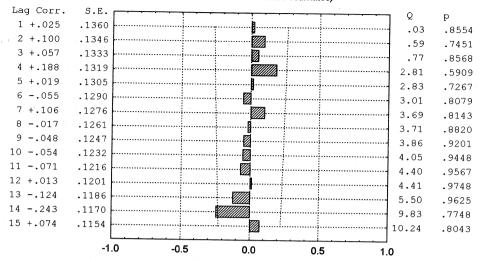
Autocorrelation Function: Portage County Landfill (2966) W-10 Specific Conductance



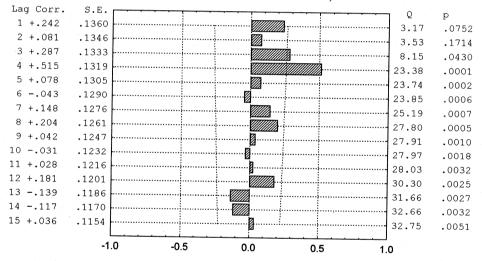
Autocorrelation Function: Portage County Landfill (2966)

W-11 Specific Conductance

(Standard errors are white-noise estimates)

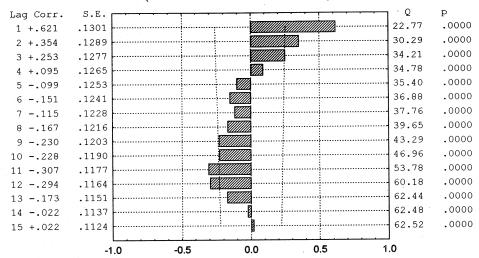


Autocorrelation Function: Portage County Landfill (2966) W-12 Specific Conductance

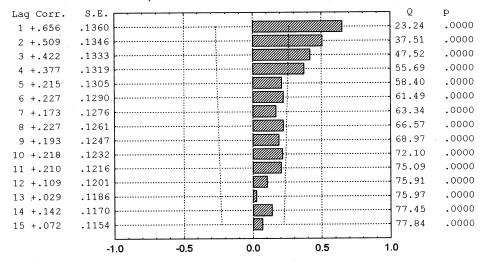


Autocorrelation Function: Grede Foundries Landfill (2974) B-3 Alkalinity

(Standard errors are white-noise estimates)

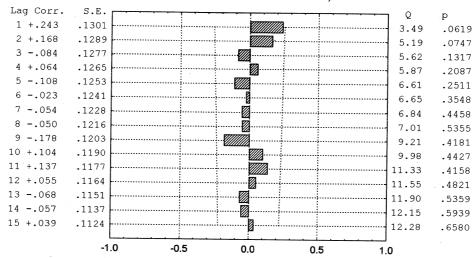


Autocorrelation Function: Grede Foundries Landfill (2974) B-5 Alkalinity

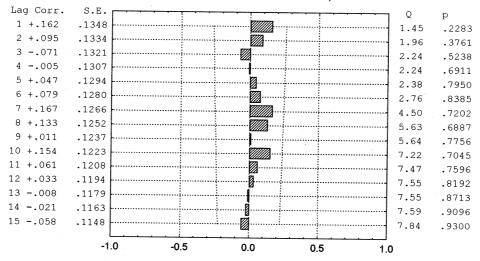


Autocorrelation Function: Grede Foundries Landfill (2974) B-3 Hardness

(Standard errors are white-noise estimates)

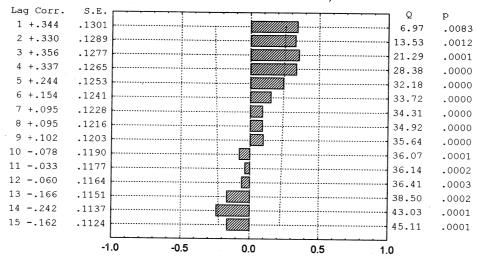


Autocorrelation Function: Grede Foundries Landfill (2974) B-5 Hardness

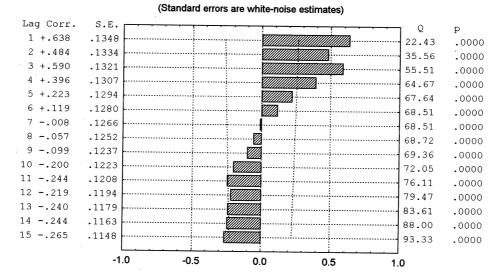


Autocorrelation Function: Grede Foundries Landfill (2974) B-3 Specific Conductance

(Standard errors are white-noise estimates)

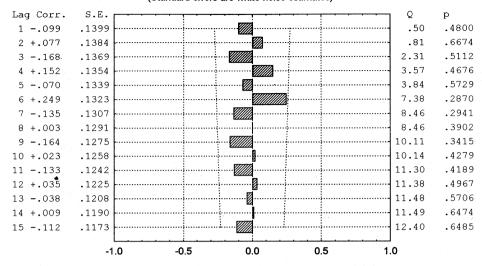


Autocorrelation Function: Grede Foundries Landfill (2974) B-5 Specific Conductance



Autocorrelation Function: Sauk County Landfill (2978) W-30 Alkalinity

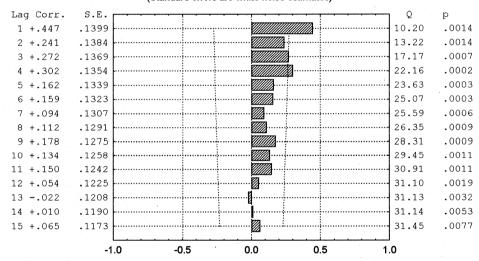
(Standard errors are white-noise estimates)



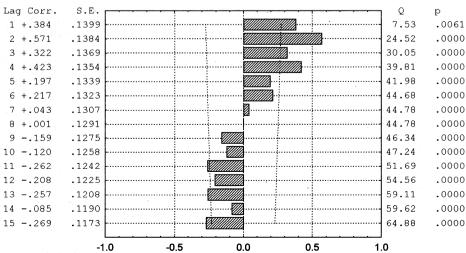
Autocorrelation Function: Sauk County Landfill (2978)

W-30A Alkalinity

(Standard errors are white-noise estimates)

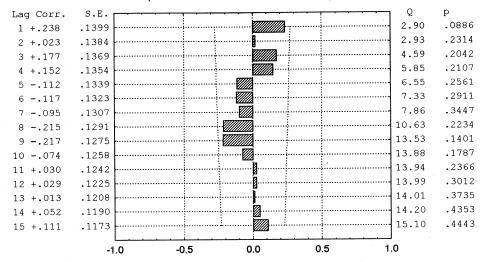


Autocorrelation Function: Sauk County Landfill (2978) W-31 Alkalinity



Autocorrelation Function: Sauk County Landfill (2978) W-30 Hardness

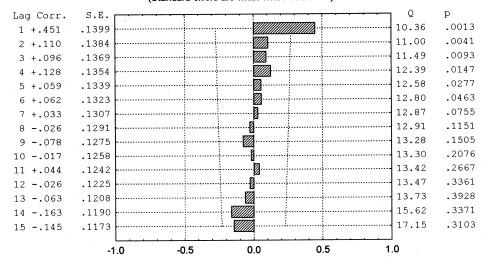
(Standard errors are white-noise estimates)



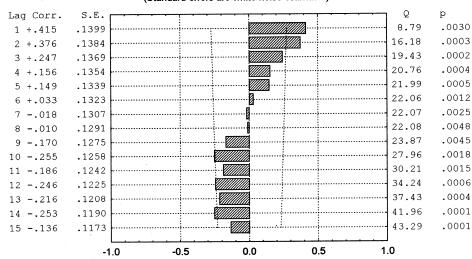
Autocorrelation Function: Sauk County Landfill (2978)

W-30A Hardness

(Standard errors are white-noise estimates)

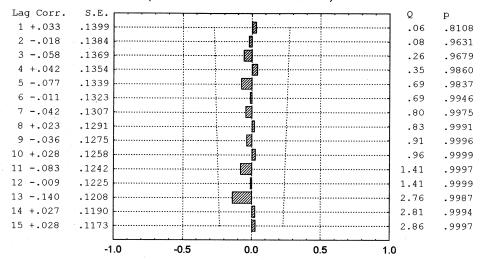


Autocorrelation Function: Sauk County Landfill (2978) W-31 Hardness



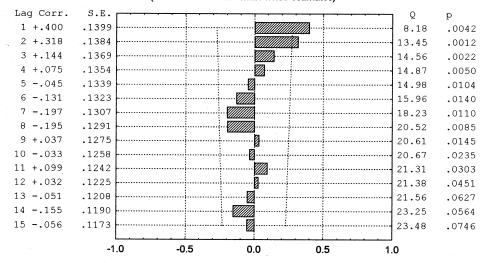
Autocorrelation Function: Sauk County Landfill (2978) W-30 Specific Conductance

(Standard errors are white-noise estimates)

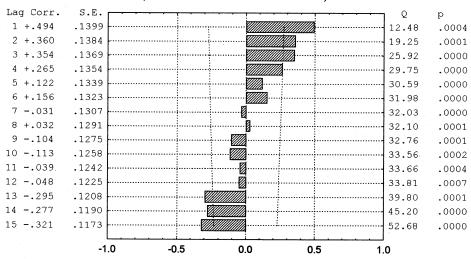


Autocorrelation Function: Sauk County Landfill (2978) W-30A Specific Conductance

(Standard errors are white-noise estimates)



Autocorrelation Function: Sauk County Landfill (2978) W-31 Specific Conductance

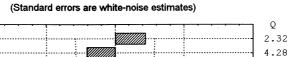


Autocorrelation Function: City of Richland Center (3065)

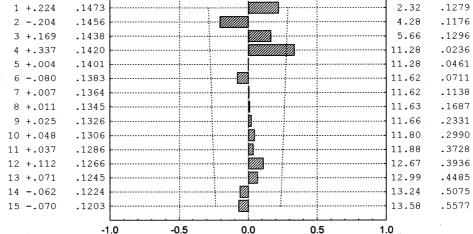
MW-6 Alkalinity

Lag Corr.

S.E.



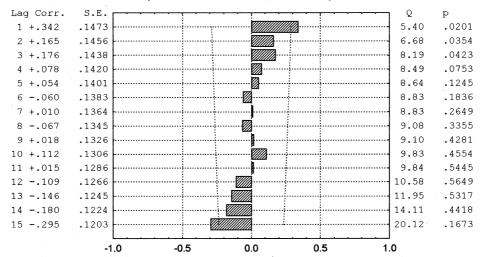
р



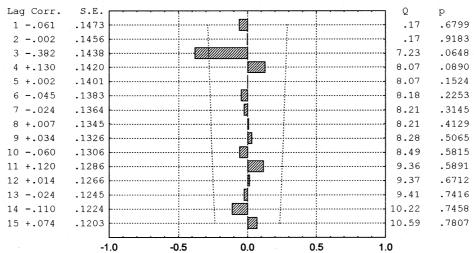
Autocorrelation Function: City of Richland Center (3065)

MW-7 Alkalinity

(Standard errors are white-noise estimates)

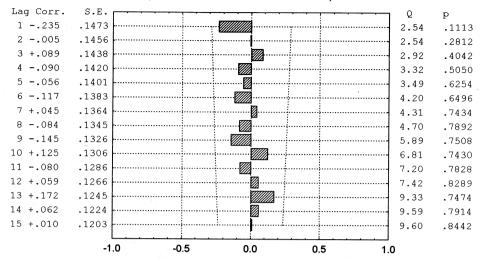


Autocorrelation Function: City of Richland Center (3065) MW-7P Alkalinity



Autocorrelation Function: City of Richland Center (3065) MW-6 Hardness

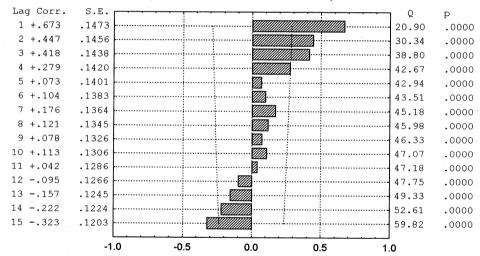
(Standard errors are white-noise estimates)



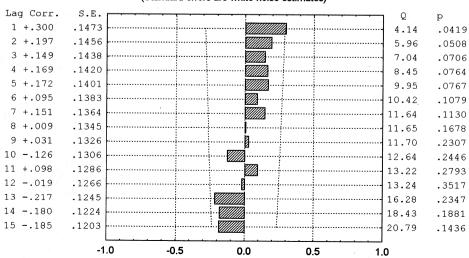
Autocorrelation Function: City of Richland Center (3065)

MW-7 Hardness

(Standard errors are white-noise estimates)

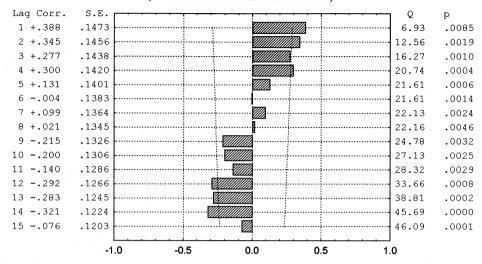


Autocorrelation Function: City of Richland Center (3065) MW-7P Hardness



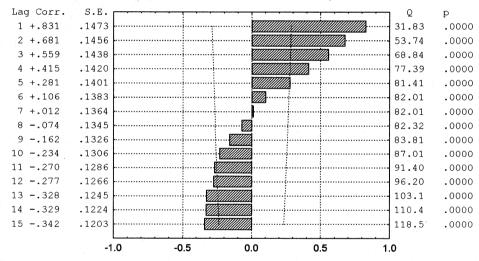
Autocorrelation Function: City of Richland Center (3065) MW-6 Specific Conductance

(Standard errors are white-noise estimates)

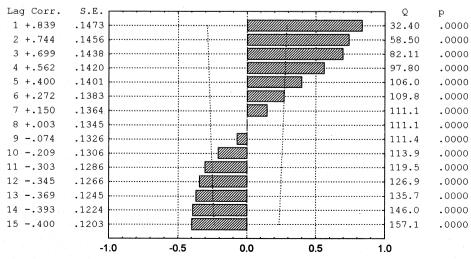


Autocorrelation Function: City of Richland Center (3065) MW-7 Specific Conductance

(Standard errors are white-noise estimates)

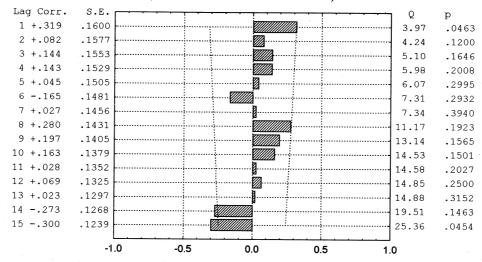


Autocorrelation Function: City of Richland Center (3065) MW-7P Specific Conductance



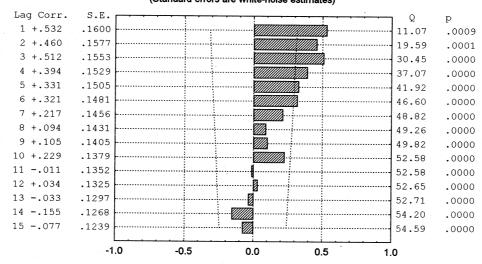
Autocorrelation Function: Juneau County Landfill (3070) OW-5 Alkalinity

(Standard errors are white-noise estimates)

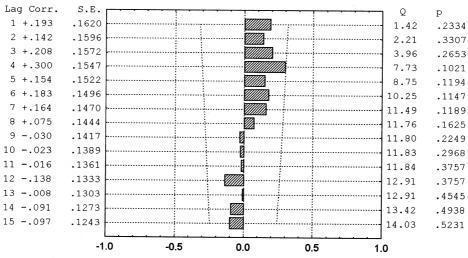


Autocorrelation Function: Juneau County Landfill (3070) OW-5 Hardness

(Standard errors are white-noise estimates)

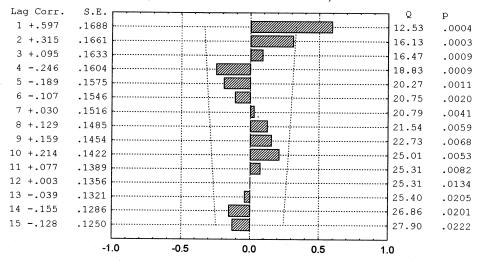


Autocorrelation Function: Juneau County Landfill (3070) OW-5 Specific Conductance



Autocorrelation Function: Troy Area Landfill (3090) B-1 Alkalinity

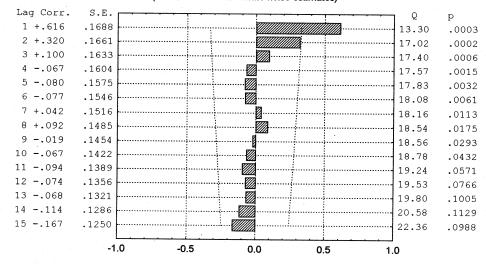
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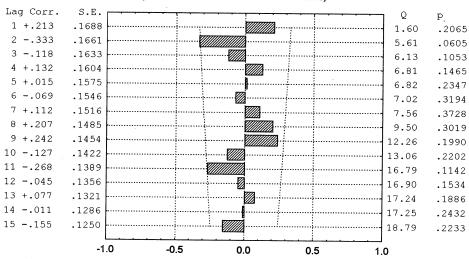
Autocorrelation Function: Troy Area Landfill (3090)

B-1B Alkalinity

(Standard errors are white-noise estimates)

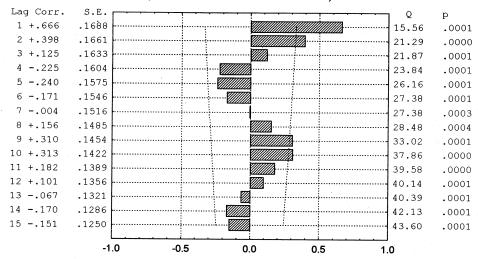


Autocorrelation Function: Troy Area Landfill (3090) B-2 Alkalinity



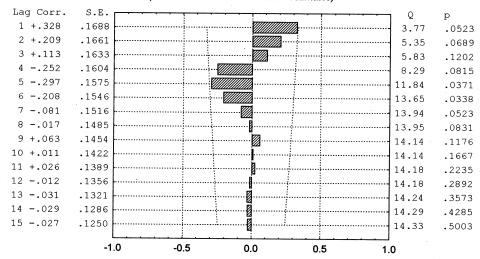
Autocorrelation Function: Troy Area Landfill (3090) B-1 Hardness

(Standard errors are white-noise estimates)

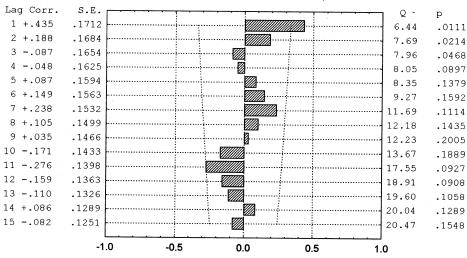


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(Standard errors are white-noise estimates)

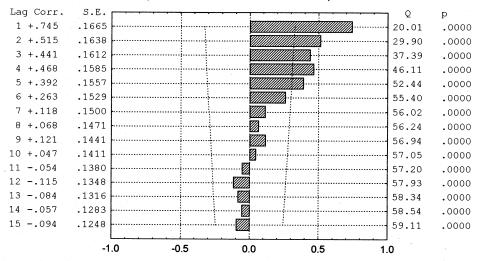


Autocorrelation Function: Troy Area Landfill (3090) B-2 Hardness



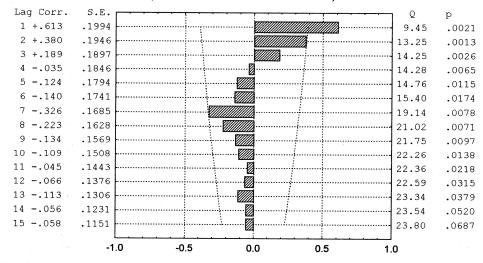
Autocorrelation Function: Troy Area Landfill (3090) B-1 Specific Conductance

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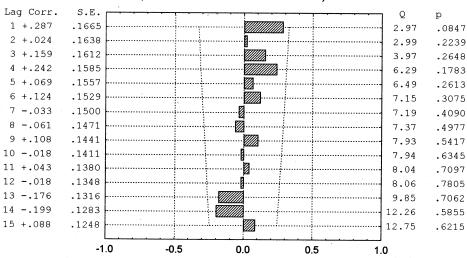


Autocorrelation Function: Troy Area Landfill (3090) B-1B Specific Conductance

(Standard errors are white-noise estimates)

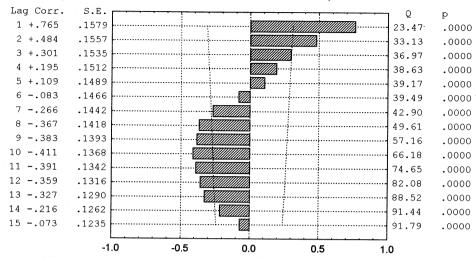


Autocorrelation Function: Troy Area Landfill (3090) B-2 Specific Conductance

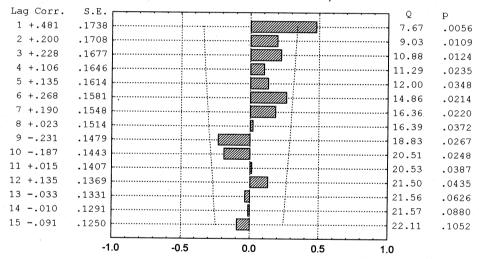


Autocorrelation Function: Lincoln County Landfill (3141) M-4 Alkalinity

(Standard errors are white-noise estimates)

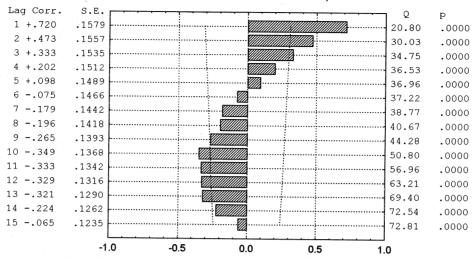


Autocorrelation Function: Lincoln County Landfill (3141) M-9 Alkalinity



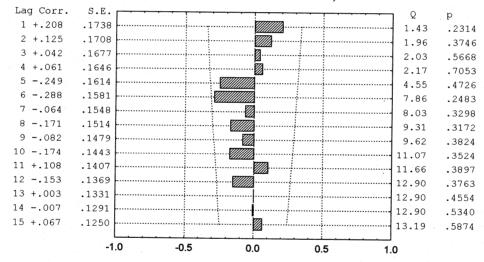
Autocorrelation Function: Lincoln County Landfill (3141) M-4 Hardness

(Standard errors are white-noise estimates)



Autocorrelation Function: Lincoln County Landfill (3141)

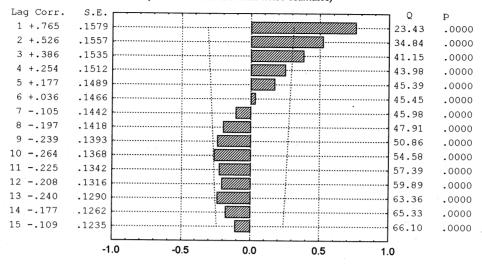
M-9 Hardness



Autocorrelation Function: Lincoln County Landfill (3141)

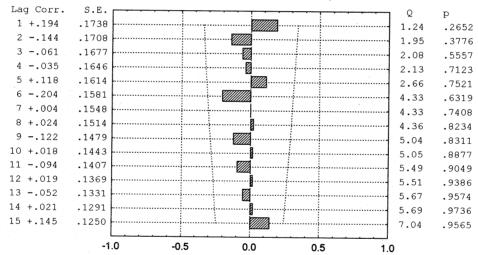
M-4 Specific Conductance

(Standard errors are white-noise estimates)



Autocorrelation Function: Lincoln County Landfill (3141)

M-9 Specific Conductance



APPENDIX D DUMPStat RESULTS: PORTAGE COUNTY CASE STUDY

APPENDIX D DUMPStat RESULTS: PORTAGE COUNTY CASE STUDY

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2966 Case Study: Table 1

Summary Statistics and Intermediate Computations for Intra-Well Prediction Limits

Constituent	Units	Well	N	Mean	SD	Factor	Limit
Alkalinity	mg/L	W-10	8	236.925	8.505	3.180	263.967
	mg/L	W-11	8	246.175	5.847	3.180	264.767
	mg/L	W-12	8	178.975	5.762	3.180	197.297
	mg/L	W-17	8	255.063	12.611	3.180	295.160
	mg/L	W-18	8	133.425	7.492	3.180	157.247
	mg/L	W-20	8	169.975	6.804	3.180	191.608
	mg/L	W-20P	8	181.863	36.830	3.180	298.968
	mg/L	W-21	8	179.225	5.555	3.180	196.889
	mg/L	W-21P	8	202.875	9.156	3.180	231.989
	mg/L	W-22	8	162.713	4.847	3.180	178.123
	mg/L	W-22P	8	178.738	7.835	3.180	203.649
	mg/L	W-23	8	147.662	37.709	3.180	267.565
	mg/L	W-23P	8	157.438	2.129	3.180	164.206
	mg/L	W-9	8	230.787	3.388	3.180	241.559
	mg/L	W-9P	8	225.287	4.934	3.180	240.977
Conductivity	MICR	W-10	8	517.875	42.549	3.180	653.166
	MICR	W-11	8	489.375	31.332	3.180	589.000
	MICR	W-12	8	374.875	26.107	3.180	457.885
	MICR	W-17	8	493.625	59.457	3.180	682.677
	MICR	W-18	8	265.750	9.603	3.180	296.284
	MICR	W-20	8	338.000	16.759	3.180	391.287
	MICR	W-20P	8	406.250	25.600	3.180	487.649
	MICR	W-21	8	361.250	23.414	3.180	435.698
	MICR	W-21P	8	410.000	24.495	3.180	487.885
	MICR	W-22	8	337.125	23.000	3.180	410.256
	MICR	W-22P	8	366.625	25.629	3.180	448.116
	MICR	W-23	8	302.500	63.552	3.180	504.573
	MICR	W-23P	8	323.500	21.119	3.180	390.650
	MICR	W-9	8	451.875	26.984	3.180	537.674
	MICR	W-9P	8	463.125	33.374	3.180	569.243
Hardness	mg/L	W-10	8	269.663	7.325	3.180	292.953
	mg/L	W-11	8	265.525	7.539	3.180	289.496
	mg/L	W-12	8	190.213	2.801	3.180	199.119
	mg/L	W-17	8	269.512	13.027	3.180	310.934
	mg/L	W-18	8	140.488	10.463	3.180	173.756
	mg/L	W-20	8	177.800	3.283	3.180	188.238
	mg/L	W-20P	8	211.762	3.913	3.180	224.203
	mg/L	W-21	8	188.600	4.837	3.180	203.979
	mg/L	W-21P	8	216.787	6.418	3.180	237.194

^{* -} Insufficient Data

^{** -} Detection Frequency < 25%

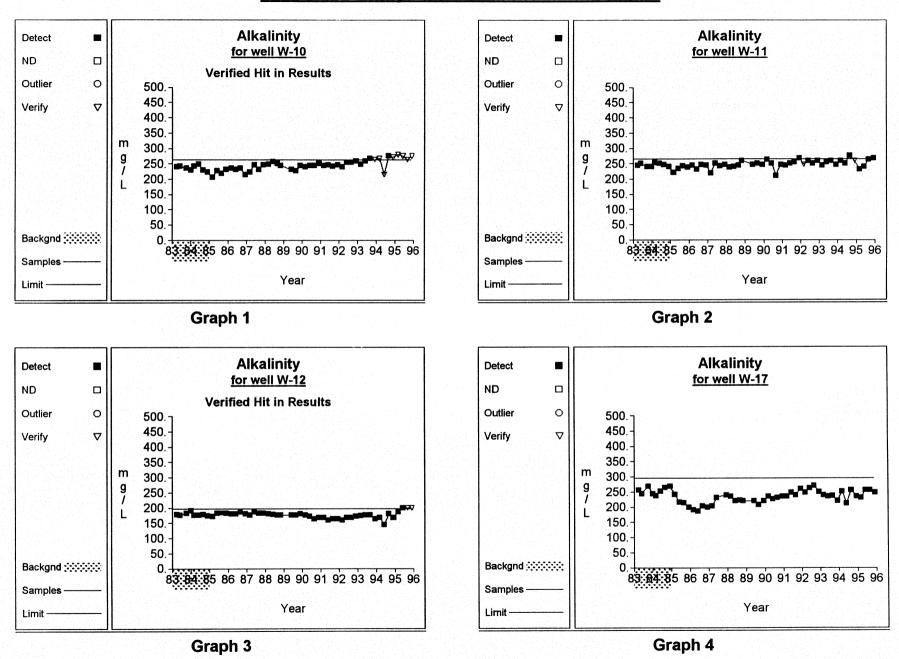
^{*** -} Zero Variance

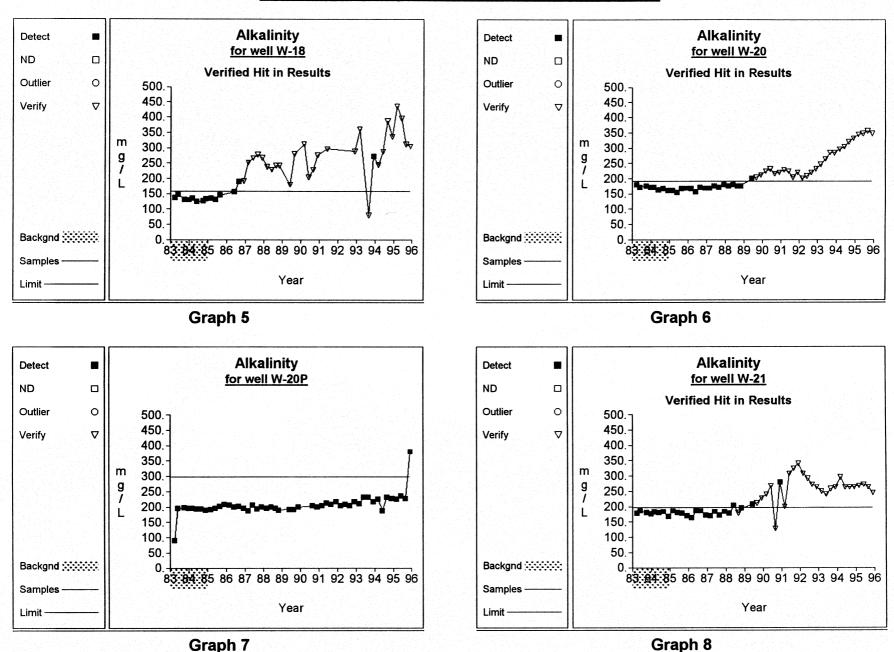
2966 Case Study: Table 1 - Continued

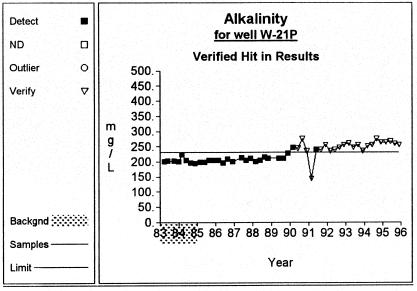
Summary Statistics and Intermediate Computations for Intra-Well Prediction Limits

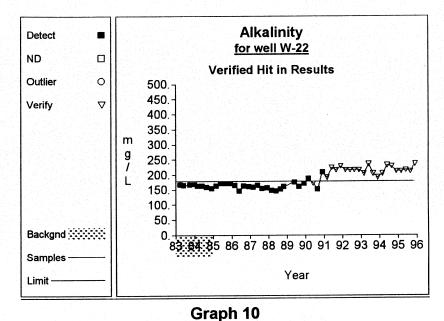
Constituent	Units	Well	N	Mean	SD	Factor	Limit
Hardness	mg/L	W-22	8	170.525	5.152	3.180	186.908
	mg/L	W-22P	8	186.075	6.602	3.180	207.067
	mg/L	W-23	8	154.075	36.476	3.180	270.056
	mg/L	W-23P	8	170.787	5.405	3.180	187.975
	mg/L	W-9	8	243.600	2.209	3.180	250.624
	mg/L	W-9P	8	247.488	2.382	3.180	255.061

^{* -} Insufficient Data
** - Detection Frequency < 25%
*** - Zero Variance









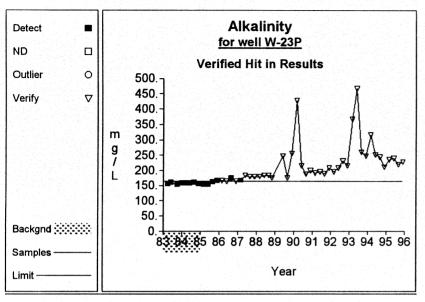
Graph 9

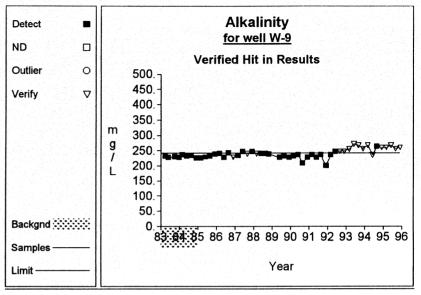
Alkalinity Detect for well W-22P ND **Verified Hit in Results** 0 Outlier 500.-450. ∇ Verify 400. 350. 300. 250. 200 150 100. 50. 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit

Alkalinity Detect for well W-23 ND **Verified Hit in Results** Outlier Ó 1000. 900. ∇ Verify 800. 700. 600. 500. 400. 300. 200 100 83:84:85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit -

Graph 11

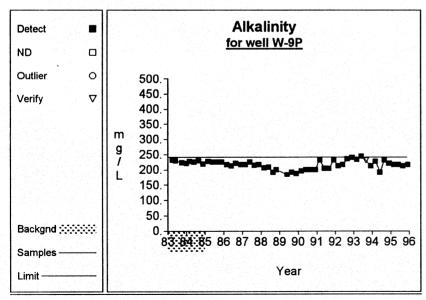
Graph 12



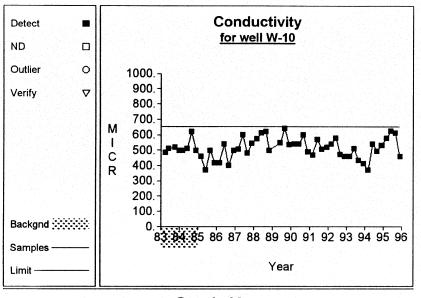


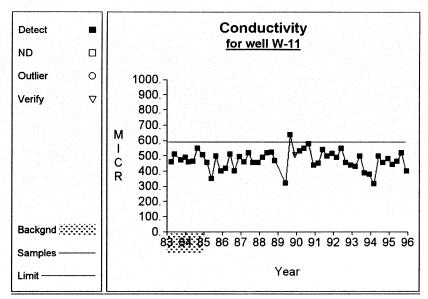
Graph 13

Graph 14



Graph 15

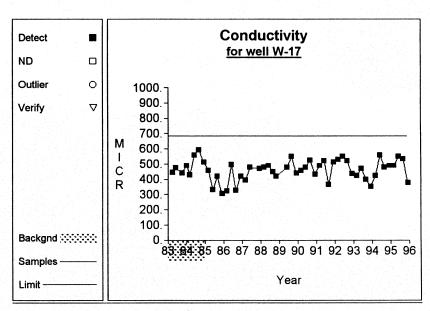




Graph 16

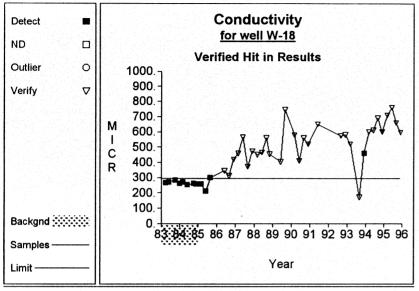
Conductivity Detect for well W-12 ND Outlier 0 500.-450. ∇ Verify 400 350 300 250. С 200. 150. 100. 50. 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit

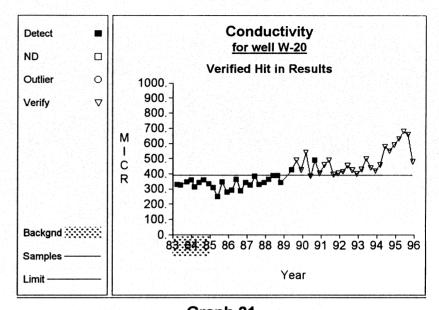
Graph 17



Graph 18

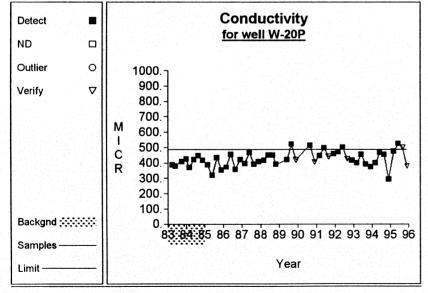
Graph 19

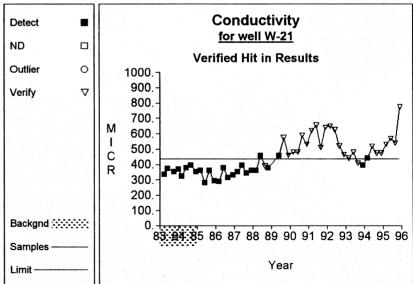




Graph 20

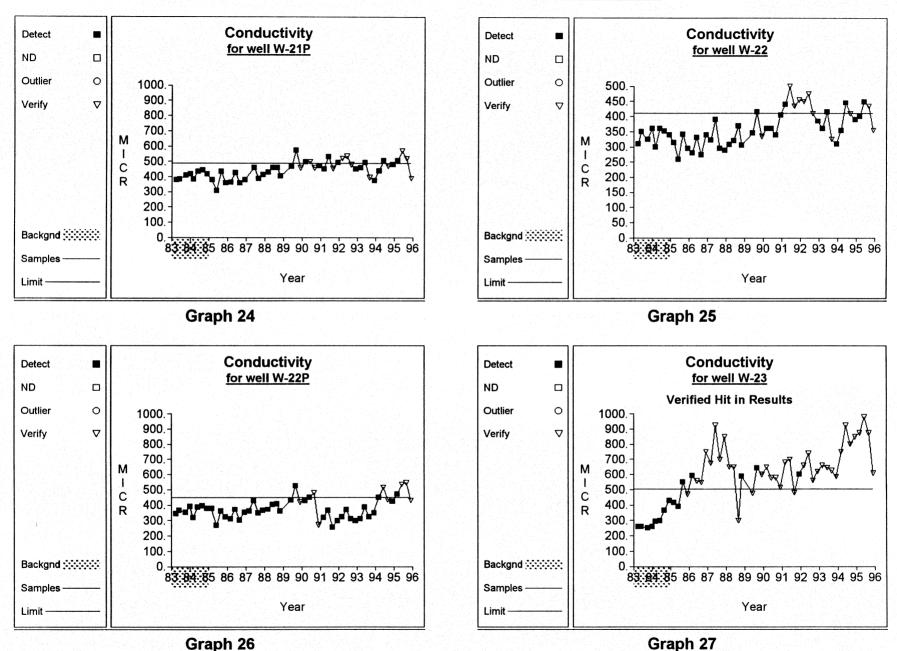
Graph 21

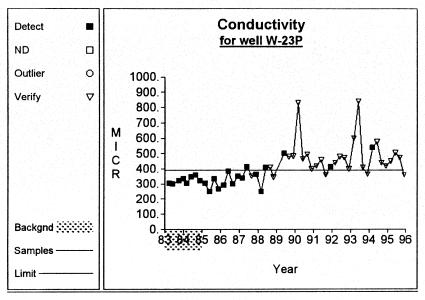


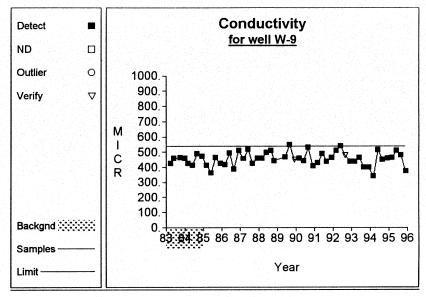


Graph 22

Graph 23

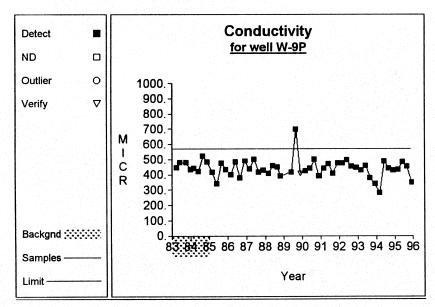




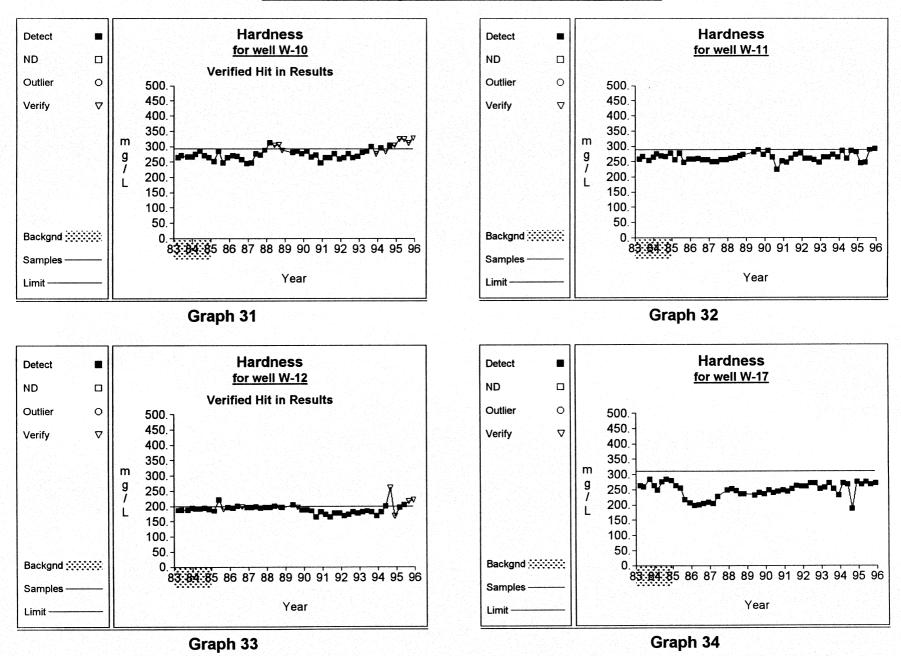


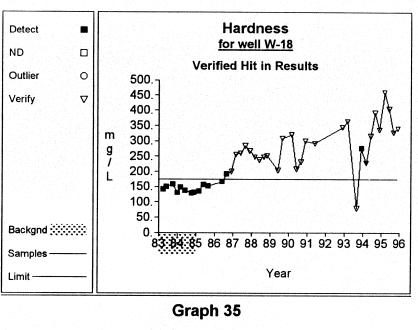
Graph 28

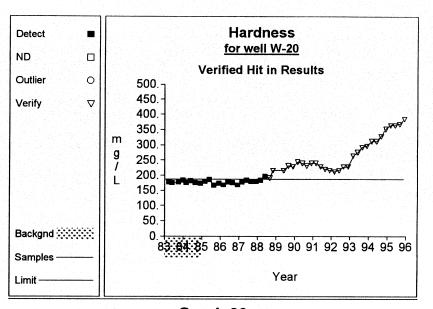
Graph 29



Graph 30

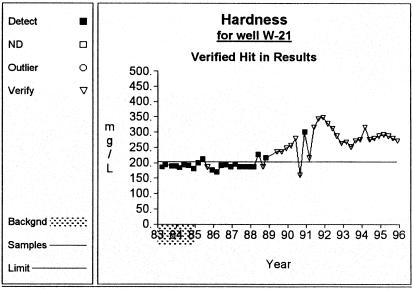






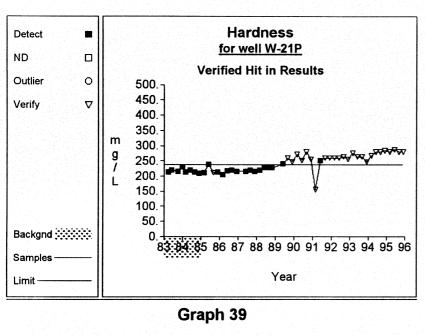
Hardness Detect for well W-20P ND **Verified Hit in Results** Outlier 0 500.-450. Verify ∇ 400. 350. 300 250 200 150 100 50. Backgnd ::::::::: 83:84:85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit

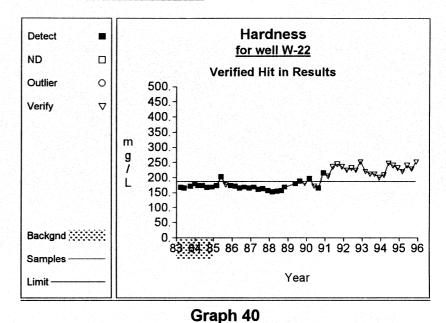
Graph 36



Graph 37

Graph 38



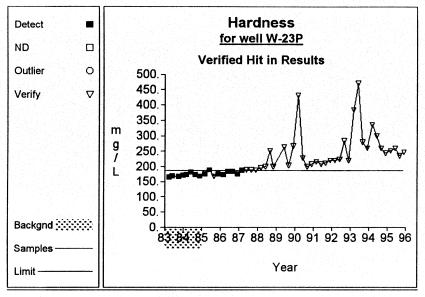


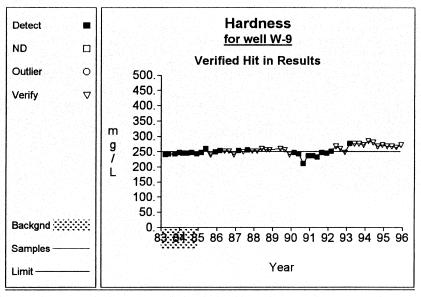
Hardness Detect for well W-22P ND **Verified Hit in Results** Outlier 0 500. 450. ∇ Verify 400. 350 300. 250. 200 150 100 50. Backgnd ::::::::: 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit

Hardness Detect for well W-23 ND Significant Trend / Verified Hit Outlier 0 1000. 900. ∇ Verify 800. 700. 600. 500. 400. 300 200 100 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Samples Year Limit

Graph 41

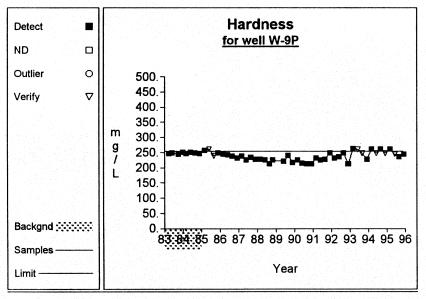
Graph 42





Graph 43

Graph 44



Graph 45

2966 Case Study: Table 1

Summary Statistics and Intermediate Computations for Combined Shewart-Cusum Control Charts

Constituent	Units	Well	N	Mean	SD	S(i-1)	S(i)	Limit
Alkalinity	mg/L	W-10	8	236.925	8.505	463.760	494.330	275.196
	mg/L	W-11	8	246.175	5.847	285.844	301.822	272.487
	mg/L	W-12	8	178.975	5.762	197.500	212.763	204.906
	mg/L	W-17	8	255.063	12.611	255.063	255.063	311.811
	mg/L	W-18	8	133.425	7.492	3730.831	3893.914	167.139
	mg/L	W-20	8	169.975	6.804	2174.613	2345.834	200.592
	mg/L	W-20P	8	181.863	36.830	236.401	397.708	347.597
	mg/L	W-21	8	179.225	5.555	2165.375	2224.594	204.224
	mg/L	W-21P	8	202.875	9.156	1015.028	1058.996	244.079
	mg/L	W-22	8	162.713	4.847	1055.089	1123.530	184.522
	mg/L	W-22P	8	178.738	7.835	709.589	815.017	213.993
	mg/L	W-23	8	147.662	37.709	7522.044	7784.672	317.355
	mg/L	W-23P	8	157.438	2.129	2614.357	2682.791	167.017
	mg/L	W-9	8	230.787	3.388	546.337	572.162	246.032
	mg/L	W-9P	8	225.287	4.934	225.287	225.287	247.492
Conductivity	MICR	W-10	8	517.875	42.549	651.603	551.179	709.345
	MICR	W-11	8	489.375	31.332	489.375	489.375	630.369
	MICR	W-12	8	374.875	26.107	392.912	374.875	492.355
	MICR	W-17	8	493.625	59.457	493.625	493.625	761.181
	MICR	W-18	8	265.750	9.603	7649.165	7969.812	308.963
	MICR	W-20	8	338.000	16.759	3617.754	3742.995	413.415
	MICR	W-20P	8	406.250	25.600	598.401	546.551	521.450
	MICR	W-21	8	361.250	23.414	3893.659	4282.995	466.613
	MICR	W-21P	8	410.000	24.495	1562.143	1517.648	520.227
	MICR	W-22	8	337.125	23.000	1315.134	1310.010	440.623
	MICR	W-22P	8	366.625	25.629	873.102	910.848	481.955
	MICR	W-23	8	302.500	63.552	12081.316	12325.264	588.484
	MICR	W-23P	8	323.500	21.119	4161.176	4176.557	418.534
	MICR	W-9	8	451.875	26.984	484.157	451.875	573.302
	MICR	W-9P	8	463.125	33.374	463.125	463.125	613.309
Hardness	mg/L	W-10	8	269.663	7.325	487.789	538.801	302.625
	mg/L	W-11	8	265.525	7.539	280.461	299.397	299.450
	mg/L	W-12	8	190.213	2.801	227.171	254.158	202.818
	mg/L	W-17	8	269.512	13.027	269.512	269.512	328.134
	mg/L	W-18	8	140.488	10.463	3731.975	3921.025	187.571
	mg/L	W-20	8	177.800	3.283	2467.149	2670.066	192.573
	mg/L	W-20P	8	211.762	3.913	867.784	900.109	229.369
	mg/L	W-21	8	188.600	4.837	2401.807	2480.371	210.365
	mg/L	W-21P	8	216.787	6.418	1199.035	1255.830	245.668

^{* -} Insufficient Data

^{** -} Detection Frequency < 25%

^{*** -} Zero Variance

2966 Case Study: Table 1 - Continued

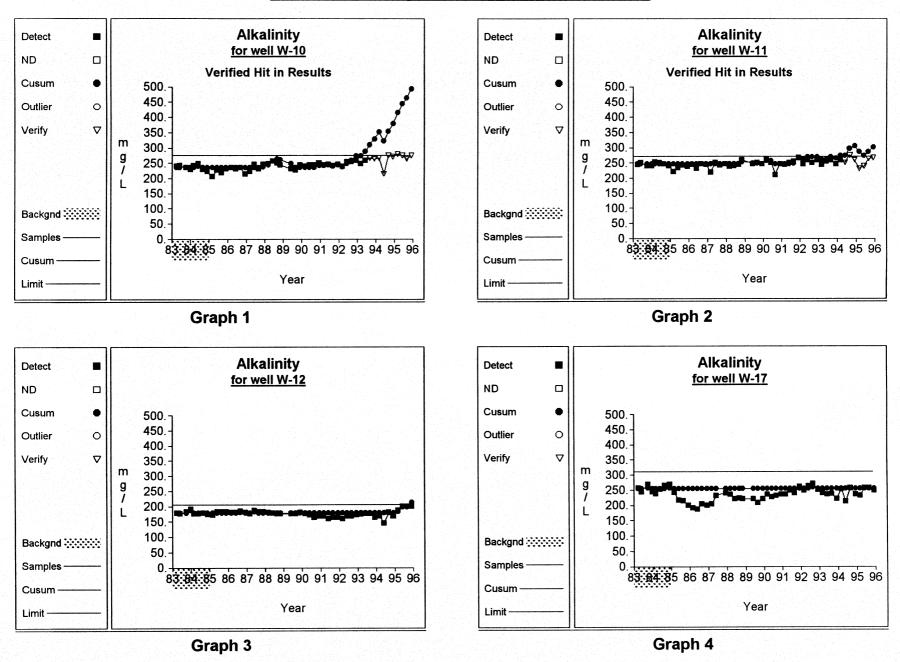
Summary Statistics and Intermediate Computations for Combined Shewart-Cusum Control Charts

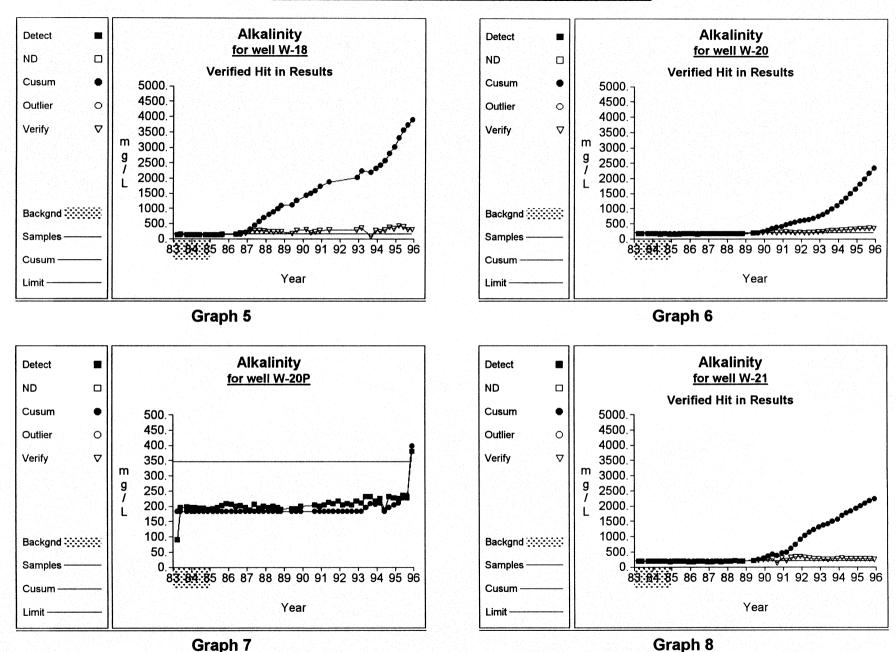
Constituent	Units	Well	N	Mean	SD	S(i-1)	S(i)	Limit
Hardness	mg/L	W-22	8	170.525	5.152	1190.588	1266.911	193.711
	mg/L	W-22P	8	186.075	6.602	737.335	864.658	215.784
	mg/L	W-23	8	154.075	36.476	7919.931	8213.380	318.217
	mg/L	W-23P	8	170.787	5.405	2682.071	2753.878	195.112
	mg/L	W-9	8	243.600	2.209	621.855	648.046	253.541
	mg/L	W-9P	8	247.488	2.382	247.488	247.488	258.205

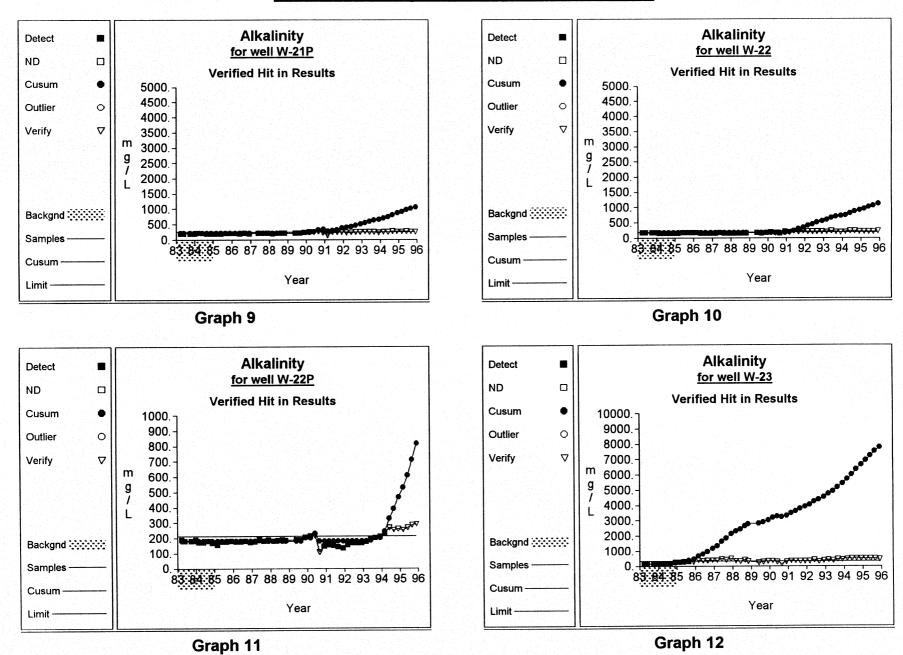
^{* -} Insufficient Data

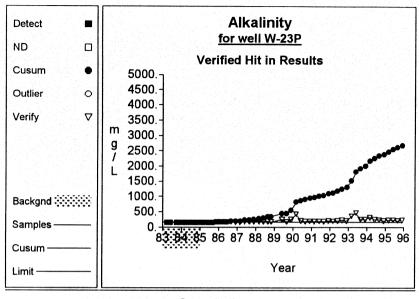
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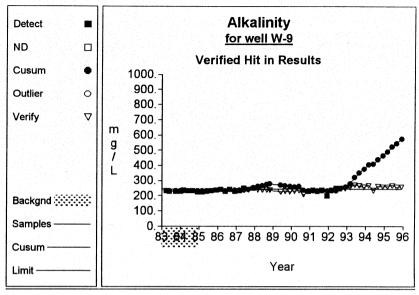
^{*** -} Zero Variance





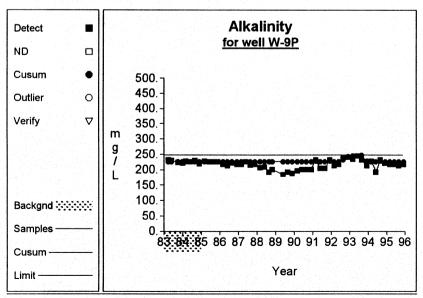




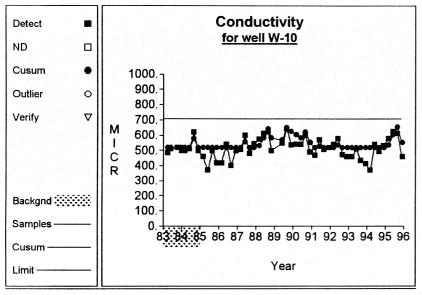


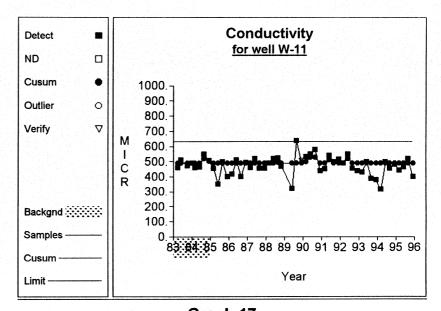
Graph 13

Graph 14



Graph 15

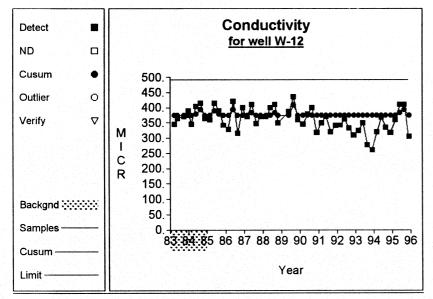


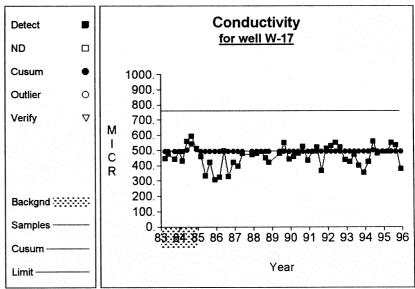


Graph 16

Graph 17

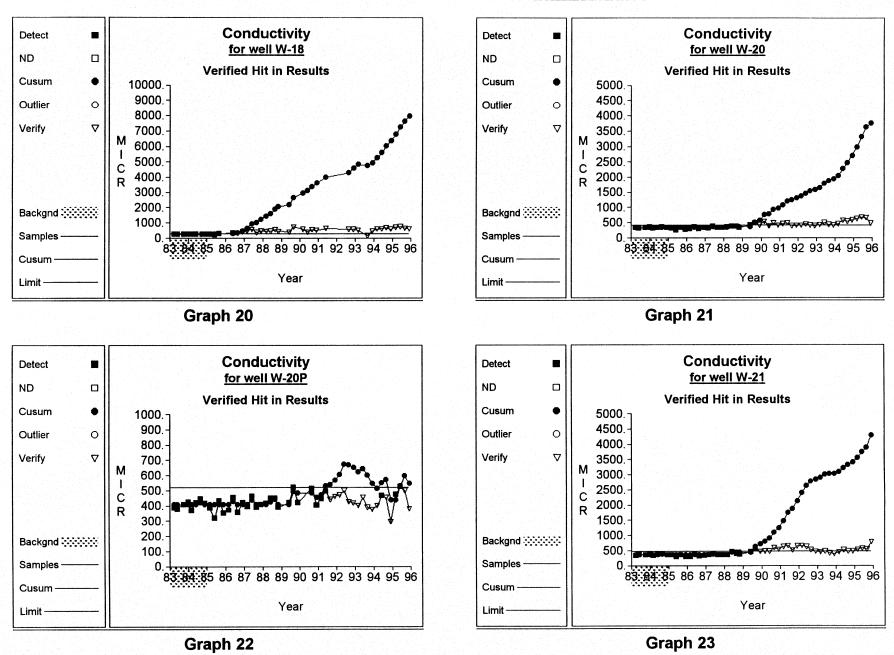
stect ■ Cond

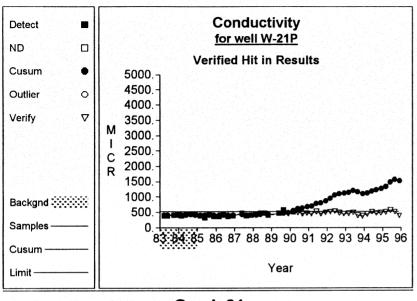


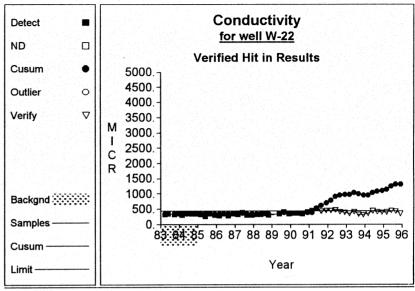


Graph 18

Graph 19



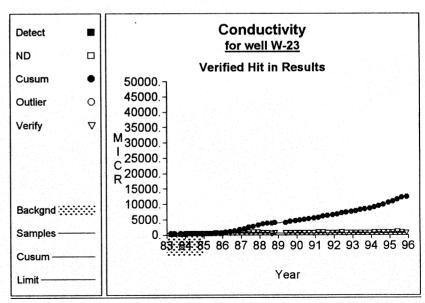




Graph 24

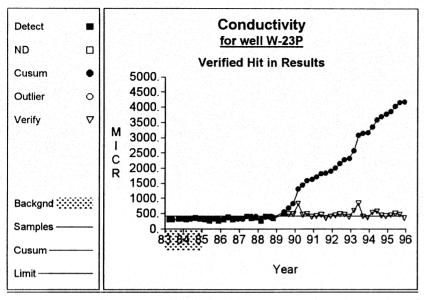
Conductivity Detect for well W-22P ND **Verified Hit in Results** Cusum 1000. 900. Outlier 0 800. Verify 700. 600. 500 C 400 300 200 100 Samples 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Cusum Year Limit

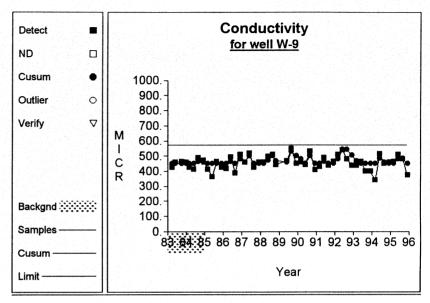
Graph 25



Graph 26

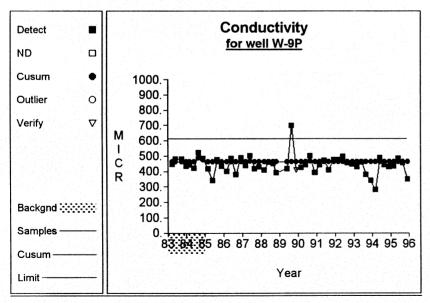
Graph 27



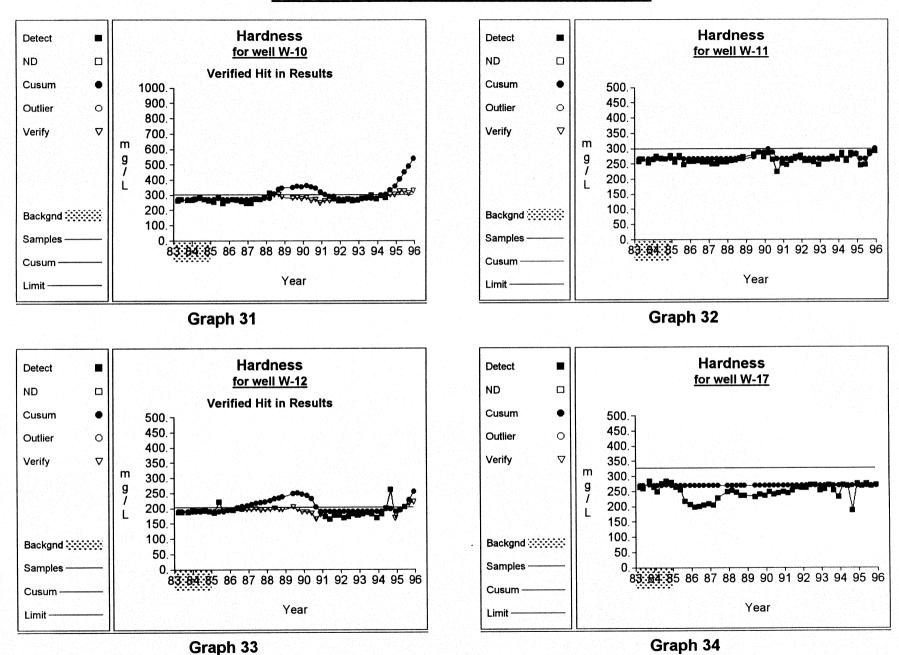


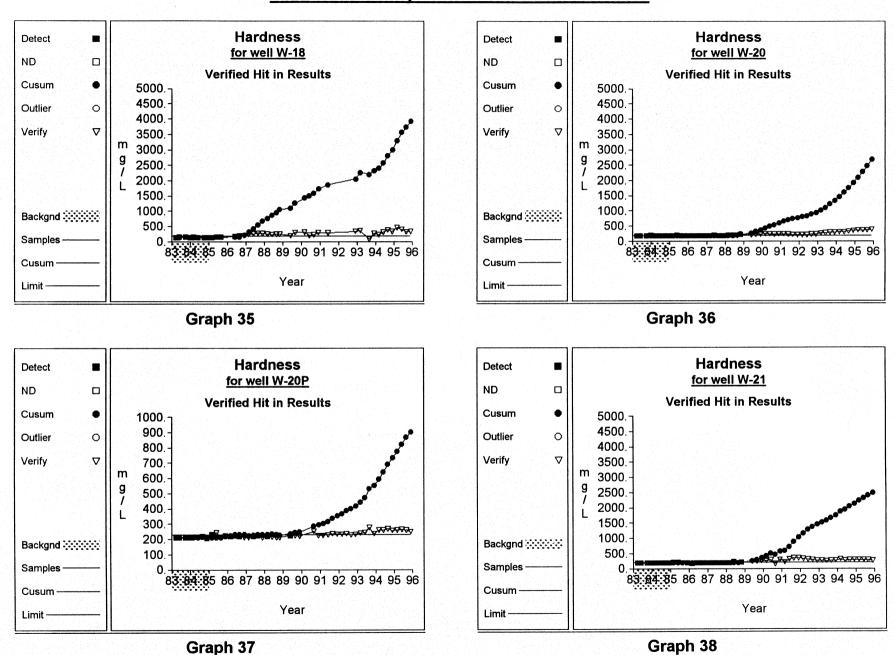
Graph 28

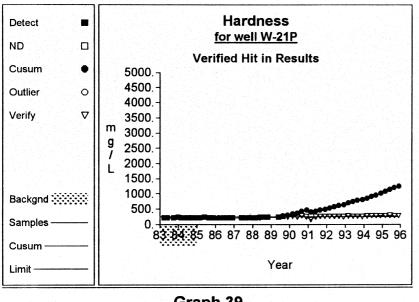
Graph 29

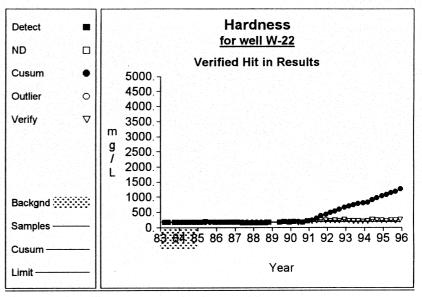


Graph 30



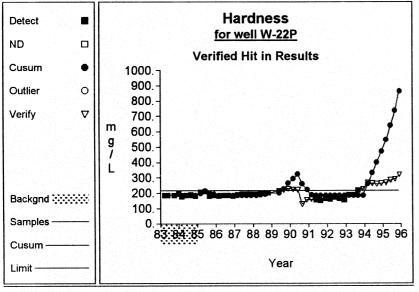


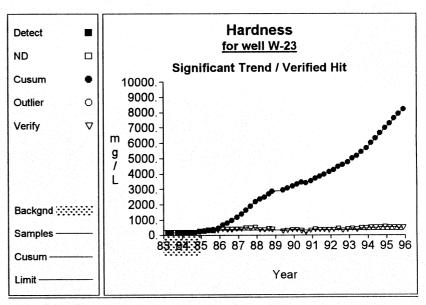




Graph 39

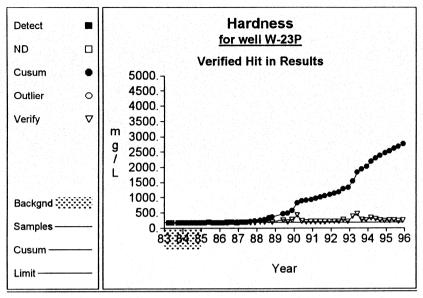
Graph 40

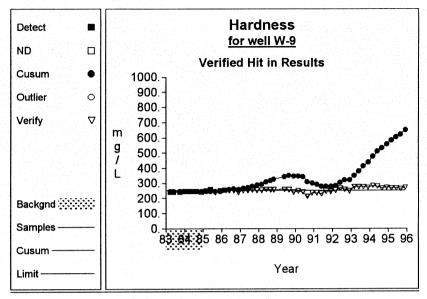




Graph 41

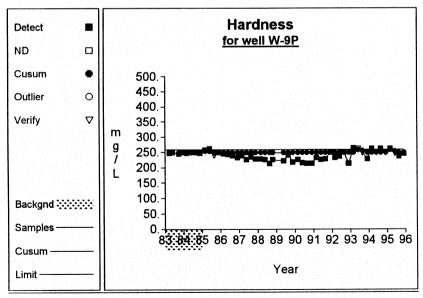
Graph 42





Graph 43

Graph 44



Graph 45

2966 Case Study: Table 1

Summary Statistics and Intermediate Computations for Intra-Well Prediction Limits

Constituent	Units	Well	N	Mean	SD	Factor	Limit
Alkalinity	mg/L	W-10	25	235.496	12.211	2.541	266.525
•	mg/L	W-11	25	242.456	9.041	2.541	265.428
	mg/L	W-12	25	179.752	4.625	2.541	191.503
	mg/L	W-17	24	228.771	24.283	2.551	290.712
	mg/L	W-18	23	183.148	55.121	2.562	324.347
	mg/L	W-20	25	170.792	9.461	2.541	194.833
	mg/L	W-20P	25	192.756	21.824	2.541	248.211
	mg/L	W-21	25	181.352	10.310	2.541	207.550
	mg/L	W-21P	24	204.583	7.217	2.551	222.992
	mg/L	W-22	25	160.908	8.092	2.541	181.471
	mg/L	W-22P	25	179.516	8.227	2.541	200.422
	mg/L	W-23	25	281.812	112.164	2.541	566.820
	mg/L	W-23P	25	170.020	19.355	2.541	219.202
	mg/L	W-9	25	234.612	6.527	2.541	251.198
	mg/L	W-9P	25	217.052	11.706	2.541	246.797
Conductivity	MICR	W-10	25	509.320	65.588	2.541	675.978
•	MICR	W-11	25	466.400	54.160	2.541	604.021
	MICR	W-12	25	375.360	29.590	2.541	450.548
	MICR	W-17	24	445.792	71.721	2.551	628.740
	MICR	W-18	23	355.478	106.569	2.562	628.466
	MICR	W-20	25	338.560	38.396	2.541	436.124
	MICR	W-20P	25	405.400	35.342	2.541	495.205
	MICR	W-21	25	360.040	43.316	2.541	470.106
	MICR	W-21P	24	407.417	39.410	2.551	507.944
	MICR	W-22	25	323.200	32.087	2.541	404.733
	MICR	W-22P	25	363.520	38.352	2.541	460.971
	MICR	W-23	25	500.600	194.489	2.541	994.797
	MICR	W-23P	25	338.320	56.851	2.541	482.779
	MICR	W-9	25	453.560	39.050	2.541	552.787
	MICR	W-9P	25	441.560	41.999	2.541	548.280
Hardness	mg/L	W-10	25	272.652	17.988	2.541	318.360
	mg/L	W-11	25	261.848	8.922	2.541	284.520
	mg/L	W-12	25	195.068	6.947	2.541	212.720
	mg/L	W-17	24	241.129	27.921	2.551	312.351
	mg/L	W-18	23	190.604	53.857	2.562	328.566
	mg/L	W-20	25	181.736	12.301	2.541	212.994
	mg/L	W-20P	25	216.884	7.977	2.541	237.154
	mg/L	W-21	25	193.792	14.958	2.541	231.801
	mg/L	W-21P	24	219.346	8.719	2.551	241.587

^{* -} Insufficient Data
** - Detection Frequency < 25%
*** - Zero Variance

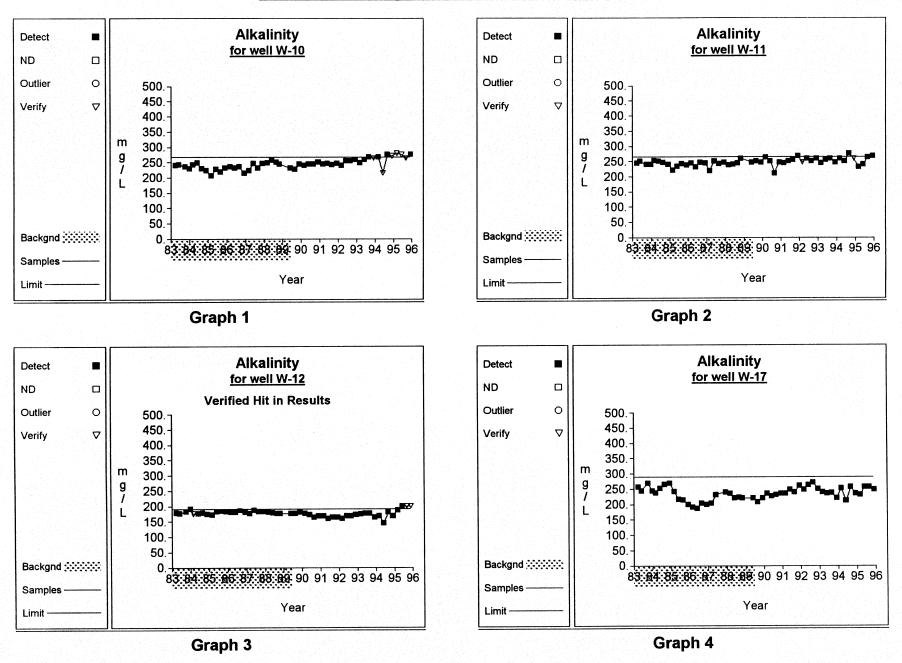
2966 Case Study: Table 1 - Continued

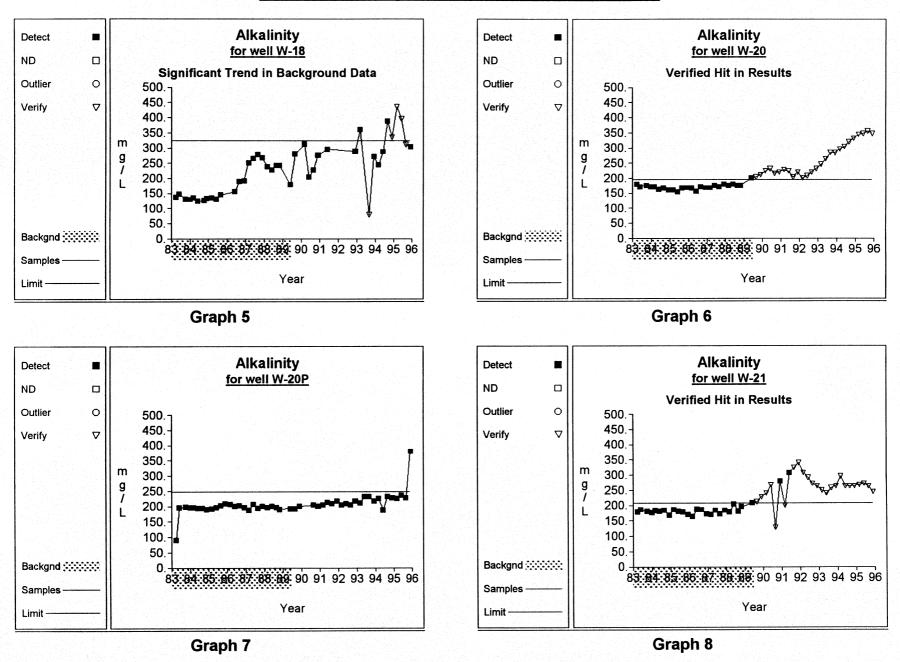
Summary Statistics and Intermediate Computations for Intra-Well Prediction Limits

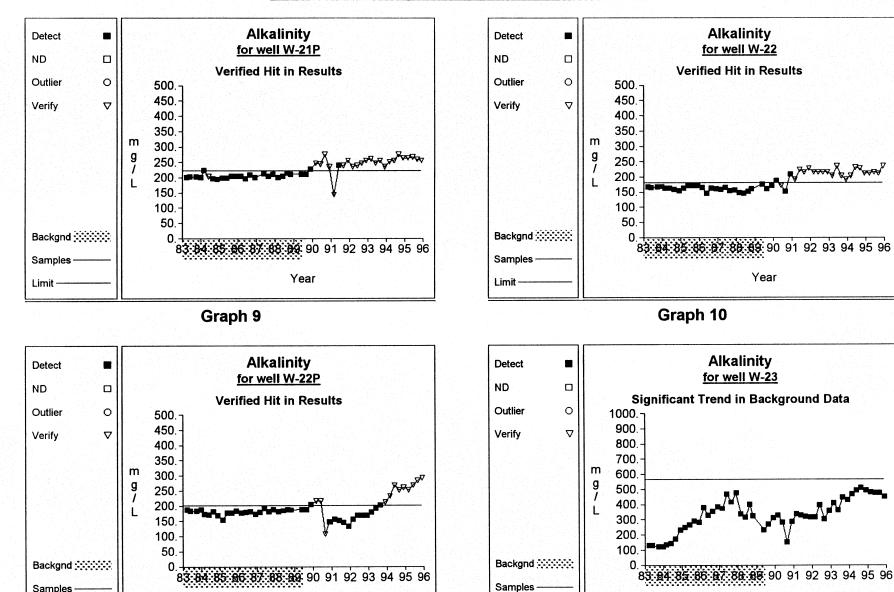
Constituent	Units	Well	N	Mean	SD	Factor	Limit
Hardness	mg/L mg/L mg/L mg/L mg/L	W-22 W-22P W-23 W-23P W-9	25 25 25 25 25 25 25	168.528 190.584 293.104 186.452 249.512 238.876	10.430 9.517 113.196 23.986 6.490 12.225	2.541 2.541 2.541 2.541 2.541 2.541	195.031 214.767 580.736 247.400 266.003 269.939

^{* -} Insufficient Data** - Detection Frequency < 25%

^{*** -} Zero Variance







Graph 11

Year

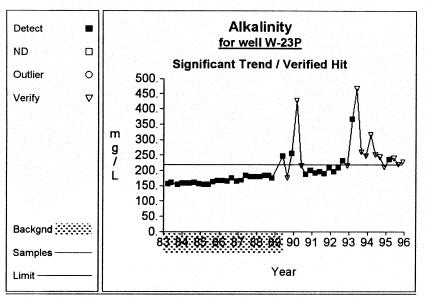
Samples

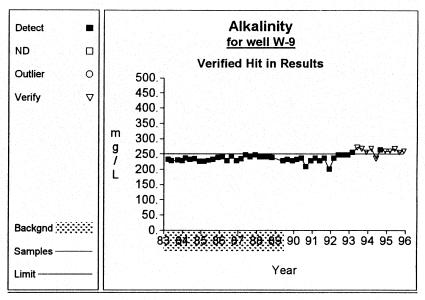
Limit

Graph 12

Year

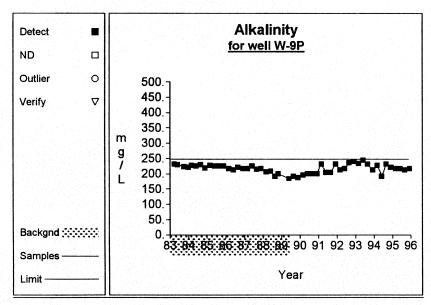
Limit



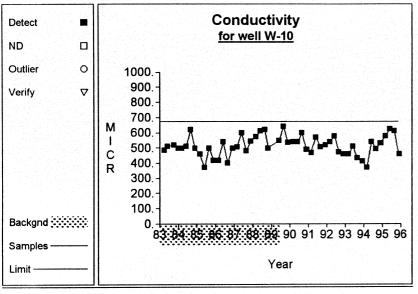


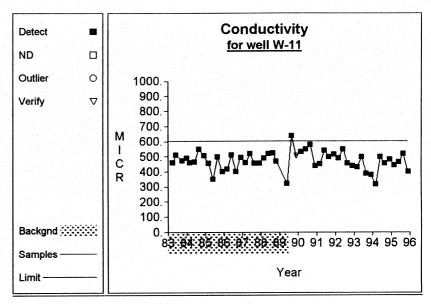
Graph 13

Graph 14



Graph 15

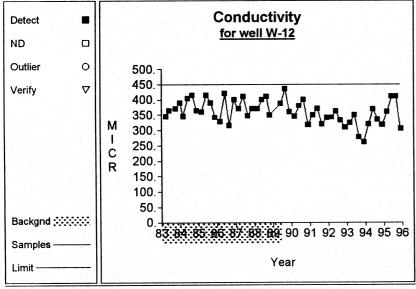


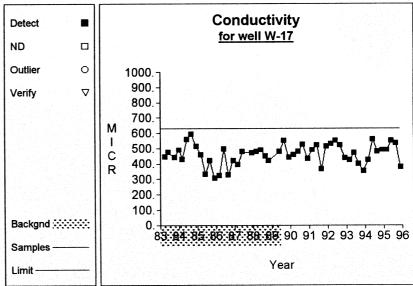


Graph 16

5 96

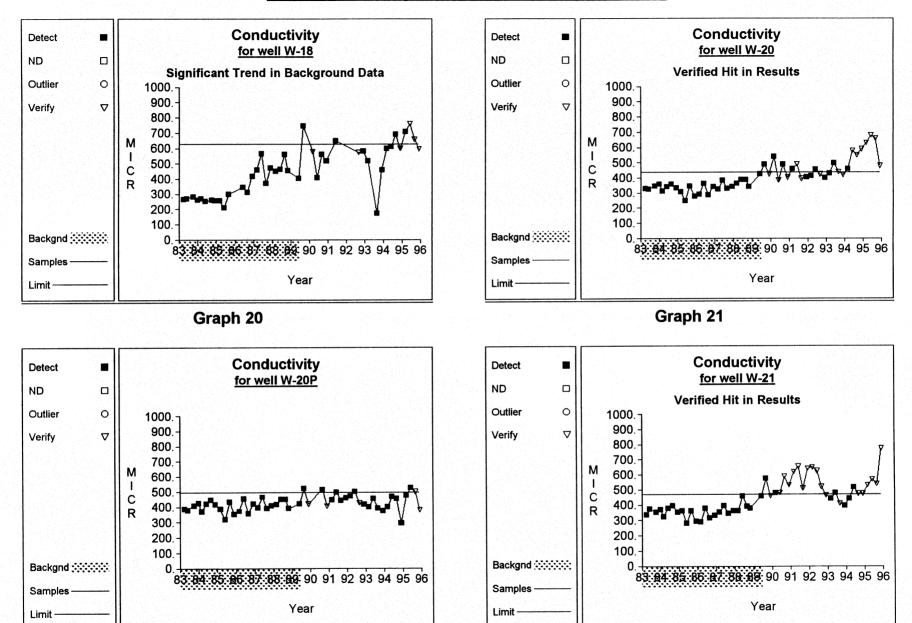
Graph 17





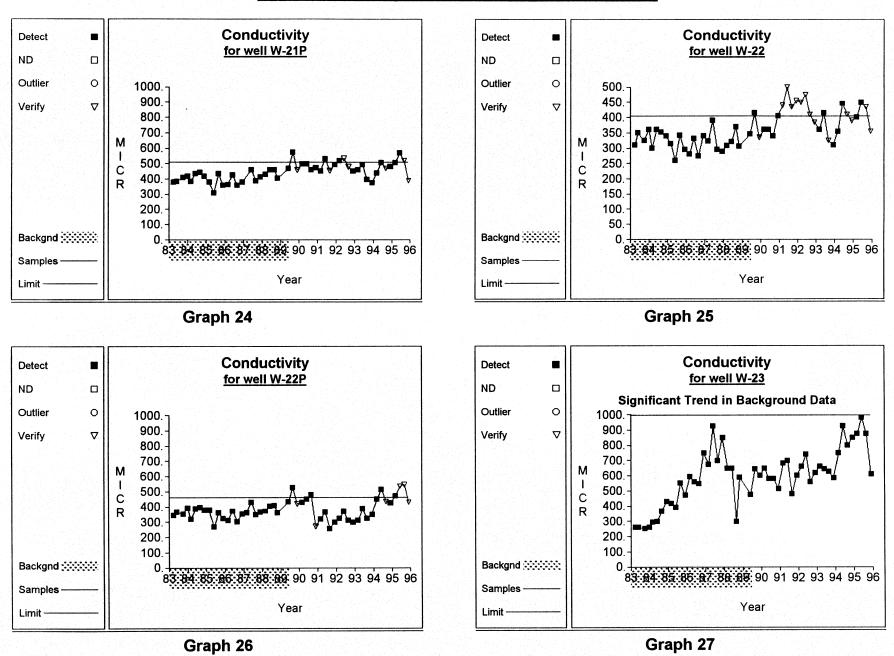
Graph 18

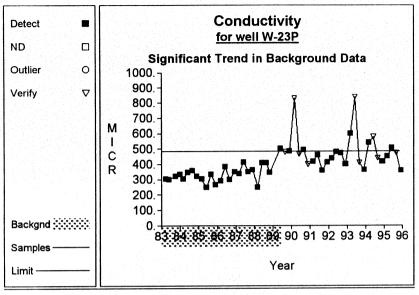
Graph 19

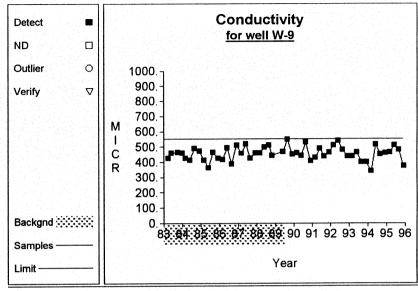


Graph 23

Graph 22

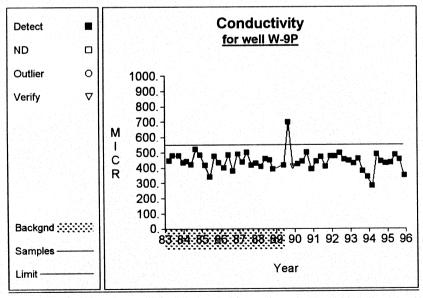




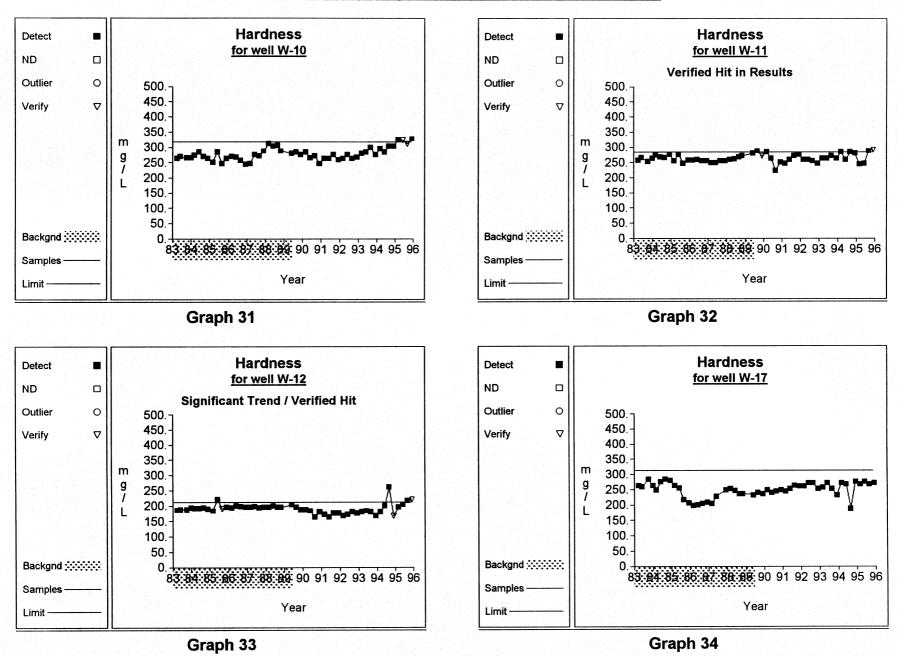


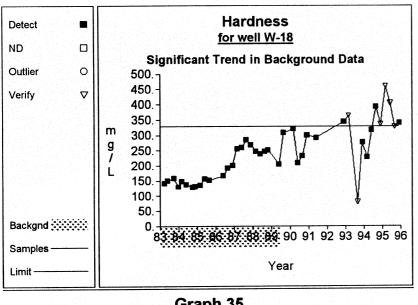
Graph 28

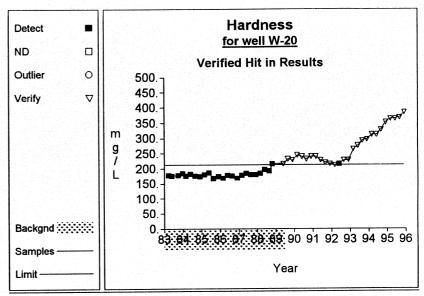
Graph 29



Graph 30

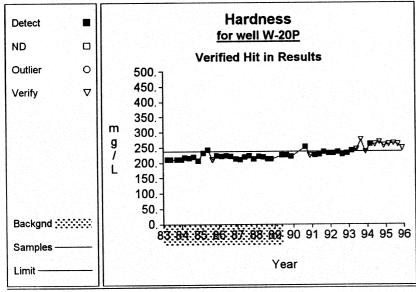


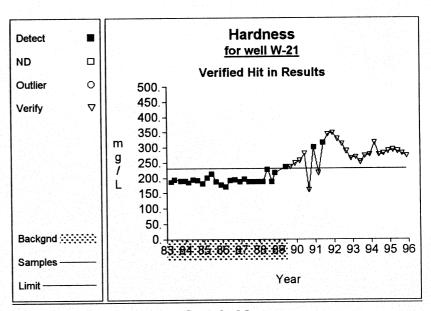




Graph 35

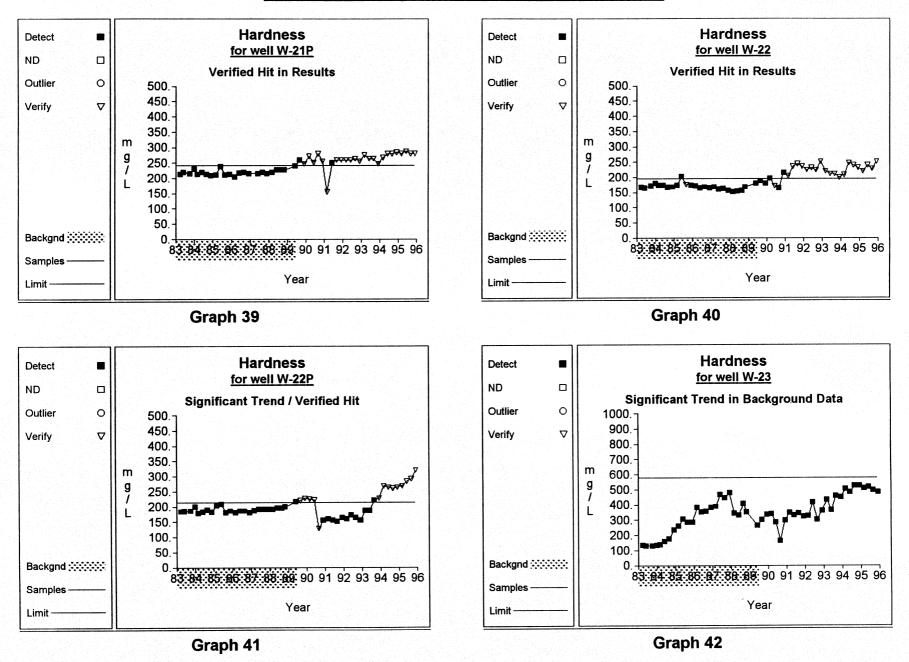
Graph 36

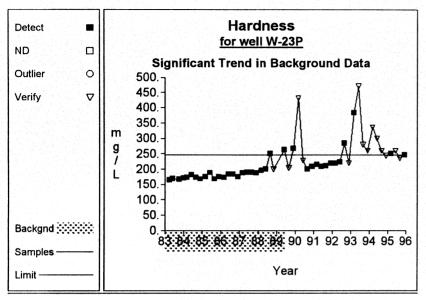


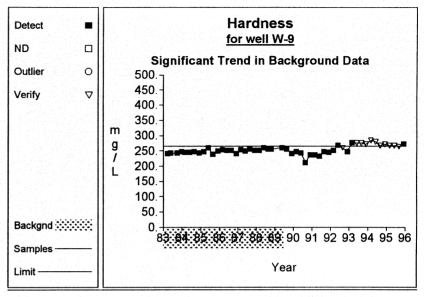


Graph 37

Graph 38

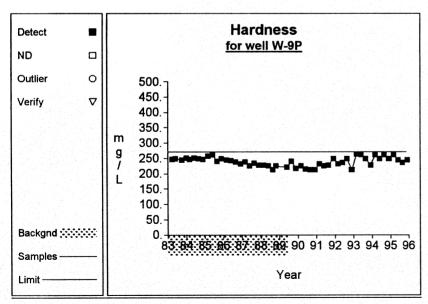






Graph 43

Graph 44



Graph 45

2966 Case Study: Table 1

Summary Statistics and Intermediate Computations for Combined Shewart-Cusum Control Charts

Constituent	Units	Well	N	Mean	SD	S(i-1)	S(i)	Limit
Alkalinity	mg/L	W-10	25	235.496	12.211	466.060	497.406	284.341
Aikaiiiity	mg/L	W-11	25	242.456	9.041	333.437	352.200	278.618
	mg/L	W-12	25	179.752	4.625	201.311	218.090	198.251
	mg/L	W-17	24	228.771	24.283	246.805	247.822	325.902
	mg/L	W-18	23	183.148	55.121	1480.928	1560.439	403.633
	mg/L	W-20	25	170.792	9.461	2143.594	2313.706	208.637
	mg/L	W-20P	25	192.756	21.824	331.392	502.268	280.052
	mg/L	W-21	25	181.352	10.310	2030.984	2085.900	222.592
	mg/L	W-21P	24	204.583	7.217	1091.697	1137.701	233.450
	mg/L	W-22	25	160.908	8.092	1077.453	1146.476	193.278
	mg/L	W-22P	25	179.516	8.227	718.337	824.651	212.426
	mg/L	W-23	25	281.812	112.164	1146.463	1228.528	730.467
	mg/L	W-23P	25	170.020	19.355	1657.678	1701.141	247.442
	mg/L	W-9	25	234.612	6.527	472.015	492.508	260.721
	mg/L	W-9P	25	217.052	11.706	217.052	217.052	263.877
Conductivity	MICR	W-10	25	509.320	65.588	648.788	550.277	771.670
	MICR	W-11	25	466.400	54.160	479.380	466.400	683.041
	MICR	W-12	25	375.360	29.590	400.255	375.360	493.720
	MICR	W-17	24	445.792	71.721	554.297	445.792	732.676
; ;	MICR	W-18	23	355.478	106.569	2901.945	3062.540	781.755
	MICR	W-20	25	338.560	38.396	3228.563	3341.206	492.144
	MICR	W-20P	25	405.400	35.342	583.264	531.358	546.769
	MICR	W-21	25	360.040	43.316	3555.281	3936.753	533.305
	MICR	W-21P	24	407.417	39.410	1490.147	1443.173	565.055
	MICR	W-22	25	323.200	32.087	1644.832	1652.566	451.548
	MICR	W-22P	25	363.520	38.352	869.818	907.534	516.926
	MICR	W-23	25	500.600	194.489	1801.799	1765.332	1278.556
	MICR	W-23P	25	338.320	56.851	3088.485	3067.526	565.725
	MICR	W-9	25	453.560	39.050	477.865	453.560	609.761
	MICR	W-9P	25	441.560	41.999	441.560	441.560	609.556
Hardness	mg/L	W-10	25	272.652	17.988	383.508	425.365	344.604
	mg/L	W-11	25	261.848	8.922	281.308	304.768	297.538
	mg/L	W-12	25	195.068	6.947	198.790	218.512	222.855
	mg/L	W-17	24	241.129	27.921	280.850	290.780	352.813
	mg/L	W-18	23	190.604	53.857	1553.668	1662.671	406.034
	mg/L	W-20	25	181.736	12.301	2170.797	2363.835	230.942
	mg/L	W-20P	25	216.884	7.977	650.412	675.545	248.793 253.625
	mg/L	W-21	25	193.792	14.958	2069.493	2136.483	1
	mg/L	W-21P	24	219.346	8.719	1106.556	1160.671	254.223

^{* -} Insufficient Data

^{** -} Detection Frequency < 25%

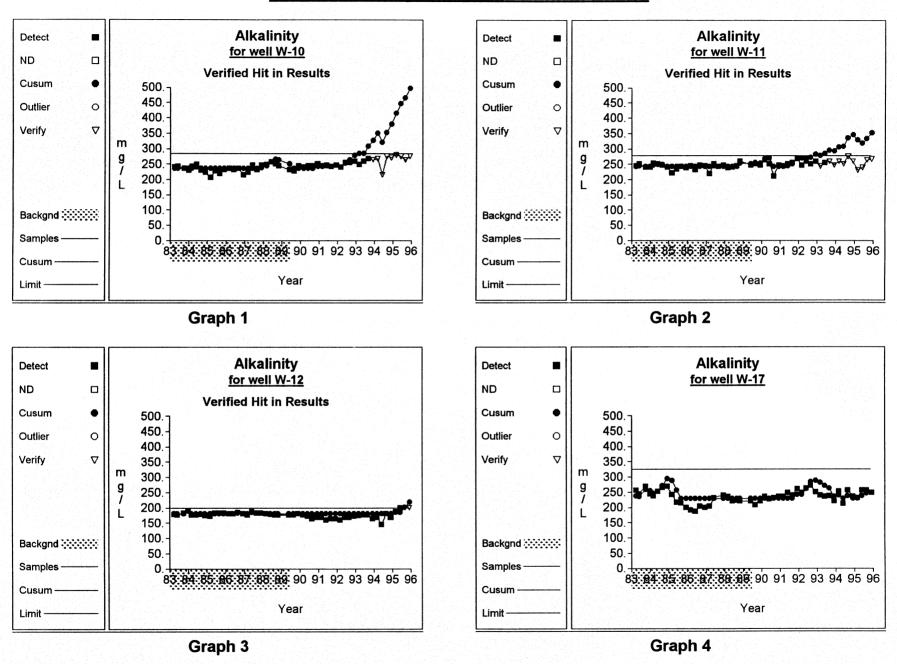
^{*** -} Zero Variance

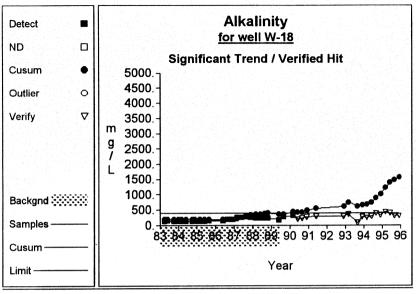
2966 Case Study: Table 1 - Continued

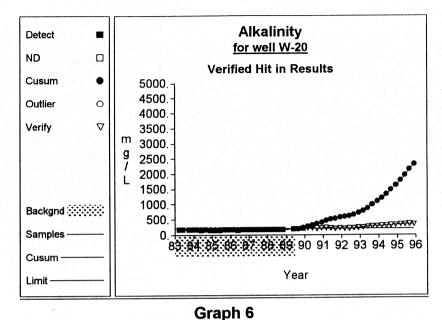
Summary Statistics and Intermediate Computations for Combined Shewart-Cusum Control Charts

Constituent	Units	Well	N	Mean	SD	S(i-1)	S(i)	Limit
Hardness	mg/L	W-22	25	168.528	10.430	1151.764	1227.414	210.248
	mg/L	W-22P	25	190.584	9.517	728.809	851.087	228.653
	mg/L	W-23	25	293.104	113.196	1253.091	1359.089	745.890
	mg/L	W-23P	25	186.452	23.986	1746.533	1790.092	282.396
	mg/L	W-9	25	249.512	6.490	446.578	464.199	275.472
	mg/L	W-9P	25	238.876	12.225	266.386	262.342	287.775

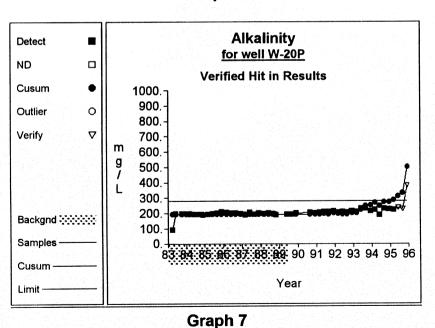
^{* -} Insufficient Data
** - Detection Frequency < 25%
*** - Zero Variance





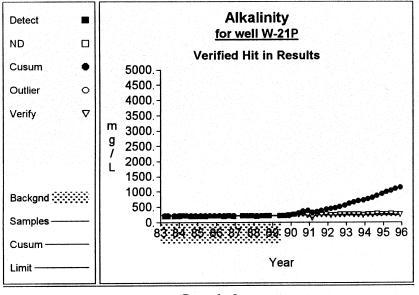


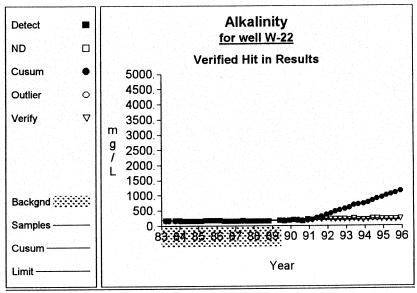
Graph 5



Alkalinity Detect for well W-21 ND Verified Hit in Results Cusum • 5000. 4500. 0 Outlier 4000. ∇ Verify 3500. 3000. 2500. 2000. 1500. 1000 Backgnd ::::::::: 500 Samples 83:84:85 86 87 88 88 90 91 92 93 94 95 96 Cusum Year Limit

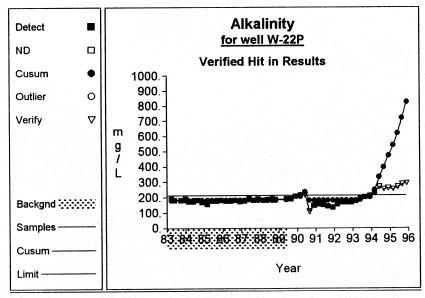
Graph 8

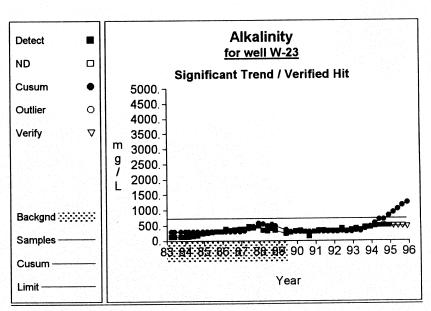




Graph 9

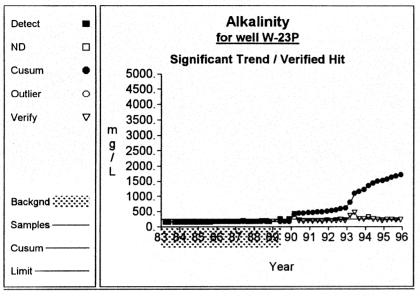
Graph 10

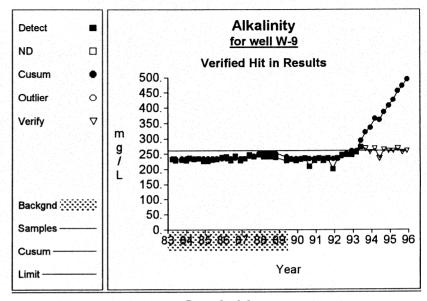




Graph 11

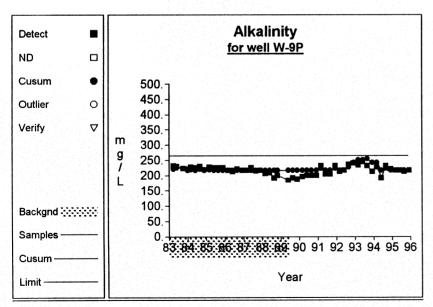
Graph 12



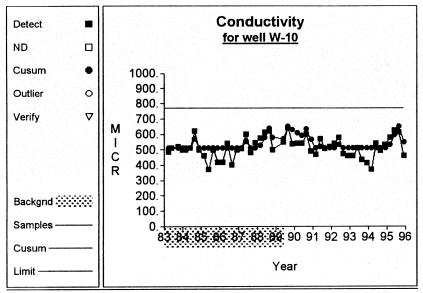


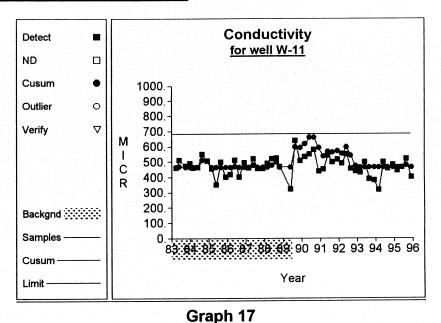
Graph 13

Graph 14



Graph 15





Graph 16

Conductivity Detect for well W-12 ND Cusum 500. 450 0 Outlier 400 ▽ Verify 350 300 250. C 200 150. 100. Backgnd ::::::::: 50. Samples 83 84 85 96 87 88 89 90 91 92 93 94 95 96 Cusum Year Limit

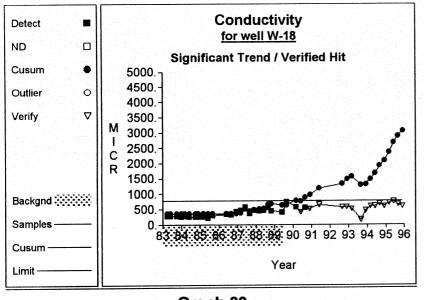
Conductivity Detect for well W-17 ND Cusum 1000. 900. 0 Outlier 800. ∇ Verify 700. 600 500 C 400 R 300. 200. Backgnd ::::::::: 100. Samples 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Cusum

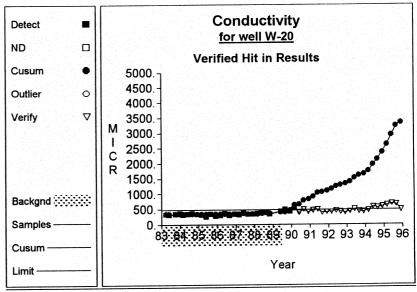
Graph 18

Graph 19

Year

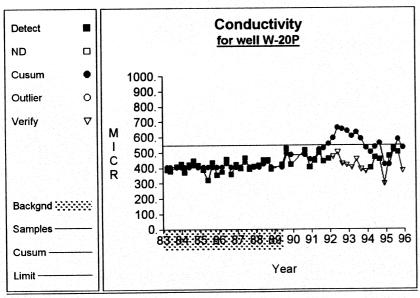
Limit

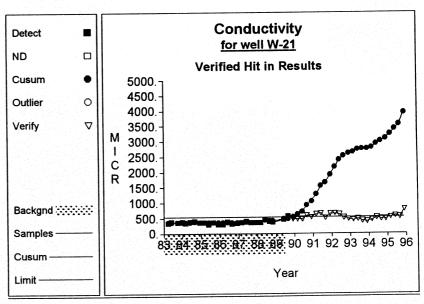




Graph 20

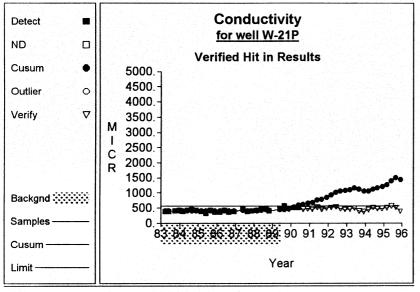
Graph 21

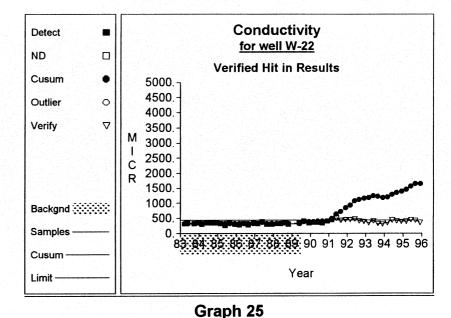




Graph 22

Graph 23





Graph 24

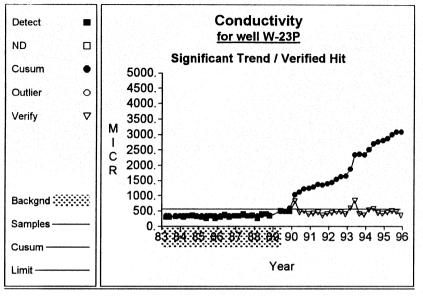
Conductivity Detect for well W-22P ND Verified Hit in Results Cusum 1000. 900. 0 Outlier 800. Verify 700. 600. 500 C 400 300 200. Backgnd ::::::::: 100. Samples 83 84 85 86 87 88 89 90 91 92 93 94 95 96 Cusum Year Limit

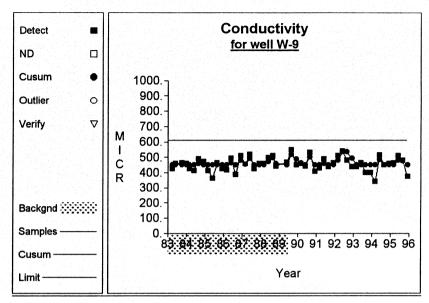
Graph 26

Conductivity Detect for well W-23 ND Significant Trend / Verified Hit Cusum lacktrian5000. 4500. Outlier 0 4000 ∇ 3500. Verify 3000 2500 2000 1500 1000 Backgnd :::::::::: 500 Samples 83:84 85:86 87 88:89 90 91 92 93 94 95 96 Cusum Year Limit

Graph 27

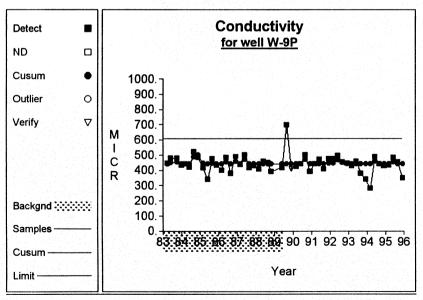
Prepared by: UW-Madison, CEE Department



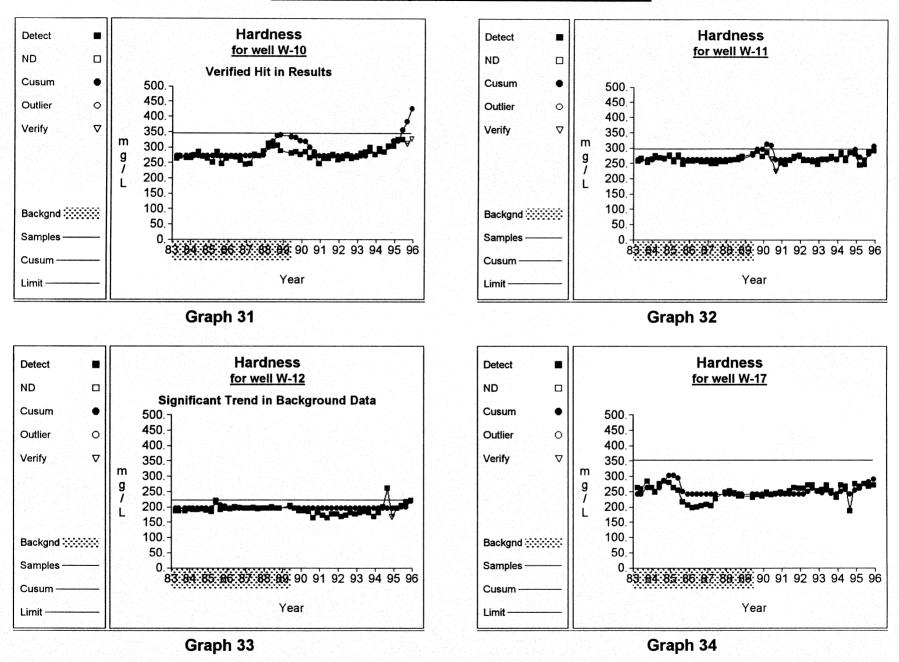


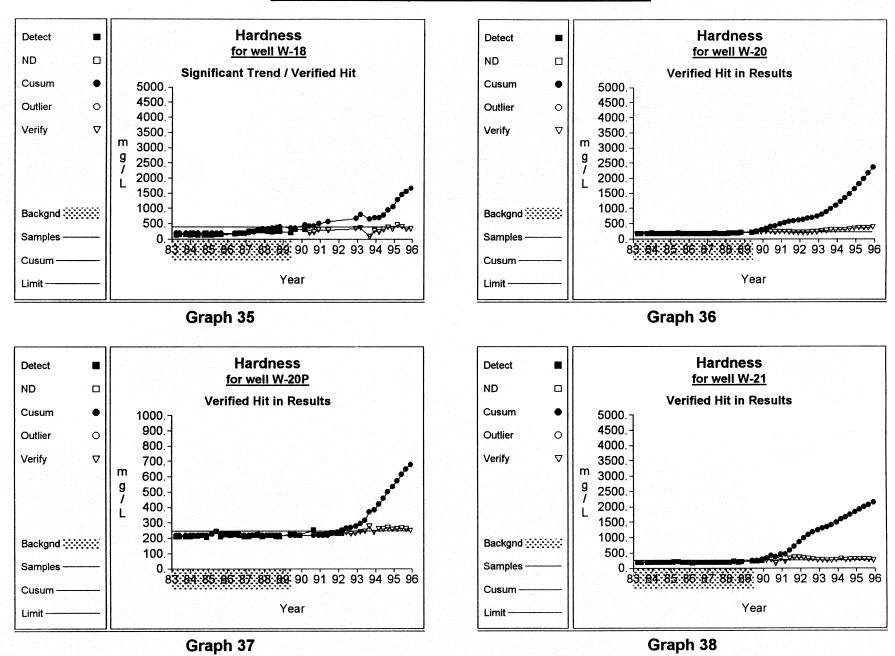
Graph 28

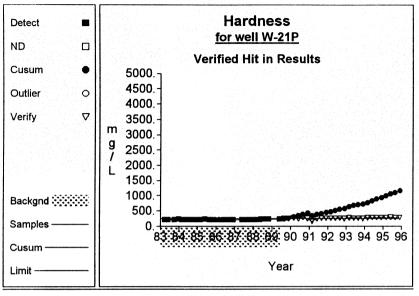
Graph 29

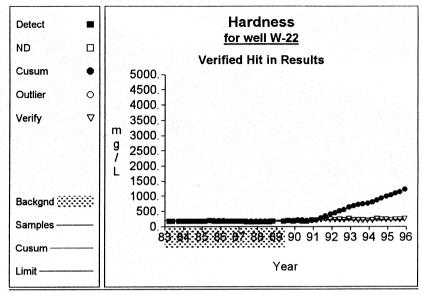


Graph 30



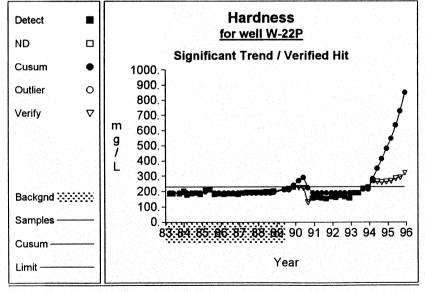


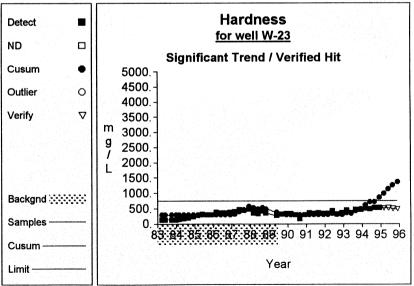




Graph 39

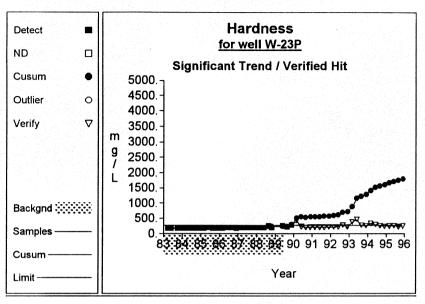
39 Graph 40

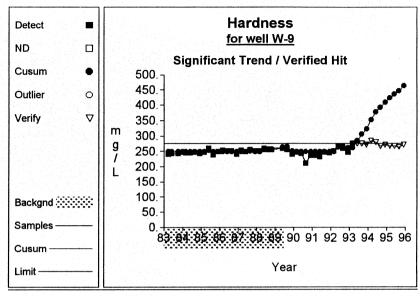




Graph 41

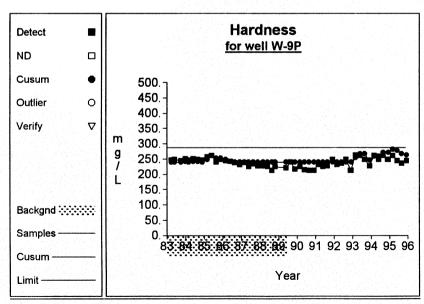
Graph 42





Graph 43

Graph 44



Graph 45